



BANCA D'ITALIA
EUROSISTEMA

Temi di discussione

(Working Papers)

Connecting the dots:
the network nature of shocks propagation in credit markets

by Stefano Pietrosanti and Edoardo Rainone

December 2023

Number

1436



BANCA D'ITALIA
EUROSISTEMA

Temi di discussione

(Working Papers)

Connecting the dots:
the network nature of shocks propagation in credit markets

by Stefano Pietrosanti and Edoardo Rainone

Number 1436 - December 2023

The papers published in the Temi di discussione series describe preliminary results and are made available to the public to encourage discussion and elicit comments.

The views expressed in the articles are those of the authors and do not involve the responsibility of the Bank.

Editorial Board: ANTONIO DI CESARE, RAFFAELA GIORDANO, MARCO BOTTONE, LORENZO BRACCINI, MARIO CANNELLA, ALESSANDRO CANTELMO, GIACOMO CARACCILO, ANTONIO M. CONTI, ANTONIO DALLA ZUANNA, VALERIO DELLA CORTE, MARCO FLACCADORO, ROSALIA GRECO, ALESSANDRO MORO, STEFANO PIERMATTEI, FABIO PIERSANTI, DARIO RUZZI.
Editorial Assistants: ROBERTO MARANO, MARCO PALUMBO, GWYNETH SCHAEFER.

ISSN 2281-3950 (online)

Designed by the Printing and Publishing Division of the Bank of Italy

CONNECTING THE DOTS: THE NETWORK NATURE OF SHOCKS PROPAGATION IN CREDIT MARKETS

by Stefano Pietrosanti* and Edoardo Rainone*

Abstract

We present a simple model of a credit market in which firms borrow from multiple banks and credit relationships are simultaneous and interdependent. In this environment, financial and real shocks induce credit reallocation across more and less affected lenders and borrowers. We show that the interdependence introduces a bias in the standard estimates of the effect of shocks on credit relationships. Moreover, we show that the use of firm fixed effects does not solve the issue, may magnify the problem and that the same bias contaminates fixed effects estimates. We propose a novel model that nests commonly used ones, uses the same information set, accounts for and quantifies spillover effects among credit relationships. We document its properties with Monte Carlo simulations and apply it to real credit register data. Evidence from the empirical application suggests that estimates not accounting for spillovers are indeed highly biased.

JEL Classification: C30, L14, G21.

Keywords: credit markets, shock propagation, networks, identification.

DOI: 10.32057/0.TD.2023.1436

* Bank of Italy, Directorate general for Economics, Statistics and Research.

Contents

1. Introduction	5
1.1 Contributions to the related literature	8
2. Model.....	11
2.1 The traditional identification challenge, with networks	11
3. The econometric framework.....	20
3.1 The credit network model.....	20
3.2 Identification.....	22
3.3 Overlapping portfolios instrumental variables	24
3.4 Bias of ICM under interdependent credit relationships.....	29
4. Monte Carlo study	32
4.1 Setting.....	32
4.2 Results with simulated networks	33
4.3 Results with networks sampled from the credit register	40
5. Empirical application.....	43
5.1 Data description.....	44
5.2 Empirical specification	46
5.3 Main results	47
6. Extensions and discussions	52
6.1 Endogenous credit relationships network.....	52
6.2 Endogenous treatments.....	54
6.3 Heterogeneous BCR and FCS	55
6.4 Comparison between OPIVs, SSIVs, and GIVs.....	57
6.5 Comparison with spillover papers in corporate finance	58
7. Conclusion.....	59
Appendix: additional tables and proofs.....	61
References	76

1 Introduction¹

Most research in empirical banking treats credit relationships in isolation, as determined only by firm, bank and relationship’s characteristics. Recent attempts at relaxing this assumption have mostly focused on competitive interaction among firms and local spillovers due to sharing geographic locations. To the best of our knowledge, no recent work *jointly* addressed simultaneity in the borrowing decision of the same firm from multiple banks, and in the lending decision of the same bank to multiple firms.

However, credit relationships are embedded in a network of links, originated by sharing the firm or bank to which they belong. As such, shifts in banks’ supply to each firm change that firm’s relative cost of borrowing across its entire portfolio of relationships. Conversely, shifts in firms’ demand for credit change the opportunity cost of lending for each bank, as equity and managerial resources are, after a certain point, scarce.

This paper aims at identifying the importance of spillover effects for the transmission of financial shocks over credit markets. We apply recent advancements in network econometrics to study how ignoring spillovers can affect standard estimates of the bank lending channel (e.g. Khwaja and Mian, 2008; Peek and Rosengren, 2000).² Moreover, we present a feasible way to correct the measurement of idiosyncratic shifts in credit demand and supply, as well as the assessment of aggregate effects in the presence of correlated demand shocks (Amiti and Weinstein, 2018; Jiménez et al., 2020), when credit-network spillovers are important.

To present our argument, we introduce an extremely simplified model of credit demand and supply joint determination. We consider the case of two banks and two firms that optimize jointly their full portfolio of relationships, nesting simultaneity in Khwaja and Mian (2008)’s model. We use this framework to argue four points. First, interdependence introduces bias in OLS estimates of the bank lending channel and any other treatment of interest. Second, in the presence of correlated demand shocks, fixed effect estimation may actually worsen the bias.³ Third, and conversely, if spillovers are large but correlated

¹Authors are Stefano Pietrosanti: Bank of Italy, stefano.pietrosanti@bancaditalia.it; Edoardo Rainone: Bank of Italy, edoardo.rainone@bancaditalia.it. We thank Alessio Anzuini, Vladimir Asriyan, Emilia Bonaccorsi di Patti, Francesco Bripi, Christian Brownlees, Andrea Caggese, Giuseppe Cappelletti, Francesco Columba, Guido Deiana, Alessio De Vincenzo, Kirill Evdokimov, Giovanni Guazzarotti, Louis-Marie Harpedanne de Belleville, Dmitry Kuvshinov, Xiaodong Liu, Geert Mesters, Alessandro Peri, Andrea Polo, Bjorn Richter, Gianmarco Ruzzier, Andres Sarto, Enrico Sette and seminar participants at the Bank of Italy, EIEF, UPF econometrics and finance seminar series, Colorado Boulder and the Stockholm School of Economics for their comments. All the errors are our own. The views expressed in the paper do not necessarily represent those of the Bank of Italy.

² The name often given in the empirical literature to the sensitivity of credit supply to banks’ shocks.

³ Using firms with multiple bank relationships to perform a fixed effect estimate of supply movements is a standard procedure in the empirical banking literature, first popularized in Khwaja and Mian (2008). A far from complete list of influential works using this strategy in order to assess the effect of different bank shocks includes Jimenez et al. (2012), Schnabl (2012), Jiménez et al. (2014), Behn, Haselmann, and Wachtel (2016), Bonaccorsi di Patti and Sette (2016), Jiménez et al. (2017).

demand shocks are not, OLS and fixed effects estimates may still differ. As a consequence, interpreting their difference as a measure of demand bias (e.g. Jiménez et al., 2020) may be misleading. Fourth and last, spillovers may contaminate the OLS estimates of firm and bank fixed effects, stopping us short of measuring pure demand and supply shifts (Amiti and Weinstein, 2018).⁴

The key challenge in addressing these concerns is the reflection problem (Manski, 1993). Reflection is the type of endogeneity that arises when shocks to one agent affect all other agents, and can induce extremely complex distortion patterns. We proceed laying down an estimation framework to address such distortions, building on the literature on spatial autoregressive models (e.g. Arduini, Patacchini, and Rainone, 2015; Bramoullé, Djebbari, and Fortin, 2009; Calvó-Armengol, Patacchini, and Zenou, 2009; Lee, 2007) as extended in Rainone (2020a), to model outcomes on links (credit relationships), instead of nodes (firms or banks). We propose a new method to construct instrumental variables based on overlapping portfolios (OPIVs).

We show that the very structure of the credit market network, in which many nodes are only indirectly connected, provides instruments for spillover identification. Indeed, firms borrowing from multiple lenders create overlapping portfolios of relationships at the firm and bank level. As idiosyncratic shocks to one relationship can affect the others involving the same parties, they can provide relevant instruments for the spillover effects. If these portfolios are not fully overlapping, such instruments are exogenous. We show how to use these OPIVs to identify spillovers and recover unbiased treatment effects, as well as idiosyncratic demand and supply movements.

The model we propose separately identifies two types of spillovers arising from the network nature of credit relationships, uncovering two channels that affect credit equilibrium outcomes. These are the *bank supply allocation effect* and the *firm demand allocation effect*. They capture the endogenous adjustment of credit supply and demand across all relationships of the same bank and firm after a shock. Noticeably, even if we focus on negative spillovers in most of the present draft, this assumption is not necessary in our framework. In future developments of this paper, we plan to study spillover heterogeneity to better understand the mechanics of shock propagation in the credit market.

Indeed, negative spillover estimates imply that substitution effects across firms' or banks' portfolios dominate endogenous demand or supply adjustments. We would expect such findings in most "normal times" situations. For example, if a monetary policy tightening induces a large contraction in supply for some banks, firms may reallocate their demand to absorb the hit. Nonetheless, even cases of complementarity across different credit relationships could be very relevant. For instance, during a sizable financial shock, banks could internalize the negative effect of the global credit contraction on firms'

⁴ This point is related to the one risen by Altavilla, Boucinha, and Bouscasse (2022). We elaborate on the differences between the two contributions in the *Related Literature* part of this Introduction.

prospects. In such a case, every decision to cut credit could reinforce the other, with positive spillovers at the bank level. Furthermore, we can augment the model to study bank- and firm-type specific spillovers and, eventually, combinations of the two depending on the specific empirical question.⁵

We explore the properties of the econometric model through Monte Carlo simulation. First, we study the bank lending channel’s estimate bias if we ignore the network nature of credit relationships. We show that both the magnitude and sign of the bias depend on *observables*, as the share of relationships hit by a shock, the number of relationship sharing firms or banks (the density of the network), and *unobservables*, as the size of spillovers. Hence, such bias cannot be addressed without a strategy to estimate cross-network demand and supply adjustments. Second, we confirm that fixed effects can exacerbate spillover bias. What is more, we show that standard estimates of bank and firm fixed effects may pinpoint nonexistent demand and supply idiosyncratic movements, especially for banks and firms that are highly connected (have high centrality in the network). Third, we document that the network estimator performs very well in finite samples, estimating spillovers, treatment effects and idiosyncratic shocks with negligible error. We confirm the economic significance of these results calibrating the network structure of our Monte Carlo simulations to mimic real networks of credit relationships observed in the Bank of Italy’s Credit Register.

Finally, we use information from 2012 to 2018 to show that spillover effects play an important role in the Italian credit market. We use complete data from the Bank of Italy’s Credit Register matched with firms’ financials from CERVED, the main Italian risk rating issuer, and banks’ financials from Italian Supervisory Reports. We exploit changes in credit granted due to changes in the interbank rate, mediated by each relationship’s exposure to other connected relationships’ revolving credit fraction. Revolving credit is a natural channel of idiosyncratic shocks’ transmission across firms and banks. For example, firms use revolvers as a buffer for unforeseen needs (see, e.g. Acharya and Steffen, 2020), exposing banks to shocks hitting firms. On the other hand, revolving credit contracts are renegotiated more often, as banks can easily change rates or adjust the granted limit. Hence, firms more dependent on revolvers may be more affected by shocks hitting banks.

All spillover parameters are significant, with the firms’ one being particularly large. We demonstrate that spillovers lead to a significant empirical bias in treatment and idiosyncratic effects’ estimates. Our exercise suggests that ignoring network spillovers would lead to overestimating the effect of interest rate changes on more revolving-intensive credit relationships by a factor of two. Furthermore, focusing on estimated fixed effects, we suggest that spillover bias may lead to an underestimation of firm fixed effects by a

⁵ For example, one may be interested in studying the substitution of credit from low-tech banks to high-tech banks by firms or the reallocation of credit from a specific type of firms to another one by banks.

factor of a half to three-fourths, as well as a two-thirds to one overestimation of bank fixed effects.

1.1 Contributions To The Related Literature

We contribute to the empirical literature on financial shocks' pass-through in two ways. First, we build a methodology to account for endogenous reallocation of demand and supply induced by shocks. Second, we apply it, deriving numerical and empirical evidence on the importance of the credit market structure for the pass-through of shocks, and quantifying how endogenous reallocation plays an important role in mediating the effect of monetary policy.

The methodology we propose addresses an important limitation of standard instrumental variable (e.g. Paravisini, 2008; Peek and Rosengren, 2000) and within-firm (Jiménez et al., 2014, 2017; Khwaja and Mian, 2008) estimates of bank shocks' effect on firms' credit access. These approaches cannot control for endogenous reallocation of demand and supply across portfolios of credit relationships. We show how this may affect (i) attempts at using the difference between OLS and fixed effects estimates to track the extent of demand bias (Jiménez et al., 2020), and (ii) attempts at using estimates of bank and firm fixed effects as direct measures of pure credit demand and supply shifts (Amiti and Weinstein, 2018). We offer a solution to these problems, based on novel and classical results in network econometrics (e.g. Ballester, Calvó-Armengol, and Zenou, 2006; Calvó-Armengol, Patacchini, and Zenou, 2009).⁶

Such contribution is complementary to works deepening the identification of financial shocks' pass-through to credit and real outcomes. While we focus on the identification challenges stemming from endogenous substitution of credit, Paravisini, Rappoport, and Schnabl (2023) highlights those implied by a lack of perfect substitutability. In particular, it shows that bank specialization causes shocks to distribute unevenly across the relationships of the same firm, limiting the effectiveness of within-firm strategies to control for unobservable confounders.⁷ Research building on these premises, such as Bripi (2021) and Altavilla, Boucinha, and Bouscasse (2022), estimates credit demand systems, focusing on the interest rate elasticity of credit. These works, though, need price data for estimation and do not address cross-substitution within the same firm or bank's relationship portfolios.

⁶ See Jackson (2010), Jackson and Zenou (2015) and Jackson, Rogers, and Zenou (2017) for a complete critical survey of the theoretical literature on the economics of networks. See Bramoullé, Djebbari, and Fortin (2020), De Paula (2020) and Graham and De Paula (2020) for insightful reviews of the literature on network econometrics.

⁷ In the same thread, though not focusing on identification problems *per se*, are papers such as De Jonghe et al. (2020) and Giometti and Pietrosanti (2022). Within-firm differences across relationships are also a standard topic of investigation in the relationship-lending literature. See, e.g., Petersen and Rajan (1994), Berger and Udell (1995), Bharath et al. (2007), Bartoli et al. (2013), Dewally and Shao (2014), Bolton et al. (2016).

Moreover, our work closely relates to the growing strand of papers in macro-finance and banking focusing on spillovers' implications for the identification of shocks' effects. For example, Mian, Sarto, and Sufi (2022) proposes a method to recover general equilibrium multipliers from differences in the regional impact of credit supply shocks. Berg, Reisinger, and Streitz (2021) studies how spillovers from firms' interactions may affect the measurement of shocks' effect on firm level outcomes. Moreover, it highlights how within group (fixed effect) identification may worsen spillover bias, in line with our evidence. Finally, Huber (2022) provides overall guidance to empirical researchers on how to deal with multiple contemporaneous spillovers (spatial, competitive, agglomeration, general equilibrium and so on) and in the presence of mechanical bias from the mismeasurement of treatment.

Whereas these works focus on the impact of peers' treatment status (being more or less hit by shocks, i.e., contextual effects) in reduced form, we uniquely focus on identifying and controlling for the effects of peers' outcomes (endogenous effects). Our approach leads us to recover parameters that can be precisely mapped to the primitives of a structural model, thus having a direct behavioral interpretation. Measuring these primitives is especially interesting in the case of credit relationships. The bank-firm relationship is a key determinant of financial shocks' pass-through, but its working is a black box. Our model allows us to unpack the black box and uncover the importance of bank supply and firm demand allocation across different relationships to determine the effect of shocks in equilibrium.

Our investigation also complements the recent works by Darmouni and Sutherland (2021) and Gupta et al. (2023), addressing spillovers in banks' credit contract design, the first in SME lending and the other in syndicated lending to large firms, both in the US. The latter paper, in particular, exploits the network structure of banks' overlapping lending portfolios through a spatial autoregressive model to estimate the degree of complementarity in banks' interest rates setting. Nonetheless, using Dealscan data which does not track credit commitments' evolution over time, Gupta et al. (2023) cannot follow changes in credit quantity and, ultimately, how shocks bolster or impair credit access. Instead, we exploit each bank's borrowers' and each firm's lenders' networks to do precisely that.

From a methodological standpoint, our work is close to other corporate finance papers that directly address Manski (1993)'s reflection problem. Recent works did so to quantify peer effects in firms' capital structure (Grieser et al., 2022; Leary and Roberts, 2014), corporate governance (Foroughi et al., 2022), and banks' liquidity choices (Silva, 2019). We cover firm access to credit and financial shocks' pass-through. From the financial shocks' pass-through perspective, our contribution is also complementary to Alfaro, García-Santana, and Moral-Benito (2021) and Huremovic et al. (2020). The latter two works study the propagation of financial shocks across the production network among

firms. Instead, we focus on the propagation of shocks through the credit network among bank-firm links.

More in general, our approach is related to studies decomposing market’s aggregate outcomes to derive instrumental variables, such as shift share instrumental variables (SSIVs), used initially in Bartik (1991), Blanchard et al. (1992) and recently in Borusyak, Hull, and Jaravel (2022) and Goldsmith-Pinkham, Sorkin, and Swift (2020), and granular instrumental variables (GIVs), proposed in Gabaix and Koijen (2020) and applied to banking in Galaasen et al. (2020). However, our approach differs from the GIVs and SSIVs approaches substantially. The GIVs and SSIVs are procedures designed to estimate price elasticities, while OPIVs is designed to estimate objects more similar to elasticities of substitution. OPIVs and both these approaches are actually complements, because they can be used in different types of markets. GIVs and SSIVs consider centralized markets where there is only one price, while the OPIVs consider decentralized markets where the price varies at the pair level. In Section 3.3, where we introduce the OPIVs, we compare them with SSIVs and GIVs in more detail.

In suggesting and applying this approach, we provide our second contribution: A unique quantification of spillover effects across credit relationships, and empirical evidence on the role of credit market structure for shock propagation. Such quantification is closely related to a recent stream of works measuring the link between credit market structure and the effect of financial shocks. Important examples are Andreeva and García-Posada (2021); Benetton (2021); Benetton and Fantino (2021); Corbae and D’Erasmus (2021); Giannetti and Saidi (2019). Again, our work is complementary to Huremovic et al. (2020) and Alfaro, García-Santana, and Moral-Benito (2021). Whereas they document that financial shock propagation worsens when firms’ market power is greater, we show that banks’ market power has a similar effect in our Monte Carlo study. This distinct amplification mechanism may counterbalance the stabilizing role of credit concentration documented by Giannetti and Saidi (2019).

Finally, and with respect to other works studying peer effects in networks, our context allows us to achieve identification of spillover effects under lighter assumptions. Differently from social networks, credit networks are easier to observe. The detail in credit register data mitigates the concern of unobservable relevant connections, a widespread worry when studying other types of interactions (see Battaglini et al., 2020; Battaglini, Patacchini, and Rainone, 2022; De Paula, Rasul, and Souza, 2019; Miraldo, Propper, and Rose, 2021, among others). Indeed, the crucial assumption of non-overlapping bank portfolios is testable and always verified in our data.

The rest of this paper is structured as follows: In Section 2 we introduce the toy model of a credit relationships network. In Section 3 we lay down the estimation framework, and explain how network econometrics allows us to identify spillover effects and recover unbiased estimates of spillovers, treatment effects and idiosyncratic shocks. In Section 4

we explore the estimator’s characteristics in finite samples, documenting bias behavior if spillovers are ignored and the unbiasedness of our estimator. In Section 5 we use Italian credit register data to show that spillover effects through firm links are highly statistically and economically significant. Section 6 presents a few extensions to the model. We take stock in Section 7.

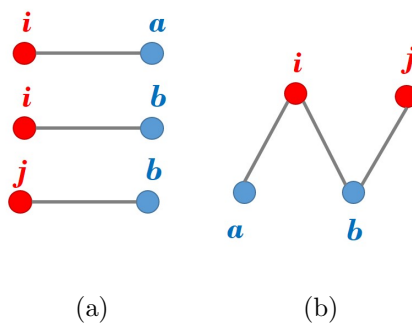
2 Model

In Section 2.1 we draft a simple example to highlight how (i) network spillovers affect the identification of financial shocks’ effects; (ii) network spillovers can harm the effectiveness of fixed effects as controls for demand and supply biases; (iii) network spillovers may distort firms and banks’ fixed effects estimates, lowering their correlation with actual idiosyncratic demand and supply shifts.

2.1 The Traditional Identification Challenge, With Networks

We present the logic of our results using a simple static model of bank-firm relationships, which modifies Khwaja and Mian (2008)’s to allow firms and banks to optimize simultaneously their full portfolio of relationships. Consider a network of two firms, i and j , and two banks, a and b . Bank a and b both lend to firm i , while bank b also lends to firm j . Figure 1 provides the visual counterpart. While the standard approach takes credit relationships in isolation (panel (a)), we stress how they form a network (panel (b)). Credit relationship ia links to relationship ib , as they share the same borrower (i), and to jb , as they share the same lender b .

Figure 1: A Toy Credit Network



Notes: Nodes a, b are banks, i, j are firms, edges are credit relationships.

The following is a static model of an exogenous network between banks and firms. Banks fund their loans with a given mix of costless (insured) deposits and costlier resources (equity, uncovered bonds, etc). As loans cannot be fully funded with costless

resources, banks face a resource allocation problem, which gives rise to the bank lending channel, i.e. the elasticity of credit to sudden changes in banks' balance sheets. In the application (Section 5), we will focus on the classic bank lending channel of monetary policy (Jiménez et al., 2014; Kashyap and Stein, 2000). By the “bank lending channel of monetary policy”, we mean the sensitivity of heterogeneous banks' lending to monetary policy changes. We start assuming that such elasticity is the same for all relationships. We make this choice to isolate our point clearly at the start, but heterogeneity in effects is very important too, mattering for identification. We will explore the topic of heterogeneity in further updates to this draft.

Formally, we assume that banks set their credit supply maximizing:

Assumption 1. $\pi_b(c_{ib}, c_{jb}) = (r_{ib} - \omega(c_{ib}, x_{ib}, c_{jb}, \nu_{ib}))c_{ib} + (r_{jb} - \omega(c_{jb}, x_{jb}, c_{ib}, \nu_{jb}))c_{jb}$
 where : $\omega(c_{ib}, x_{ib}, c_{jb}, \nu_{ib}) = \omega \frac{c_{ib}}{2} - \omega(\xi x_{ib} + \nu_{ib} - \theta c_{jb})$

We specify the ω cost function as linear, ω is thus a parameter that captures the baseline cost to the bank of one more dollar of commitment.⁸ Such functional form adapts Khwaja and Mian (2008) to our setting. $r_{i,jb}$ are the interest rates bank b obtains for its credit from firms i and j ; c_{ib}, c_{jb} are the quantity of credit supplied to firms i and j ; $x_{i,jb}$ are our **treatments of interest**, tracking some relationship's characteristic that changes bank b 's marginal cost of lending to the specific firm by $-\xi$ dollars when monetary policy changes; $\nu_{i,jb}$ are unobservable random shifts in supply.

At difference with Khwaja and Mian (2008), c_{jb} enters the supply choice to firm i directly (and viceversa), capturing the supply-side of interdependence in lending decisions. Everything else equal, if bank b already lends one more dollar to firm j , this rises the cost of lending to i by θ dollars, capturing the scarcity in bank resources. The resulting supply equations follow from first order conditions (we display derivations in Appendix A.1):

$$\begin{aligned}
 r_{ib} &= \omega c_{ib} - \underbrace{\omega(\xi x_{ib} + \nu_{ib} - \theta c_{jb})}_{u_{ib}} \\
 r_{jb} &= \omega c_{jb} - \underbrace{\omega(\xi x_{jb} + \nu_{jb} - \theta c_{ib})}_{u_{jb}} \\
 r_{ia} &= \omega c_{ia} - \underbrace{\omega(\xi x_{ia} + \nu_{ia})}_{u_{ia}}
 \end{aligned} \tag{1}$$

and enter each firm's demand problem as the firm's cost function.

On the firm side, we consider a setting in which firms sustain multiple relationships in equilibrium due to decreasing return to scale (increasing costs) in borrowing more from only one lender (for empirical evidence on the cost of credit captivity, see Ioannidou and

⁸ The assumption of a common ω parameter across banks implies that banks face the same capital market. Deviations from this assumption imply that impact effects are actually heterogeneous in the system we estimate. We will empirically explore the relevance of this possibility in the further sections.

Ongena, 2010).⁹ We start from a case in which, as in Khwaja and Mian (2008) and at difference with Paravisini, Rappoport, and Schnabl (2023), demand shocks will distribute equally across the different relationships of the same firm. At difference with Khwaja and Mian (2008), we add that all credit funds the same project, hence it is perfectly substitutable and there is going to be scope for endogenous reallocation. Formally, the above translates to:

Assumption 2. *The firm knows the banks' pricing rules, and decides its credit demand maximizing $\pi_f(c_{fa}, c_{fb}) = (e_i - \alpha(c_{ia} + c_{ib}))(c_{ia} + c_{ib}) - \sum_{k=a,b} c_{ik}r_{ik}(c_{ik})$.*

Here, e_i is the productivity of firm f 's use of funds and the source of demand shocks, α tracks the quadratic decrease in returns to scale and r_{fK} is the loan's cost derived above. Firms' profit maximization results in the following structural demand system:¹⁰

$$\begin{aligned} c_{ia} &= \rho c_{ib} + \beta x_{ia} + \delta_i + \epsilon_{ia}, \\ c_{ib} &= \rho c_{ia} + \phi c_{jb} + \beta x_{ib} + \delta_i + \epsilon_{ib}, \\ c_{jb} &= \phi c_{ib} + \beta x_{jb} + \delta_j + \epsilon_{jb}. \end{aligned} \tag{2}$$

Relaxing the assumption that banks and firms optimize their choices only at the single relationship level, and allowing them to more realistically maximize their profits considering all their relationships together, we obtain a simultaneous system of equations.

We summarize the convolution of structural parameters in each equation with β , for the supply shift we want to measure; $\delta_{i,j}$ captures firms' demand shifters common across relationships; ρ , the spillover from credit relationships of the same firm; ϕ , the spillover from credit relationships of the same bank. In what follows, we focus on structural parameters (ω, θ, α) such that ρ and ϕ are negative. Which is, we study the case in which the substitutability of credit across different relations of the same firm drives the ρ -spillover, and the opportunity cost of lending across different relationships of the same bank drives the ϕ one. Nevertheless, the model does not need these assumptions *a priori*; complementarities (positive ρ , or ϕ , or both) are also possible.

For example, lenders may be market-leaders in specific credit markets and internalize the externalities implied by cutting lending to some firms (Giannetti and Saidi, 2019). Such internalization would lead to a complementarity in lending to different firms in that market, and thus positive ϕ . On the other hand, banks may be specialized in funding specific projects (Paravisini, Rappoport, and Schnabl, 2023). If we think of an exporting firm, it may deal with a bank which has comparative advantage in funding export towards a certain market, and another bank focused on funding purchases of specif machines useful

⁹ For the moment, we focus on active credit relationships and their intensive margin. We thus ignore corner solutions in which links between a firm and a bank may not be active. We will explore matching models in future extensions.

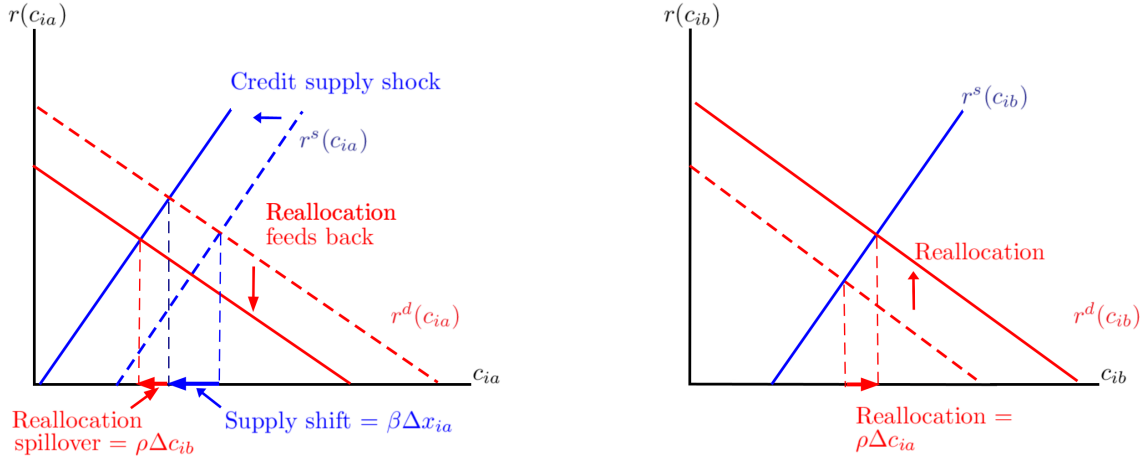
¹⁰ We display derivations in Appendix A.1.

to produce exactly that export good. Thus, if that firms wants to scale up export, it must scale up borrowing from both banks, with a complementarity in demand and a positive ρ .¹¹

In such an environment, the amount of credit that firm i borrows from bank a (c_{ia}) depends on the amount that firm i borrows from bank b . On the other hand, the amount of credit that bank b lends to firm j depends on the amount that bank b lends to firm i (c_{ib}). In turn, c_{ib} depends on both c_{ia} and c_{jb} , i.e. the single relationship's outcome receive impulses from all relationships sharing the same firm or bank. These endogenous effects induce spillovers among credit relationships, based on the structure of links in the credit network.

The model illustrates how simultaneity in credit demand and supply decisions affect the final credit consumption observed in equilibrium. It can be represented graphically as a system of linear supply and demand schedules that are co-determined at the intercept level across relationships belonging to the same firm or bank. In Figure 2, we show how this translates in a dependence of each relationship's amount of credit on both the supply shift we want to measure and reallocation spillovers, due to the change in relative costs.

Figure 2: Interdependence in a Toy Network



Notes: The Figure represents graphically the interaction between firm i and bank a and b , captured by Equation 2, in the case bank a receives a negative supply shock. The blue lines are banks' credit supply schedules to firm i , while the red lines are firm i 's demand for credit from a and b respectively. Solid lines are post-shock curves, while dashed lines are pre-shock curves.

Focusing on the two relationships belonging to firm i in System 2, say a supply shock $\Delta x_{ia} < 0$ hits bank a , contracting its supply curve, while bank b supply stays still.¹² Graphically, the shock moves a 's supply from the old, dashed line, to the new, solid one by $\beta \Delta x_{ia}$. Then, for each amount of credit offered, the bank is asking an higher interest

¹¹ To solve the system, we also assume here that $|\phi| < 1$ and $|\rho| < 1$. This is not strictly required in more general cases. We define formally the space of these parameters in the next sections.

¹² Figure 2 ignores reallocation through banks, i.e. ϕ . This is done just for clarity, as reallocation across different relationships of the same firm is enough to convey the intuition of the problem at hand.

rate to i . As a consequence, i will not only comply with the restriction imposed by a , but demand less from a and more from b overall, as a 's credit became relatively more costly with respect to credit from b . This will translate in a $\rho\Delta c_{ia}$ shift in the equilibrium consumption of c_{ib} , partly compensated by an opposite shift by $\rho\Delta c_{ib} = \rho^2\Delta c_{ia}$ in c_{ia} . In conclusion, the final change in equilibrium credit from bank a to firm i will be composed by two elements, one directly related to the supply shift and governed by parameter β , and one indirect, due to reallocation and governed by parameter ρ .

We then consider how all the above relates to the standard identification problem in the empirical banking literature, the recovery of a treatment effect β , which can be correlated with firms' demand shocks. To further simplify the example, we assume that

Assumption 3. *Treatment only hits relationship ia ($x_{ib}, x_{jb} = 0$), $\delta_j = 0$, and $\beta < 0$ to mimic a contraction in credit.*

If the researcher ignores the simultaneity of choices across credit relationships and estimates the following system of equations instead of (2),

$$\begin{aligned} c_{ia} &= \beta x_{ia} + \varepsilon_{ia}, \\ c_{ib} &= \beta x_{ib} + \varepsilon_{ib}, \\ c_{jb} &= \beta x_{jb} + \varepsilon_{jb}. \end{aligned} \tag{3}$$

the OLS estimator for β will be biased:

Proposition 1. *Under Assumptions 1, 2, and 3 the estimator of β for the system of equations in (3) is biased. The bias can be expressed as*

$$\hat{\beta}_{OLS} = \frac{\text{cov}(c_{ia}, x_{ia})}{\text{var}(x_{ia})} = \beta + \underbrace{\rho \frac{\text{cov}(x_{ia}, c_{ib})}{\text{var}(x_{ia})}}_{\text{spillover bias}} + \underbrace{\frac{\text{cov}(x_{ia}, \delta_i)}{\text{var}(x_{ia})}}_{\text{demand bias}} \tag{4}$$

Proof. Looking at the first Equation in the estimated System 3 and comparing it with the real Equation in System 2, we can see that the error term ε_{ia} actually equals to:

$$\varepsilon_{ia} = \delta_i + \rho c_{ib} + \epsilon_{ia}$$

The structural demand system in 2 can be expressed in terms of its reduced form components. In particular, we can express c_{ib} as:

$$c_{ib} = \frac{(1+\rho)}{1-\phi^2-\rho^2} \delta_i + \beta \frac{\rho}{1-\phi^2-\rho^2} x_{ia} + \frac{\rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2}$$

From which $\hat{\beta}_{OLS} = \frac{\text{cov}(c_{ia}, x_{ia})}{\text{var}(x_{ia})} = \beta + \rho \frac{\text{cov}(x_{ia}, c_{ib})}{\text{var}(x_{ia})} + \frac{\text{cov}(x_{ia}, \delta_i)}{\text{var}(x_{ia})}$ derives.

That $\frac{\text{cov}(x_{ia}, c_{ib})}{\text{var}(x_{ia})} = \frac{1+\rho}{1-\rho^2-\phi^2} \frac{\text{cov}(x_{ia}, \delta_i)}{\text{var}(x_{ia})} + \beta \frac{\rho}{1-\rho^2-\phi^2} \neq 0$, $\frac{\text{cov}(x_{ia}, \delta_i)}{\text{var}(x_{ia})} \neq 0$ and that $\frac{\text{cov}(x_{ia}, \delta_i)}{\text{var}(x_{ia})} \neq \rho \frac{\text{cov}(x_{ia}, c_{ib})}{\text{var}(x_{ia})}$ if not for specific values of the parameters conclude the proof. Details are provided in Appendix A.1. ■

The presence of spillovers impacts the estimate **directly** (the *spillover bias* component in Equation (4)). A supply shift induces reallocation of credit demand, as firms demand more from the relationships that became relatively more convenient. Moreover, spillovers affect the bias **indirectly**, interacting with *demand bias*. Demand bias can be at play even in absence of spillovers, but the presence of spillovers can amplify or reduce it, depending on their magnitudes and signs. In this example, the spillover-demand bias interaction component is captured by the role of ϕ and ρ parameters in the $\frac{1+\rho}{1-\rho^2-\phi^2}$ multiplier in front of $\frac{\text{cov}(x_{ia}, \delta_i)}{\text{var}(x_{ia})}$. We can see that, when the absolute value of spillovers is high, the denominator of the interaction component increases, magnifying the demand bias.

In this simplified framework, we can also see how the inclusion of **firm fixed effects** cannot control for credit relationship interdependence, unless we also address the spillovers problem directly. Say we attempt a within estimation, but ignore spillovers represented in System (2). The resulting system is:

$$\begin{aligned} c_{ia} &= \beta x_{ia} + \delta_i + \varepsilon_{ia}, \\ c_{ib} &= \beta x_{ib} + \delta_i + \varepsilon_{ib}, \\ c_{jb} &= \beta x_{jb} + \delta_j + \varepsilon_{jb}. \end{aligned} \tag{5}$$

and the resulting problem is that we will not even address demand bias.

Indicating averages with bars, so that, for example, $\bar{c}_i = \frac{c_{ia} + c_{ib}}{2}$, we can state:

Proposition 2. *Under Assumptions 1, 2 and 3, the estimator of β for the system of equations in (5), the shift in banks' supply curve, is biased and the bias can be expressed as*

$$\begin{aligned} \hat{\beta}_{FE} &= \frac{\text{cov}(c_{ia} - \bar{c}_i, x_{ia} - \bar{x}_i)}{\text{var}(x_{ia} - \bar{x}_i)} \\ &= \beta(1 - \rho) + \rho(1 - \rho) \frac{\text{cov}(c_{ib}, x_{ia})}{\text{var}(x_{ia})} - \rho \frac{\text{cov}(\delta_i, x_{ia})}{\text{var}(x_{ia})} - \phi \frac{\text{cov}(c_{jb}, x_{ia})}{\text{var}(x_{ia})}. \end{aligned} \tag{6}$$

Proof. From Assumption 3 it follows that $\hat{\delta}_i = c_{ib}$. From the structural demand system $\varepsilon_{ia} = \rho c_{ib} + \epsilon_{ia}$ and $\varepsilon_{ib} = \rho c_{ia} + \phi c_{jb} + \epsilon_{ib}$. Then we have that:

$$\begin{aligned} \hat{\beta}_{FE} &= \frac{\text{cov}(c_{ia} - \bar{c}_i, x_{ia} - \bar{x}_i)}{\text{var}(x_{ia} - \bar{x}_i)} = \frac{\text{cov}(c_{ia} - c_{ib}, x_{ia})}{\text{var}(x_{ia})} = \frac{\text{cov}(\beta x_{ia} + \varepsilon_{ia} - \varepsilon_{ib}, x_{ia})}{\text{var}(x_{ia})} = \dots \\ &\dots \beta + \frac{\text{cov}(\rho((1-\rho)c_{ib} - \beta x_{ia} - \delta_i) - \phi c_{jb}, x_{ia})}{\text{var}(x_{ia})} = \dots \\ &\dots \beta(1 - \rho) + \rho(1 - \rho) \frac{\text{cov}(c_{ib}, x_{ia})}{\text{var}(x_{ia})} - \rho \frac{\text{cov}(\delta_i, x_{ia})}{\text{var}(x_{ia})} - \phi \frac{\text{cov}(c_{jb}, x_{ia})}{\text{var}(x_{ia})} \end{aligned} \tag{7}$$

From the above, and the reduced form of System 2, it is evident that β_{FE} is biased, and that correlated demand shocks still play a role, as they are reflected back in the estimator through reallocation spillovers. β_{FE} is indeed a function of δ_i in two ways. First, through the $-\rho \text{cov}(\delta_i, x_{ia}) / \text{var}(x_{ia})$ element, due to demand reallocation within the relationships of the same firm. Second, through the impact of δ_i on *all* other bias components. Details are provided in Appendix A.1. ■

From Proposition 2 and its proof, we can also drive another important conclusion.

Proposition 3. *Under Assumptions 1 and 2, $\hat{\beta}_{FE} \neq \hat{\beta}_{OLS}$ is possible even in the absence of demand bias ($cov(x_{ia}, \delta_i) = 0$).*

Proof. Using Assumption 3 and the absence of demand bias to simplify our calculation, we can thus express the reduced form for c_{ib}, c_{jb} in System (2):

$$\begin{aligned} c_{ib} &= \beta \frac{\rho x_{ia}}{1-\phi^2-\rho^2} + \frac{\rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2} \\ c_{jb} &= \beta \frac{\rho\phi x_{ia}}{1-\phi^2-\rho^2} + \frac{\phi\rho\epsilon_{ia} + (1-\rho^2)\epsilon_{jb} + \phi\epsilon_{ib}}{1-\phi^2-\rho^2} \end{aligned}$$

In the absence of demand bias, the OLS estimator equals:

$$\hat{\beta}_{OLS} = \beta \frac{1-\phi^2}{1-\rho^2-\phi^2} \tag{8}$$

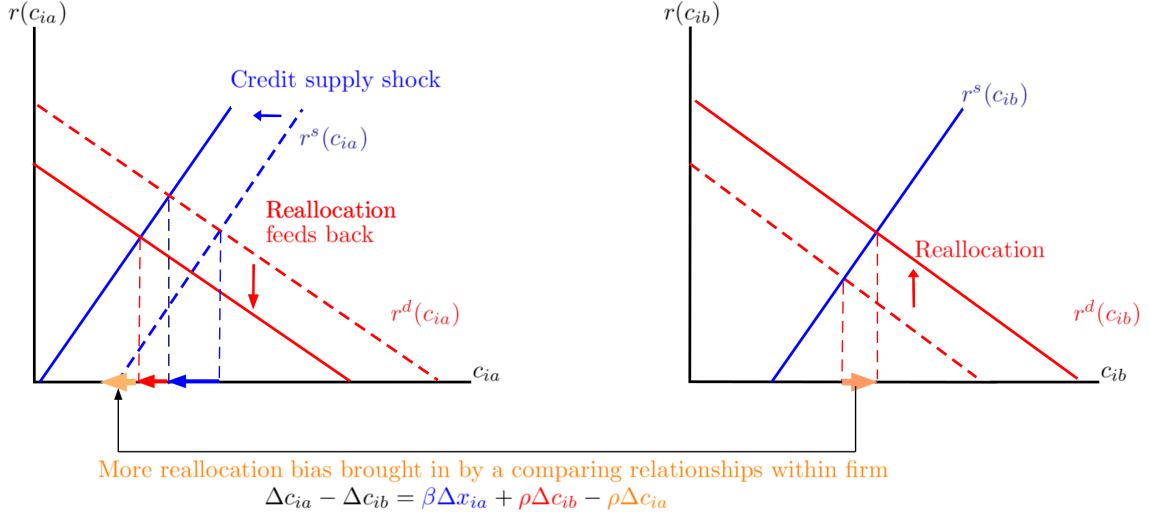
Using the bias expression in Proposition 2, the fixed effect estimator is:

$$\hat{\beta}_{FE} = \beta(1-\rho) + \beta \frac{\rho^2(1-\rho)}{1-\phi^2-\rho^2} - \beta \frac{\phi^2\rho}{1-\phi^2-\rho^2} = \beta \frac{1-\phi^2+\rho}{1-\phi^2-\rho^2}$$

which are different except for specific values of the reduced form parameters. Details are provided in Appendix A.1. ■

This difference implies that interpreting the distance between $\hat{\beta}_{FE}$ and $\hat{\beta}_{OLS}$ as informative on the sign of the demand bias (see, e.g. Jiménez et al., 2020) may, at least in some cases, lead to misguided conclusions. For an intuition, we shall use again our even more simplified graphical framework. In Figure 3, focusing on within-firm reallocation, we show how differentiating within firm in the absence of demand bias is tantamount to add further spillover bias back. This may lead a within-firm assessment of a supply shift even more off than a simple OLS.

Figure 3: Interdependence and Fixed Effects in the Absence of Demand Bias



Notes: The Figure represents graphically how the interaction between firm i and bank a and b , captured by Equation 2, can affect within-firm assessment of supply shifts. The blue lines are banks' credit supply schedule to firm i , while the red lines are firm i 's demand for credit. Solid lines are post-shock curves, while dashed lines are pre-shock curves. Orange segments highlight how within-firm differentiation may worsen our assessment of the supply shift.

Consider again the $\beta \Delta x_{ia}$ supply shift we want to quantify, and suppose we mistakenly believe both c_{ia} and c_{ib} are affected by a demand shock specific to firm i that can bias our quantification of $\beta \Delta x_{ia}$. If we differentiate $\Delta c_{ia} - \Delta c_{ib}$ to address this nonexistent demand confounder, we add to the feedback effect of reallocation on $c_{ia} = \rho \Delta c_{ib}$, the initial reallocation from ia to ib , $\Delta c_{ib} = \rho \Delta c_{ia}$. Within comparison may actually increase the role of reallocation bias in assessing the supply shifts.

Finally, the example highlights an issue with retrieving the bank and firm fixed effects and interpreting them as credit supply and demand shifters (Amiti and Weinstein, 2018).

Proposition 4. *Under Assumptions 1 and 2, firm fixed effects' estimates contain supply shock spillovers and bank fixed effects' estimates may contain demand shock spillovers. As such, they cannot be regarded as pure measures of each firm or bank demand and supply shocks, respectively.*

Proof. We start considering an alternative version of Assumption 3, that allows for $\delta_j \neq 0$. Then, we notice that if the econometrician tries and estimate System (5), then she obtains:

$$\begin{aligned} \hat{\delta}_i &= \frac{(1+\rho)}{1-\phi^2-\rho^2} \delta_i + \frac{\phi}{1-\phi^2-\rho^2} \delta_j \\ \hat{\delta}_j &= \frac{\phi(1+\rho)}{1-\phi^2-\rho^2} \delta_i + \frac{(1-\rho^2)}{1-\phi^2-\rho^2} \delta_j \end{aligned} \quad (9)$$

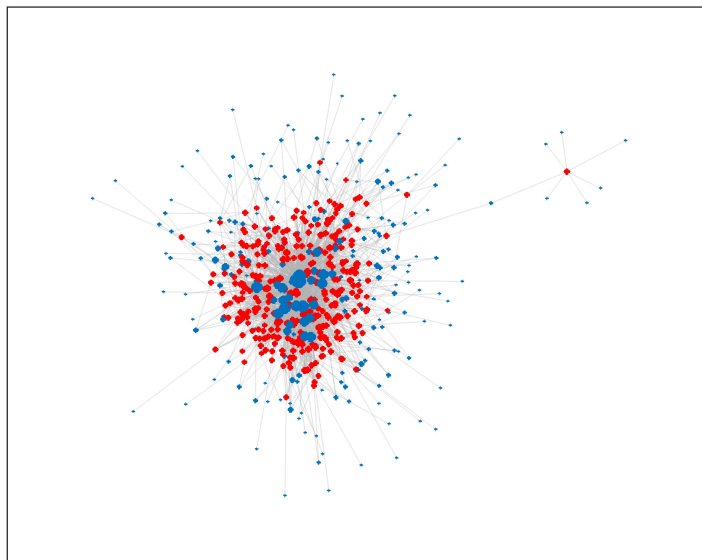
from the the solution of System (2). ■

Even in this simple setting, where we focus on firms fixed effects only, we can see that fixed effect estimate in (9) are already affected by two problems. First, fixed effects

do not capture the pure demand shock exactly, but an amplified or attenuated version of it (on the base of the signs and relative magnitudes of ϕ, ρ). Focusing on $\hat{\delta}_i$, this is exemplified by the $\frac{(1+\rho)}{1-\phi^2-\rho^2}\delta_i$ element. Second, bank links reflects other firms' shocks back in the estimate, further biasing the fixed effect estimates. Focusing again on $\hat{\delta}_i$, this is exemplified by the $\frac{\phi}{1-\phi^2-\rho^2}\delta_j$ element.

This example is useful to get an intuition of the basic mechanics at play when banks and firms optimize their portfolios in an integrated way, and the simple yet powerful results that follow if these forces are ignored and standard estimation is performed. Nevertheless, actual credit networks can be much more complex, involving thousands of firms and hundreds of banks. Figure 4 provides a visual example of how complex the network formed by real credit relationships could be from a sample of only 500 real credit relationships from the Italian Credit Register data.

Figure 4: A Sampled Real Credit Network



Notes: The network is derived from a sample of 500 credit relationships observed in 2012. Banks are represented in blue and firms in red. The estimated network is represented with a force-directed layout with five iterations. It uses attractive forces between adjacent nodes and repulsive forces between distant nodes. To ease the visualization, the size of the nodes is equal to the (log) of their degree. See Fruchterman and Reingold (1991) for more details.

As a consequence, the resulting signs and magnitudes of the biases are more difficult to derive. First, as we can already see from the simple example, if the shock hit relationship ib instead of relationship ia , all bias expressions would change. This is actually an instance of a standard result in the network literature, finding that the extent of spillovers is mediated by each node's location in the network (Ballester, Calvó-Armengol, and Zenou, 2006). Furthermore, as the number of links increases, the complexity of the feedback

effects increases too. A network with three links incorporates feedback loops of order two at most, i.e. a shock can affect another link and this movement can come back hitting the initial node. A network of n links can exhibit much more complicated dynamics.

Fortunately though, the intuition we built in this Section carries through to contexts of greater complexity, where existence and uniqueness of solution has been proven under mild conditions in Ballester, Calvó-Armengol, and Zenou (2006). In the rest of the document, we will introduce and estimate a generalized network model of credit relationship between F firms and B banks.

3 The Econometric Framework

In this section we introduce an econometric framework encompassing spillovers among credit relationships between banks and firms. The framework adds in four ways to the toolkit currently available to the empirical researcher in banking. First, it allows us to identify and measure spillovers through banks and firms. Second, it can be used to consistently estimate direct and indirect effects of treatments and shocks to firms and banks outcomes without imposing strong independence assumptions. Third, such estimates allow to measure and understand the deep links between shocks' effects and the credit market structure, especially how the centrality of banks and firms in the network of relationships affects the passthrough. Finally, as we derive the econometric framework from the simple microfoundation of banks and firms behavior discussed in Section 2.1, the estimates map with salient structural parameters that have an explicit economic interpretation.

3.1 The Credit Network Model

Suppose that there are two sets, \mathbb{F} and \mathbb{B} , of firms and banks in the market with cardinality respectively equal to F and B . We can easily generalize System (2) in Section 2.1 to any number of banks, firms and relationships in the credit network as:

$$c_{ib} = \alpha + \phi \sum_{j \in \mathbb{F} \setminus i} a_{ib,jb} c_{jb} + \rho \sum_{k \in \mathbb{B} \setminus b} a_{ib,ik} c_{ik} + \delta_i + \gamma_b + x_{ib} \beta + \epsilon_{ib}, \quad (10)$$

where c_{ib} is credit from bank b to firm i . δ_i and γ_b are the firm and bank fixed effects. x_{ib} is a vector of exogenous characteristics of the loan, which may include a specific treatment administered to the relationship ib . ϵ_{ib} is the error component. The term $a_{ib,jb}$ captures the connections among credit relationships by the lender side, being equal to one if both i and j borrow from b . The term $a_{ib,ik}$ captures the connections among credit relationships

by the borrower side, being equal to one if both b and k lend to i .¹³

In this model, credit relationships are not i.i.d., credit granted bilaterally from banks to firms is jointly determined. On the one hand, the amount of credit from bank b to firm i depends on the credit that bank b gives to other firms j , as firms compete on the demand side to get credit from bank b , which is budget constrained. It captures the effect that a change in c_{jb} has on c_{ib} . We call this effect, captured by ϕ , the *bank supply allocation effect* (BSA), as it captures shifts in the supply between credit relationship involving the same bank that are driven by its reallocation policies. On the other hand, the amount of credit from bank b to firm i depends on the credit that firm i takes from other banks, as banks compete on the supply side to grant credit to firm i , whose demand is not unlimited. It captures the effect that a change in c_{ik} has on c_{ib} . We call this effect, captured by ρ , the *firm demand allocation effect* (FDA), as it captures shifts in the demand between credit relationship involving the same firm that are driven by its substitution choices.

Consider again System (2) in Section 2.1. The credit relationship ib is influenced by ia as both share the same borrower, and by jb as they share the same lender. The relationship ia is not influenced directly by jb , because it does not share any counterparty with it. Nevertheless, as we discuss in more detail below, ia is indirectly exposed to jb through adjustments in ib , as highlighted in Section 2.1. The matrix form of the *credit network model* (CNM) is:

$$\begin{aligned} C &= \alpha + \phi A_B C + \rho A_F C + X\beta + \Delta + \Gamma + \epsilon, \\ &= \phi A_B C + \rho A_F C + Z\mu + \epsilon. \end{aligned} \tag{11}$$

where X is the matrix of loans covariates. Δ is the matrix containing the firm fixed effects. Γ is the matrix containing the banks fixed effects. C is the vector containing all the N credit relationships between banks and firms in the market. A_B is the $(N \times N)$ adjacency matrix of the network that keeps track of connections among loans through banks whose generic element $a_{ib,jk}$ is equal to one iff $b = k$. A_F is the $(N \times N)$ adjacency matrix of the network that keeps track of connections among loans through firms whose generic element $a_{ib,jk}$ is equal to one iff $i = j$. For both adjacency matrices, we let $a_{ib,ib} = 0$ for all ib , following convention. The vector $A_B C$ contains for each loan the amount of credit granted by the same bank to other firms. The vector $A_F C$ contains for each loan the amount of credit obtained by the same firm from other banks.

We define the *isolated-credit model* (ICM):

$$C = \alpha + X\beta + \Delta + \Gamma + \epsilon, \tag{12}$$

¹³ Given that we want to consider the total credit, we do not row-normalize these terms as sometimes is done in network econometrics, see Liu, Patacchini, and Zenou (2014) for a discussion.

a model in which the credit relationships are forced to be independent, i.e. if we impose the restriction $\phi = \rho = 0$ in Equation (11). It is worth observing that (i) the CNM nests the standard ICMs commonly used in the literature, and (ii) it exploits exactly the same information set of the ICM, because the network structure (summarized by A_B and A_F) comes directly from the observable firms and banks' identities. From this perspective, our model can be used by every researcher working with credit register data.

The matrix form of the model makes it clearer that we deal with a simultaneous system of equations, in which the credit vector C enters the equation both on the left and the right hand side, through $A_B C$ and $A_F C$, the endogenous terms. This feature captures the more realistic assumption that credit choices are not independent, but comes at the cost of additional complexity in the econometric model and its identification. It can not be estimated by simple OLS. Nevertheless, model (11) belongs to the spatial autoregressive (SAR) models class, thus we can exploit some key results in this literature, especially the branch of the literature that extended these models to the analysis of networks (see Arduini, Patacchini, and Rainone, 2020; Bramoullé, Djebbari, and Fortin, 2009; Hsieh and Lee, 2016; Johnsson and Moon, 2021; Lee, 2007; Lee, Liu, and Lin, 2010; Patacchini, Rainone, and Zenou, 2017, among others). However, there are peculiarities and problems in our framework that deserve discussion and could need tailored solutions. For example, standard SAR models usually consider outcomes at the node level, while in our case outcomes are at the link level. In addition, nodes here belong to two different types of agents and links can be formed only between the two types, not within, and we have multiple endogenous terms and parameters.

3.2 Identification

The main issue that arises when we want to estimate equation (11) is the endogeneity of $A_B C$ and $A_F C$, negating the consistency of OLS estimation. The simultaneity of equations in model (11) creates an intrinsic endogeneity problem if

$$\begin{aligned} E[(A_F C)' \epsilon] &= E[(A_F (I - \phi A_F - \rho A_B)^{-1} (\alpha + Z\mu + \epsilon))' \epsilon] \neq 0, \\ E[(A_B C)' \epsilon] &= E[(A_B (I - \phi A_F - \rho A_B)^{-1} (\alpha + Z\mu + \epsilon))' \epsilon] \neq 0. \end{aligned}$$

The last inequalities hold if

$$\begin{aligned} E[(A_F (I - \phi A_F - \rho A_B)^{-1} \epsilon)' \epsilon] &= \sigma_\epsilon^2 \text{tr}(A_F (I - \phi A_F - \rho A_B)^{-1}) \neq 0, \\ E[(A_B (I - \phi A_F - \rho A_B)^{-1} \epsilon)' \epsilon] &= \sigma_\epsilon^2 \text{tr}(A_B (I - \phi A_F - \rho A_B)^{-1}) \neq 0, \end{aligned}$$

where tr is the matrix trace operator. Endogeneity is basically determined by the structure of the observed network, represented by A_F and A_B .

In SAR models, spatial lags of both the endogenous (here $A_B C$ and $A_F C$) and ex-

ogenous variables (here $A_B X$ and $A_F X$) can be included on the right hand side. The social networks literature calls the latter contextual effects, as they capture the direct influence of peers' characteristics. While endogenous effects arise quite naturally in our context, and we explicitly discussed them in Section 2, exogenous effects are plausible too. For just one example, a reversal in the fortunes of some upstream and downstream firms may impact all firms sharing the same value chain. Outcomes of such firms will thus react directly to these shocks. Similar contextual effects have been recently studied by works such as Berg and Streitz (2019); Huber (2018). Here, we will focus on our unique contribution, the handling of direct effects. Nonetheless, we stress that our model can be extended to accommodate contextual effects, if needed.

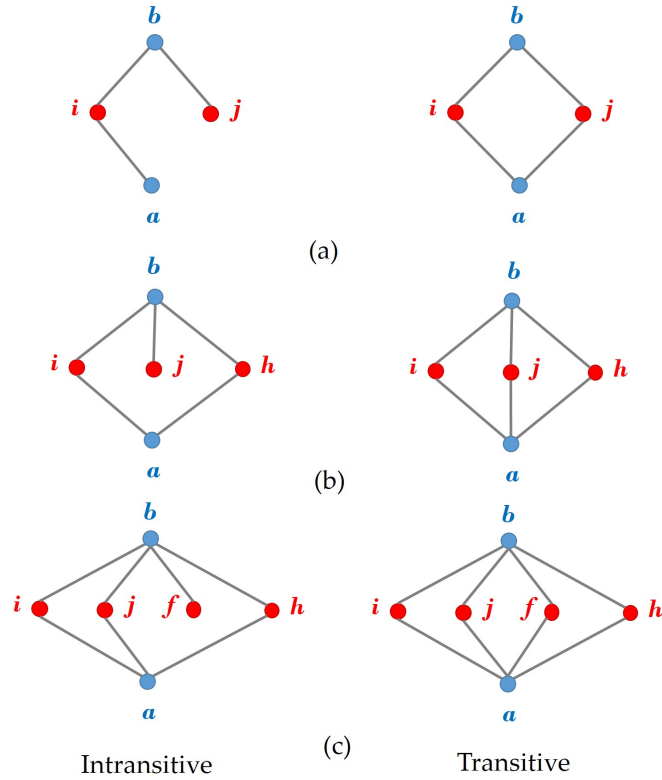
The following proposition establishes sufficient conditions under which the parameters in model (11) are identified, even if augmented with exogenous effects $A_B X$ and $A_F X$.

Proposition 5. (*Identification of the Credit Network Model*). *The credit network model in (11) is identified if I_F , $A_B A_F A_B$ and A_F are linearly independent and I_B , $A_F A_B A_F$ and A_B are linearly independent -i.e. there are intransitive quadriads in the credit network- and $\phi\beta \neq 0$ and $\rho\beta \neq 0$.*

A nice feature of this result is that it translates into the easy-to-check requirement that banks do not have fully overlapping portfolios. In other words, it requires that not all the banks lend to the same set of firms. It follows that the credit market structure itself can provide the solution to the endogeneity problem, if it meets certain conditions. Precisely, we need the market structure to show a certain degree of *intransitiveness*. The level of intransitivity is the ratio between the number of intransitive quadriads and the total number of quadriads. A quadriad is a set of four nodes, two banks and two firms. The quadriad is transitive if all the between-type links are realized. For identification purposes, the market must not be composed only by transitive quadriads. The presence of intransitive triads is a sufficient (but not necessary) condition for the identification of the model's parameters.

The intuition is that intransitivity provides exclusion restrictions that allow us to identify the system of simultaneous equations in (11). Figure 5 provides examples of market structures that allow (networks on the left) and do not allow (networks on the right) for identification of spillovers among credit relationships, with two (panel (a)), three (panel (b)) and four (panel (c)) firms in the market. Let us consider the simplest networks in panel (a). In the left network, the fact that bank k does not have a relationship with firm j allows jb to be excluded from ik 's equation, because j is not connected with k . It follows that jb can be used as an instrument to estimate the effect of ib on ik , as it has a direct impact on the former but not on the latter (which in turn it influences only through the former).

Figure 5: Network Structure and Identification



Notes: Nodes a, b are banks, red nodes are firms, edges are credit relationships.

Intuitively, when the number of firms grows, the number of intransitive quadriads has to grow as well. In the credit market, a transitive quadriad appears when a bank b lends to a different set of firms w.r.t. another bank k , or specularly when a firm i borrows from a different set of banks w.r.t. another firm j . If the market is composed only by transitive quadriads, we cannot identify the parameters in the system, there is no valid exclusion restriction. This situation is extremely rare credit markets.

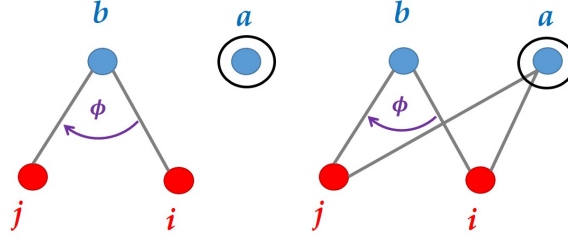
3.3 Overlapping Portfolios Instrumental Variables

In this section, we propose a new method to construct instrumental variables based on overlapping portfolios (OPIVs). In credit markets, firms can borrow from multiple lenders. This feature creates overlapping portfolios of relationships both at the firm and bank level. As idiosyncratic shocks to one relationship can affect the others involving the same parties, they can provide relevant instruments for the FDA and the BSA. If these portfolios are not fully overlapping, such instruments are also exogenous.

Intuition. Let us make a simple example in Figure 6. In panel (a), the credit network is disconnected, bank a not lending at all. In panel (b), we have full overlap. Everybody lends/borrows from everybody. Panel (c), instead, respects all conditions for identification. For simplicity, let's focus on how a shock to bank a helps us identify ϕ . In

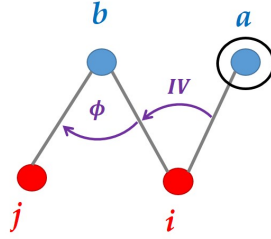
panel (a), a shock to bank a is not relevant, as the two banks are not connected through any firm. In panel (b), the same shock is not exogenous, because it directly influences jb through ja . In panel (c), a shock to bank a is a valid instrument, instead. Indeed, it is relevant, as it directly influences bi through ai ; moreover, it is exogenous to jb , as the two relationships are only indirectly connected. The same reasoning applies to shocks that are firm-specific or relationship-specific, for the identification of ρ .

Figure 6: Exogeneity and Relevance of OPIVs



(a) Not relevant

(b) Not exogenous



(c) Relevant and exogenous

Notes: Nodes a, b are banks, i, j are firms, edges are credit relationships.

Let's consider again the example we introduced in Section 2.1, in its most basic form, where there is no correlated demand confounder ($\frac{\text{cov}(x_{ia}, \delta_i)}{\text{var}(x_{ia})} = 0$) and the following assumption holds:

Assumption 4. x_{jb} and x_{ia} do not affect directly c_{ib} and are uncorrelated with ϵ_{ia} and ϵ_{jb} .

Then we have:

Proposition 6. Under Assumptions 1, 2, 4, and the network structure in Section 2.1, we can identify the spillover parameters (ϕ and ρ) with a 2SLS procedure, and deliver an unbiased estimate of β .

Proof. The System in 2 can be rearranged as

$$\begin{aligned}
 c_{ia} &= \rho\pi_\rho x_{jb} + \left(\beta + \frac{\beta\rho^2}{1-\phi^2-\rho^2}\right)x_{ia} + \rho\mu + \epsilon_{ia} \\
 c_{ib} &= \pi_\rho x_{jb} + \pi_\phi x_{ia} + \mu \\
 c_{jb} &= \phi\pi_\phi x_{ia} + \left(\frac{\phi^2\beta}{1-\phi^2-\rho^2} + \beta\right)x_{jb} + \phi\mu + \epsilon_{jb}
 \end{aligned} \tag{13}$$

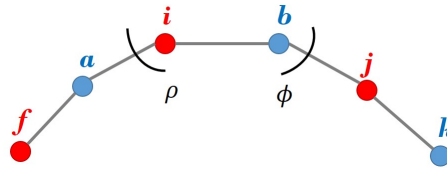
where π_ρ and π_ϕ are the reduced form parameters for the instruments' effect on c_{ib} and μ summarizes all other parameters we are not interested into. We use the c_{ib} equation of the System in 13 as first stage, recovering $\hat{\pi}_{\rho,OLS}, \hat{\pi}_{\phi,OLS}$, which in turn we use to deflate $\widehat{\pi}_{\rho} \widehat{\rho}_{OLS}, \widehat{\pi}_{\phi} \widehat{\phi}_{OLS}$ in the second stage, to finally obtain $\hat{\rho}_{IV} = \rho, \hat{\phi}_{IV} = \phi$. Endowed with unbiased estimates of the spillovers through bank and firms' parameters, we can correct the $\hat{\beta}_{OLS}$ and derive an unbiased $\hat{\beta}_{IV}$. Details in Appendix A.1. ■

The key insight comes from the fact that, under Assumption 4, credit relationship jb (ia) provides exogenous variation through x_{jb} (x_{ia}) that does not affect directly c_{ib} , but does affect it indirectly through c_{jb} (c_{ia}). It then allows us to identify ϕ (ρ) with a 2SLS estimator. Identification of ρ and ϕ allows us, in turn, to retrieve an unbiased estimator of β .

Depending on the circumstances, Assumption 4 may be too strong, for example because there are no credible pairwise observables (for example x_{jb}) that do not affect c_{ib} . An example could be that the econometrician does not observe other pairwise or firm specific variables at all, and only have bank specific variables at hand. If only firm-specific variables are available, j -specific variables (say x_j) can be used as an instrument for c_{jb} . If only bank-specific variables are available, b specific variables (say x_b) can not be used as an instrument, because they affect directly c_{ib} .

In such cases, we can use variation in more distant credit relationships for identification. For instance, let us add two other credit relationships to our toy example and assume that bank a supplies credit also to firm f and firm j demands credit also to bank k . Figure 7 provides the relative graph. In this highly intransitive credit network, variation induced by fa (jk) to ia (jb) can be used as an instrument to identify ρ (ϕ) -i.e. the effect of c_{ia} (c_{jb}) on c_{ib} -. Here, even if the econometrician only observe banks' characteristics, a (k) specific variables can be used as an instrument for c_{ia} (c_{jb}) and identify ϕ (ρ). A specular strategy can be used if the econometrician only observe firms' characteristics, say f (j).

Figure 7: IV exclusion restrictions



Notes: nodes a, b, k are banks, f, i, j are firms, edges are credit relationships.

In the full model. Following Proposition 5, OPIVs are a natural way to proceed with the estimation of equation (11). The OPIVs are substantially “network embedded”; in other words, we can use the network topology to create IVs that are correlated with

the endogenous variables, but independent from the error term.¹⁴ Indeed, the expected values of the two endogenous variables, $E(A_FC)$ and $E(A_BC)$, meet these two conditions. Taking advantage of the reduced form, we thus derive the theoretically best OPIVs as

$$TIV_F = E(A_FC) = E[(A_F(I - \phi A_F - \rho A_B)^{-1}(\alpha + Z\mu))], \quad (14)$$

$$TIV_B = E(A_BC) = E[(A_B(I - \phi A_F - \rho A_B)^{-1}(\alpha + Z\mu))], \quad (15)$$

where we can get to this conclusion as $E[(I - \phi A_F - \rho A_B)^{-1}\epsilon] = 0$.

Given that the parameters in Equations (14)-(15) are unknown, TIV_B and TIV_F are unfeasible. Assuming $\|\phi A_F + \rho A_B\|_\infty < 1$,¹⁵ the term $(I - \phi A_F - \rho A_B)^{-1}$ is an infinite sum of elements $\sum_{k=0}^{\infty} (\phi A_F + \rho A_B)^k$. A linear approximation, EIV_F and EIV_B , of vectors appearing in equation (14) and (15) can thus be used for the empirical IV. Exploiting only the variation of covariates, X , first order approximations of TIV_F and TIV_B are respectively:

$$EIV_F^1 = A_F X, \quad (16)$$

$$EIV_B^1 = A_B X, \quad (17)$$

second order approximations are

$$EIV_F^2 = [A_F, A_F A_B] X, \quad (18)$$

$$EIV_B^2 = [A_B, A_B A_F] X. \quad (19)$$

and so on and so forth.¹⁶ Observe that here $A_F^k \equiv A_B^k \equiv \mathbf{0}$, which is the zero matrix, for $k > 1$, as we have a bipartite network with only cross types links. Let us use again our example in Figure 7, in order to consistently estimate ρ (ϕ), a first order approximation uses x_{ia} (x_{jb}) as an IV. A second order approximation would use x_{fa} (x_{jk}) in addition in

¹⁴ The literature of spatial and network econometrics investigated in depth several methods to treat the endogeneity created by these simultaneous equations, Kelejian and Prucha (1999) and Liu and Lee (2010) proposed a GMM approach, and Lee (2004) used a Quasi-Maximum Likelihood Estimator. In this paper we use an IV approach in the spirit of Lee, Liu, and Lin (2010), Lee (2007) and Kelejian and Prucha (1998).

¹⁵ This is a sufficient condition for the invertibility of $(I - \phi A_F - \rho A_B)$; it also determines the parameter space for spillover effects.

¹⁶ The approximation is as follows. $E(A_FC) = E[(A_F(I - \phi A_F - \rho A_B)^{-1}(\alpha + Z\mu))] = E[A_F[\sum_{k=0}^{\infty} (\phi A_F + \rho A_B)^k](\alpha + X\beta + \Delta + \Gamma)] = E[A_F X\beta] + E[A_F A_B X\beta] + E[A_F(\alpha + \Delta + \Gamma)] + E[A_F A_B(\alpha + \Delta + \Gamma)] + E[A_F[\sum_{k=2}^{\infty} (\phi A_F + \rho A_B)^k](\alpha + X\beta + \Delta + \Gamma)]$. This is due from the fact that we have a bipartite network without within type connections and thus $A_F A_F = A_B A_B = \mathbf{0}$. A specular approximation can be derived for TIV_B .

the IV.¹⁷ The 2SLS estimator is consequently

$$\hat{\theta}_{2SLS} = (W'P_QW)^{-1}(W'P_QC), \quad (20)$$

where $Z = [A_FC, A_BC, Z]$, $P_Q = Q(Q'Q)^{-1}Q'$, $Q = [EIV_F, EIV_B, X]$ and $\hat{\theta}_{m,t,2SLS} = [\hat{\phi}_{2SLS}, \hat{\rho}_{2SLS}, \hat{\mu}_{2SLS}]$. Given that we do not employ Γ and Δ in the construction of the OPIVs, the number of IVs does not grow with the number of banks and firms in the sample. It follows that the main asymptotic properties of the estimator are standard and follow those of the ICM.

Intuitively, the estimator is consistent as long as the number of credit relationships grows faster than the number of banks and firms in the market, as it allows to include the FEs in the model.¹⁸ Using Liu and Lee (2010)'s terminology, if we wanted to use a more efficient "many IVs" strategy, instead of a "few IVs" strategy we adopt, we could have employed also Γ and Δ in the OPIVs. Nonetheless, in case, we should have derived the asymptotic properties of such estimator to develop a bias-correction procedure. In this paper, our main objects of interest are the identification, the finite sample properties, and the empirical analysis of this few IVs estimator. We leave the analysis of the asymptotic performance of a more efficient many IVs estimator for future research.

In some cases, the dimension of credit register or, more generally, bilateral credit relationship data can be very high, eventually millions of observations. In such situations, the curse of dimensionality can severely constrain or eventually prevent the computation of the estimator in (20), because some of the matrices involved can not be manipulated with standard software. We can address this problem using a within estimator. Let $F = [\Delta, \Gamma]$ and $J = I - F(F'F)^{-1}F'$. Then we have

$$JY = \phi JA_BY + \rho JA_FY + JX\beta + J\epsilon.$$

One can thus first estimate $\omega = (\phi, \rho, \beta)'$ by 2SLS.

Let $R = [A_BY, A_FY, X]$ and $P_H = JH(H'JH)^{-1}H'J$ where H is a matrix of IVs of linearly independent columns of $[X, A_BX, A_FX]$.¹⁹ Then

$$\hat{\omega}_{2SLSW} = (R'P_HZ)^{-1}R'P_HY, \quad (21)$$

¹⁷ Observe that using X in the empirical IV corresponds to the 'few IV' estimator strategy in Liu and Lee (2010), we abstract from efficiency considerations and bias correction issues that would emerge from the analog of the 'many IV' estimator strategy using Z instead.

¹⁸ We abstract here from the potential consequences of violating this assumption and selecting only firms that borrow from multiple banks in the sample, as it is a standard practice in this literature, which could affect both the ICM and CNM. Usually this operation is justified by observing that firms with only one relationship account for a very small share of the market.

¹⁹ Observe that if pairwise exogenous X are observable in the data and exogenous effects are not considered, a first order approximation is the preferred solution if the curse of dimensionality is binding, because products of high dimensional matrices such as A_B and A_F . Unless differently specified we will use such approximation in what follows.

is a consistent estimator for the spillover effects. Next, one can estimate $\phi = (\gamma', \delta')$ by OLS. Let $\hat{U} = Y - R\hat{\omega}_{2SLSW}$. Then,

$$\hat{\phi}_{2SLSW} = (F'F)^{-1}F'\hat{U}. \quad (22)$$

3.4 Bias of ICM under Interdependent Credit Relationships

Here, we derive analytical results on the sign and magnitude of treatment effects and idiosyncratic shocks' biases when the DGP is a CNM and the econometrician estimates a ICM. Which is, what happens when we ignore the network nature of credit relationships?

3.4.1 Bias of Treatment Effects

Let us first abstract from the presence of fixed effect and let the DGP be

$$C = \phi A_B C + \rho A_F C + X\beta + \epsilon, \quad (23)$$

For simplicity, let also $\rho = \phi$. Suppose we estimate

$$C = X\beta + U, \quad (24)$$

then the error term has the following form

$$\begin{aligned} U &= \phi A_B C + \rho A_F C + \epsilon = (\phi A_B + \rho A_F)(I - \phi A_B - \rho A_F)^{-1}[X\beta + \epsilon] + \epsilon \\ &= MX\beta + (M + I)\epsilon. \end{aligned} \quad (25)$$

Suppose that X is univariate and $X \perp \epsilon$,²⁰ we have

$$\begin{aligned} X'U &= X'MX\beta + X'(M + I)\epsilon = X'\phi A \sum_{k=0}^{\infty} (\phi A)^k X\beta \\ &= X' \sum_{k=1}^{\infty} (\phi A)^k X\beta = \beta \sum_{k=1}^{\infty} \phi^k X' A^k X = S. \end{aligned} \quad (26)$$

If $\beta, \phi > 0$ then $B = \hat{\beta} - \beta = (X'X)^{-1}S > 0$. The positive bias is given by the amplification generated by spillovers, which is not disentangled in the reduced form estimate. If $\phi < 0$ then the sign of B depends on A , i.e. the network structure, and the intensity of the spillovers, as it contains decaying functions of ϕ .

Suppose X is binary. For example, a stark change in regulation hits only part of our credit relationships population. Then, for each treated relationship, we call $X' A^k X = p_k$

²⁰To ease the notation we assume independence here, the same conclusions can be reached assuming $E[\epsilon'X] = 0$.

the sum of the number of k -distant treated credit relationships. It follows that

$$S = \beta \sum_{k=1}^{\infty} \phi^k p_k = \beta \left(\sum_{k \text{ odd}} \phi^k p_k + \sum_{k \text{ even}} \phi^k p_k \right) = \beta(OD + EV).$$

Given that OD is negative and EV is positive, $B > 0$ if $OD > -EV$. The intuition behind the indeterminate sign is that, when the spillovers are negative, they lower the outcomes of odd-distant agents (at distance 1, 3, 5, etc) but increase the outcomes of even-distant agents (at distance 2, 4, etc). Nevertheless, the negative sign is likely to prevail as the first round effects have a higher weight, because $\phi < 0$, especially in denser networks. The inclusion of fixed effects, whether they are correlated or not with treatment, does not change these last results. We show the details in Appendix A.1

3.4.2 Bias with Spillovers and Endogenous Treatments

In the previous analysis, we assumed that the main regressor is exogenous. Let us now allow X to be an endogenous regressor. In credit markets, as in other fields in which experiments are not possible or easy to implement, regressors are often endogenous. Endogeneity can arise because of self-selection in the extensive (see Jiménez et al., 2014), or in the intensive margin (see Paravisini, Rappoport, and Schnabl, 2023). Also, it can arise due to the omission of relevant variables on the RHS. In some cases, a credible instrument (a selection step, or a clever IV) can be found, but in many situations this is not a possibility and the omission of relevant, and eventually endogenous, variables is always difficult to assess in practice.

In what follows, we want to understand analytically the consequences of endogeneity for the bias due to spillovers. For simplicity, we still focus on $\phi = \rho$ and we assume fixed effects away. At difference with the previous sections, we assume that $\epsilon = \iota X + V$, with V being an error term such that $V \perp X, \epsilon$. Suppose we estimate the following CNM's and ICM's parameters

$$C = \phi A_B C + \rho A_F C + X\beta + \epsilon, \quad (27)$$

$$C = X\beta + U, \quad (28)$$

When the DGP comes from Equation 27, the error term of Equation 28 has the following form

$$\begin{aligned} U &= \phi A_B C + \rho A_F C + \iota X + V \\ &= (\phi A_B + \rho A_F)(I - \phi A_B - \rho A_F)^{-1}[X\beta + \iota X + V] + \iota X + V, \end{aligned}$$

then, given that $V \perp X$, we have

$$\begin{aligned} X'U &= S + X'(M + I)(\iota X + V) \\ &= \underbrace{S}_{\text{spillovers}} + \underbrace{\iota X'X}_{\text{endogeneity}} + \underbrace{\iota X'MX}_{\text{combination}}, \end{aligned}$$

$$X'\epsilon = \iota X'X + \iota X'MX.$$

In this case, the estimate of β is biased for *both* Equations (27) and (28). Nonetheless, the bias of (28), B_{ICM} , has three drivers: The pure *spillover* component; the pure *endogeneity* component; the *combination* of the two. On the other hand, the bias of (27), B_{CNM} , is only affected by the last two. It follows that even in a context rid with endogeneity, we can use our model to study firms and banks credit reallocation, system-wide. More formally $D = B_{ICM} - B_{CNM} \neq 0$ does not imply that $\iota \neq 0$, but it does imply that $S \neq 0$, and thus D can inform about the presence of (and the bias induced by) spillovers even if the estimate of β is biased.

We get this result because unbiased spillovers can be recovered even in the presence of treatment endogeneity: the network lags used as instrumental variables for the identification of the spillovers, for example the EIV in Equations (16) and (17), are still uncorrelated with the error term, i.e. $E[\epsilon'AX] = 0$. The intuition is that the treatments to other agents in the economy can still be valid instruments for their outcomes, even if they are endogenous.²¹ The key requirement is that they are not endogenous to their network lags.²²

3.4.3 Bias of Idiosyncratic Firm and Bank Shocks

Again, we find that estimating and retrieving banks and firms fixed effects we face an indeterminate bias sign. The intuition behind this result is similar to the one for the treatment effect bias' indeterminacy: Portfolio wide optimization and higher order loops let idiosyncratic shocks diffuse through the credit network. Propagation may stress different shocks reflecting throught the network, giving rise to oscillating sums and indeterminate biases. We provide the details in Appendix A.1.

²¹ Such result would not stand if the order of the *EIV* used to approximate the *TIV* is higher, because $X'A^kX$ could contain powers of the same x_i when $k \geq 2$. Using a first order approximation helps to avoid it.

²² Simulation in Section 4 confirm the result, while we show that our method can easily accommodate a instrumental variable strategy in Section 6.

4 Monte Carlo Study

In this section, we simulate networks to illustrate our method’s validity and its properties, in finite samples under different settings, using both simulated and real credit (sampled from the credit register) networks. Furthermore, we use these numerical experiments to study the bias of the isolated-credit network estimator. The key take away from our simulation is that the sign and size of the bias from ignoring the network structure cannot be derived *ex-ante*. Indeed, the bias is a convolution of observables, such as the market structure (in particular its density) and the number of treated units, and unobservables, such as the structural parameters determining the spillovers’ magnitude. Thus, we can experience high non linearity of the bias, sign uncertainty, and we need to tackle spillovers directly.

4.1 Setting

We use a given set of parameters ϕ , ρ and β , randomly generated characteristics X , error terms ϵ , and a network G as inputs. We generate G as a ‘circular network’, ordering nodes according to natural numbers from 1 to N . If node i is odd, it is a bank, if it is even, it is a firm. We link node i to all opposite type nodes, till node $i + j$ for $j \leq z_i$, where z_i is an independent realization from a uniform distribution $U(0, m)$ for each node i .²³ For example, if z_1 (z_2) = 10, bank 1 (firm 2) links with firms 2,4,6,8 and 10 (banks 3,5,7,9 and 11). In this way, there are only links between banks and firms, none within the two types.

We label this experiment the circular network, because, when $z_i = 1$ for all i , nodes’ connections are one to one (banks can share one borrower at most), with the last reconnecting to the first, describing a circle. The model we use is indeed a generalization of a simple circle that allows for a random number of connections. The variable m defines network density. When m increases, we are increasing the average number of shared borrowers, as well as the number of banks between which they are shared. This setup is useful for the comparative statics exercise, because it allows us to easily change features (size, sparsity, etc.).

In our benchmark simulation, we generate 500 networks G with $n = 200$ nodes in each, 100 banks and 100 firms. Once G is randomly generated, we then derive the links’ adjacency matrix A from G . For each link we create an observable (x) and an unobservable (ϵ) variable. x is a dummy equal to 1 for a certain share s of the population and zero otherwise, tracking whether the credit relationship received the shock. We extract ϵ from a normal distribution with mean equal to zero and variance equal to σ . We generate lender and borrower fixed effects from independent normal distributions

²³ We have also repeated our simulation experiment using other distributions for z_i (different from uniform). The results are not sensitive to different specifications. Such results are available upon request.

$\theta N(0, 1)$ draws, constraining them to be positive by adding the absolute value of the minimum draw. We thus generate credit relationships in reduced form from: $C = (I - \phi A_B - \rho A_F)^{-1}(\beta X + \Delta + \Gamma + \epsilon)$. In particular, our pivotal setting has $(\beta, N, \sigma, \delta) = (-2, 200, 1, 0.1)$ and $R = 500$ simulated samples.

We perform two classes of exercises within this framework. First, we look at the bias of $\hat{\beta}$ when we ignore the interdependence among credit relationships. Second, we study the performance of our estimators for the treatment effect β , the spillovers ϕ , ρ , and firms and banks' fixed effects in finite samples, under different settings and assumptions. Across all these exercises, on top and above negative treatment effect, we also assume negative spillovers. Doing so, we focus on a framework resembling a typical negative shock hitting the credit supply of some banks to some firms (e.g. a monetary policy contraction, to which some banks are more sensitive than others), where these firms and banks can reallocate credit freely across relationships in their portfolio.

4.2 Results with Simulated Networks

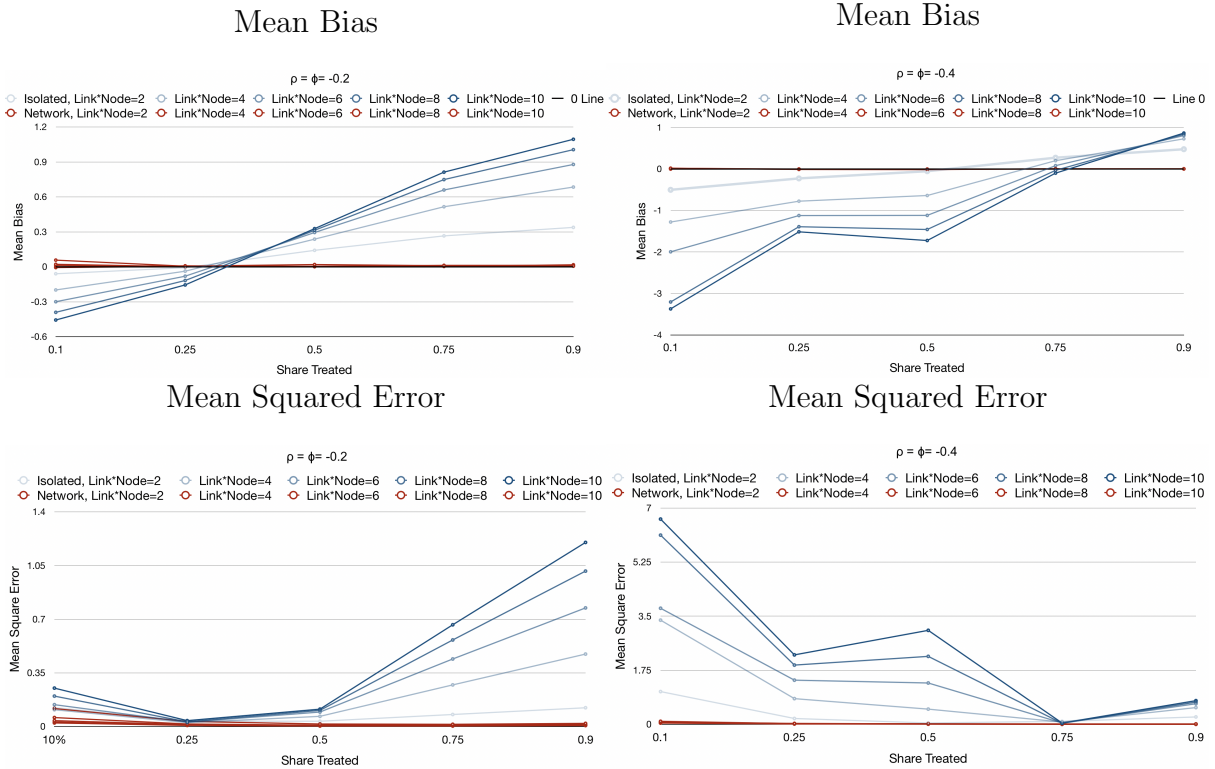
4.2.1 Treatment Effects

Bias under Interdependencies. We first study the bias of β in absence of fixed effects. We set the spillover parameters ϕ and ρ such that $\phi = \rho$ for simplicity, but our results hold in more general settings, as shown in the Appendix. Across exercises, we vary the value of spillovers, the share of treated relationships in the population, and we attach different amounts of credit relationships to each bank and firm node (i.e., we vary the network's number of connections, its density).

Figure 8 reports the mean bias (upper panel) and mean square errors (MSE) from estimating the treatment effect parameter β in the isolated and connected credit models. Blue lines represent ICM results, while red line represent CNM results. Shade intensity is darker as we increase the network density, while we report across different elements of each panel the results we obtain changing the strength of spillovers. In particular, we display results for $\rho = \phi = -0.2$ in the first element of each panel, and $\rho = \phi = -0.4$ in the second.²⁴

²⁴ We report only two such plots per panel for expositional clarity, and we will do so through other Figures over this Section. Nonetheless, in Appendix Tables A.1 through A.7, we report all underlying numbers, also presenting results for the intermediate value of $\rho = \phi = -0.3$ and additional experiments that add to the scope of our results.

Figure 8: Treatment Effect Estimates - ICM and CNM Performance



Notes: Mean bias and mean squared error for treatment effect (β). We compute the mean bias across the 500 simulated samples. For each sample, the number of nodes N in the network is 200, 100 firms and 100 banks. In black, we plot the zero line. In different shades of blue, we plot how the mean bias (upper two figures) and the mean square error (lower two figures) of the ICM's estimate change. In red, we do the same for the CNM's estimates. Darker shades signal denser networks ($m=2,4,6,8,10$). In left panels $\phi = -0.2$, in right panels $\phi = -0.4$. We display the plots' underlying data in Appendix Table A.1.

In Figure 8's upper panel, we can appreciate that the sign and intensity of the bias depend on observable features of the network, i.e. the treated share and the density of relationships. Nonetheless, comparing the left and right elements of Figure 8's upper panel, we can also see that the bias depends on the magnitude of spillovers too. We cannot observe directly these parameters and, as a consequence, we cannot mend the bias without directly confronting it in estimation.

Furthermore, we note that when we treat a small fraction of units, we overestimate the magnitude of β to be much more negative than what it really is, especially if the spillover is larger (right upper panel), while we underestimate it eventually, as the fraction of treated units increases and reaches a certain threshold, which depends on the magnitude of spillovers and the density of the network.

The negative (βEV) and positive (βOD) components of the bias, documented in Section 3.4, explain such a pattern of over and underestimation. When there are few treated units, βEV prevails, as most higher-order effects are feedback loops triggered by own-treatment, which amplifies the negative direct effect. To provide economic interpretation,

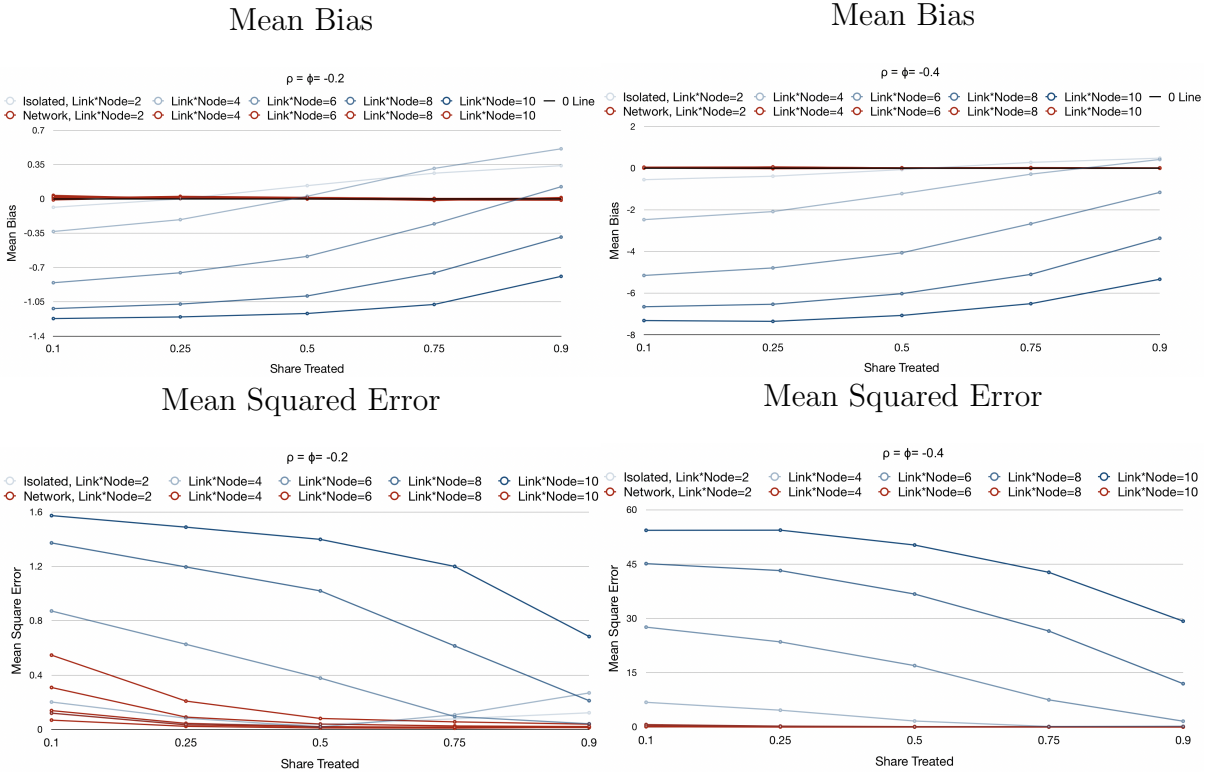
we think of a case in which credit from shocked banks becomes scarcer, and firms with multiple lenders systematically start sourcing more credit from other providers, further decreasing what they demand from the shocked banks in equilibrium. We stress that this is the numeric counterpart of what we showcased through the theoretical example in Section 2.1, when only one credit relationship faces a supply contraction.

Instead, when the supply contraction hits more nodes, the amplifying effect of feedback loops ends up being more than offset by indirect effects from other treated units, which are primarily positive under the assumption of negative treatment and negative spillovers. Again, interpreting the results through economic logic, we can think of the case in which the majority of banks must suddenly contract credit supply, causing each firm to face a sudden increase in all other firms' demand for credit at the few spared banks. Eventually, this indirect effect can be so strong and credit from spared banks so scarce that firms source more credit from the initially hit banks than they would have done in the absence of interconnections.

Higher density, i.e. more relationships per node and thus higher share of borrowers shared by lenders, allows for more feedback loops and thus amplifies the bias' magnitude in both directions. The density also determines the point in which the bias switches its sign: the higher the number of links, the higher the number of treated units needed to switch the sign of the bias to positive, as crowd out is less striking when each firm has more options. Also the value of spillover effects (ϕ and ρ) determines the point at which the bias changes its sign. The higher they are, the stronger direct effects and own-treatment feedback loops, the more treated units we need to switch the bias from negative to positive. This conclusion is the more true, the fewer units are treated, as there is more 'space' to directly absorb the supply contraction.

The results shown do not depend on the setting chosen, which is quite simple but able to point out the main forces at work. We have chosen the density as the main metric for the network, others can be considered. Here we are not particularly interested in finding a "sufficient" statistic for the contribution of the network structure to the bias of β , because we can derive it: endowed with a credit register, we observe A_F and A_B , as well as the treatment vector, and we provide in Section 5 a consistent estimator for ϕ and ρ . With all these elements, we can precisely derive the bias *ex post*.

Figure 9: Treatment Effect Estimates with Fixed Effects - ICM and CNM Performance



Notes: Mean bias and mean squared error for treatment effect (β). We compute the mean bias across the 500 simulated samples, where the data generating process includes a set of fixed effects. For each sample, the number of nodes N in the network is 200, 100 firms and 100 banks. In black, we plot the zero line. In different shades of blue, we plot how the mean bias (upper two figures) and the mean square error (lower two figures) of the ICM's estimate change. In red, we do the same for the CNM's estimates. Darker shades signal denser networks ($m=2,4,6,8,10$). In left panels $\phi = -0.2$, in right panels $\phi = -0.4$, for both panels $\theta = 0.1$. We display the plots' underlying data in Appendix Table A.2.

Finally, we show that the inclusion of bank and firm fixed effects does not solve this issue. On the contrary, the inclusion of fixed effects, ignoring spillovers, can exacerbate bias. In Figure 9's upper panel, we display with blue lines how the bias' severity increases the higher the spillovers and the denser the network. The change in sign for sparse networks confirms again the intuition we conveyed within our theoretical example in Section 2.1, especially in Figure 3. By adding back endogenous reallocation in the treatment, fixed effects put more weight on own-treatment's feedback loops, overplaying shocks' severity.

Estimator Performance. Next, we analyze the performance of our estimator. First, we report the mean bias and the MSE of $\hat{\beta}$ for the same setting used above. We visualize the results as the red lines in Figures 8 and 9. In every simulation, the mean bias is negligible in all settings, with different shares of treated units, magnitude of spillovers and density of the credit network (Figure 8's upper panel). Moreover, our estimator is precise, as measured by the low MSE, which decreases with the number of treated units

and with the density (i.e. with the increase in the number of relationships, thus the sample size; see Figure 8's lower panel). Last, we also see that such performance is not affected by the inclusion of fixed effects in the data generating process (Figure 9)

Second, we analyze the performance of the estimates of the spillover parameters ϕ and ρ , relaxing the assumption that $\phi = \rho$, and also studying what happens when n grows. In Table 1 we report the mean and the standard deviation of the estimates of ϕ and ρ at different intensities (by columns) and densities (by rows). We can see that the estimates are always centered at the true values. The standard deviation decreases with the number of nodes in the network and the density of the connections among them. With a quite small sample size, 800 nodes, the dispersion of the estimates is limited even with very small density (when $m = 2$).

Table 1: Performance of the Spillovers' Estimators ($\hat{\phi}$, $\hat{\rho}$)

n	m		true		true		true		true		true	
			ϕ	ρ	ϕ	ρ	ϕ	ρ	ϕ	ρ	ϕ	ρ
200			-0.1	-0.1	-0.1	-0.2	-0.1	-0.3	-0.1	-0.4	-0.4	-0.4
	2	mean	-0.097	-0.100	-0.100	-0.209	-0.101	-0.306	-0.093	-0.414	-0.406	-0.406
		std	0.084	0.087	0.090	0.089	0.082	0.081	0.082	0.076	0.066	0.067
	6	mean	-0.098	-0.098	-0.097	-0.197	-0.097	-0.295	-0.096	-0.398	-0.402	-0.395
		std	0.029	0.030	0.029	0.030	0.029	0.033	0.028	0.032	0.039	0.040
	10	mean	-0.102	-0.096	-0.099	-0.198	-0.099	-0.297	-0.098	-0.398	-0.402	-0.397
		std	0.021	0.020	0.023	0.023	0.022	0.024	0.021	0.024	0.031	0.030
800												
	2	mean	-0.102	-0.098	-0.100	-0.201	-0.097	-0.301	-0.097	-0.401	-0.398	-0.401
		std	0.041	0.043	0.044	0.042	0.042	0.042	0.040	0.037	0.034	0.033
	6	mean	-0.099	-0.100	-0.099	-0.200	-0.097	-0.300	-0.099	-0.400	-0.398	-0.401
		std	0.015	0.014	0.015	0.016	0.014	0.016	0.014	0.016	0.021	0.020
	10	mean	-0.100	-0.099	-0.100	-0.200	-0.099	-0.300	-0.099	-0.400	-0.399	-0.400
		std	0.010	0.011	0.010	0.011	0.011	0.012	0.010	0.012	0.015	0.015

Notes: Over this Table, we compute the mean and the standard deviation across 500 simulated samples. n is the number of nodes in the network, m regulates the network density as described in Section 4.1. We report further simulation results in Appendix Table A.5.

Bias under Endogenous Treatment. The analytical results in Section 3.4.2 show that in the presence of endogenous treatments, both the CNM and ICM estimates of β can be biased. However, the difference between the two biases can inform about the presence

of spillovers and the latter can be still estimated without distortion. In this section, we study the performance of both estimators when treatment endogeneity is introduced and sequentially increased in finite samples. To analyze this case, we start from the settings used above and generate $\epsilon = \iota X + V$, with V being a normal error term with mean equal to 0 and variance equal to σ . On the one hand, we increase sequentially the endogeneity of the treatment with $\iota = 0, 0.2, 0.5$. On the other hand, we increase the magnitude of spillovers $\phi = \rho = 0, -0.2, -0.4$, as in the previous exercises.

Table 2 reports our results. The first interesting result is that spillovers are always correctly estimated by the CNM, even in the presence of high treatments endogeneity. This is because the spillovers among agents' outcomes can be still recovered even if the treatment is endogenous, as discussed in Section 3.4.2. Moreover, we can see that while if spillovers do not really matter there is no significant difference between the ICM and CNM, when they do matter the estimator we propose performs strictly better than the ICM. The second one is that CNM's estimate bias increases steadily in ι , but is not sensitive to ϕ and ρ , at difference with the ICM estimates', where biases compound. Indeed, we can see that the bias of the ICM increases in both ι and $\phi = \rho$, almost doubling the magnitude of the estimated effect when spillovers and treatment endogeneity are high.

Table 2: Performance of All Estimators When Treatment is Endogenous

			$\iota = 0$			$\iota = -0.2$			$\iota = -0.5$		
			$\hat{\phi}$	$\hat{\rho}$	$\hat{\beta}$	$\hat{\phi}$	$\hat{\rho}$	$\hat{\beta}$	$\hat{\phi}$	$\hat{\rho}$	$\hat{\beta}$
$\phi = \rho = 0$	CNM	mean	0.001	-0.001	-2.007	0.003	0.004	-2.194	0.003	0.001	-2.495
		std	0.033	0.035	0.076	0.031	0.032	0.075	0.027	0.027	0.077
	ICM	mean			-2.008			-2.193			-2.494
		std			0.073			0.071			0.074
$\phi = \rho = -0.2$	CNM	mean	-0.197	-0.197	-2.002	-0.199	-0.197	-2.203	-0.200	-0.199	-2.500
		std	0.033	0.034	0.073	0.029	0.031	0.072	0.027	0.026	0.073
	ICM	mean			-2.231			-2.457			-2.789
		std			0.085			0.084			0.087
$\phi = \rho = -0.4$	CNM	mean	-0.398	-0.399	-2.003	-0.400	-0.398	-2.196	-0.399	-0.401	-2.496
		std	0.030	0.030	0.077	0.027	0.028	0.083	0.023	0.023	0.076
	ICM	mean			-2.915			-3.197			-3.649
		std			0.152			0.152			0.183

Notes: Over this Table, we compute the mean and the standard deviation across 500 simulated samples. The number of nodes is 800, $m = 2$ and $\beta = -2$.

4.2.2 Idiosyncratic Firm and Bank Shocks

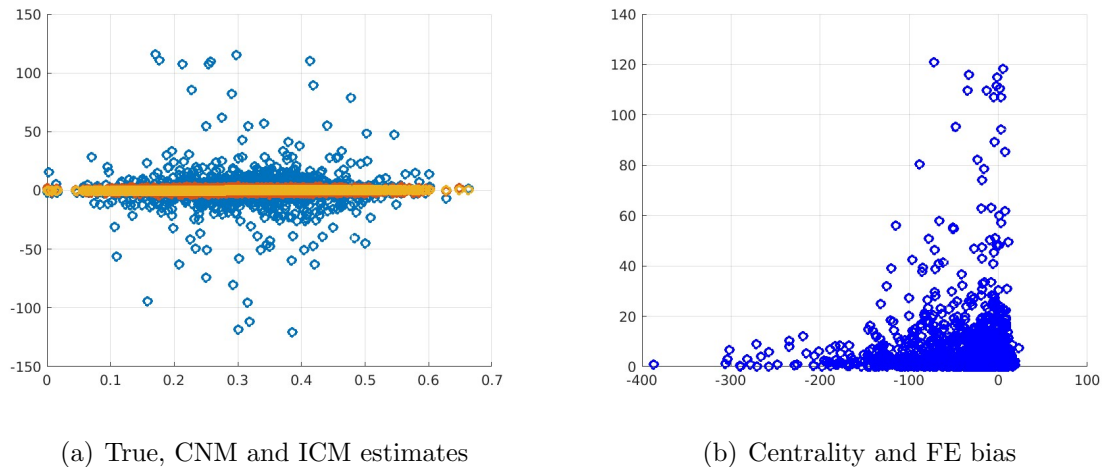
Thanks to our methodology, we can also improve on the estimation of firms and banks fixed effects, carefully separating reallocation from actual idiosyncratic credit supply and demand movements. Fixed effects are not only an important tool to mitigate the concern of unobservable demand and supply shocks biasing treatment estimates. They are also key tools to actually measure idiosyncratic credit demand and supply shifts, allowing us to quantify their real effects (Amiti and Weinstein, 2018). Nonetheless, if we ignore endogenous reallocation of demand and supply across the network of credit relationships, we may confuse it for actual idiosyncratic shocks. Moreover, idiosyncratic shocks diffuse through the credit network themselves, producing feedback loops that may unduly magnify their own estimates.

We discuss the details of our fixed effects estimators' good performance in the Appendix, commenting Table A.6. Here, we instead focus on a short graphical presentation, stressing how bias distributes across nodes in the network and its relation with centrality. We focus on a setting with $n = 2000$, $m = 10$ and $\phi = \rho = -0.4$, to see how idiosyncratic shocks estimates perform with dense networks and high spillovers.

In panel (a) of Figure 10 we plot the true value of the idiosyncratic shocks on the x-axis against themselves (in yellow), the CNM fixed effects estimates (in orange) and the ICM fixed effects estimates (in blue) on the y-axis. We can see that bias for the ICM estimates can be quite severe for some nodes, upward or downward. Indeed, the bias can be positive or negative under negative spillovers, with sign depending on the network topology.²⁵

²⁵ This is consistent with the results in Section A.1.4.

Figure 10: ICM and CNM Fixed Effects' Estimate and Distribution over Nodes



Notes: We sample this network setting $n = 2000$, $m = 10$ and $\phi = \rho = -0.4$, the other parameters are the same as in the pivotal simulation described previously. In panel (a), x-axis: true value of the idiosyncratic shock, y-axis: true value (in yellow), CNM FE estimates (in orange) and ICM FE estimates (in blue). In panel (b), x-axis: our node i centrality measure $D_i' M D_i$, y-axis: ICM FE estimate for node i .

Panel (b) of Figure 10 plots the nodes' centrality on the x-axis, against the absolute value of the ICM-estimates' bias on the y-axis. We measure the centrality of relationships in which node i is involved with $D_i' M D_i$. As shown in Equations (A.11)-(A.12), the bias sign and magnitude depends on the number of loops in which the firm or bank is involved in. In particular, spillovers distort more the fixed effects estimate of more 'central' lenders and borrowers in the credit network.²⁶ In the Figure, we can indeed see that the higher the centrality of the node, the higher the value of the ICM estimates' bias.

In other words, even if their idiosyncratic variation is negligible, lenders or borrowers that are more central in the credit network may show particularly high values of their ICM-estimated fixed effects. Nevertheless, these large values may just be the result of shocks originated by other nodes in the network. Given that central nodes are more exposed to other nodes' influence, they may accumulate a large amount of variation which is not originated by themselves, but which comes from other banks and firms instead.

4.3 Results with Networks Sampled from the Credit Register

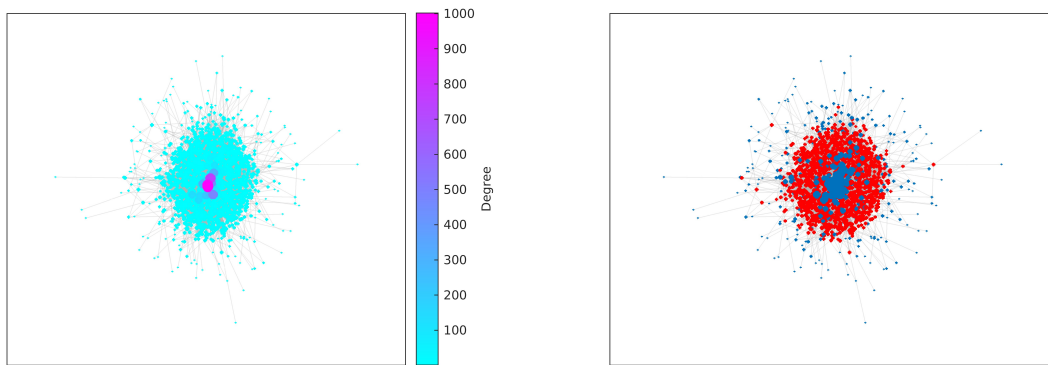
We conclude our simulation study with Monte Carlo experiments based on an observed set of real credit relationships. So far we considered fairly simple network structures. Here, we test whether our method still works when the complexity of the credit relationships topology increases, resembling a real credit market structure. We thus construct G using realized credit relationships between banks and firms from the credit network we use in

²⁶ Still, we remind that under negative spillovers the bias can be close to zero even for central nodes, as positive spillovers from even loops can offset negative ones from odd loops.

the empirical application in Section 5. We randomly extract $n = 400, 800, 2000$ nodes from the full set of relationships observed in 2012 in the Italian credit market between banks and firms and use the links among them in the simulation exercise. The other elements of the simulation are generated in the same way described above.

Figure 11 depicts the network for the 2000 nodes sample. We highlight two features of such a credit network. First, the high interconnectedness between banks and firms; second, the high concentration of connections. In panel (a), the color of each node changes with its degree, i.e. the number of connections it has; more violet nodes represent more connected banks and firms. We can see that the real credit network features high skewness, with some nodes, especially banks, having a very high number of relationships. Some banks lend to thousands of firms in the sample, others only to few of them. This feature of the real credit network is also important for identification, as it guarantees that banks have not fully overlapping portfolios (see Section 3.2 and Proposition 5).

Figure 11: The Credit Network



(a) Degree

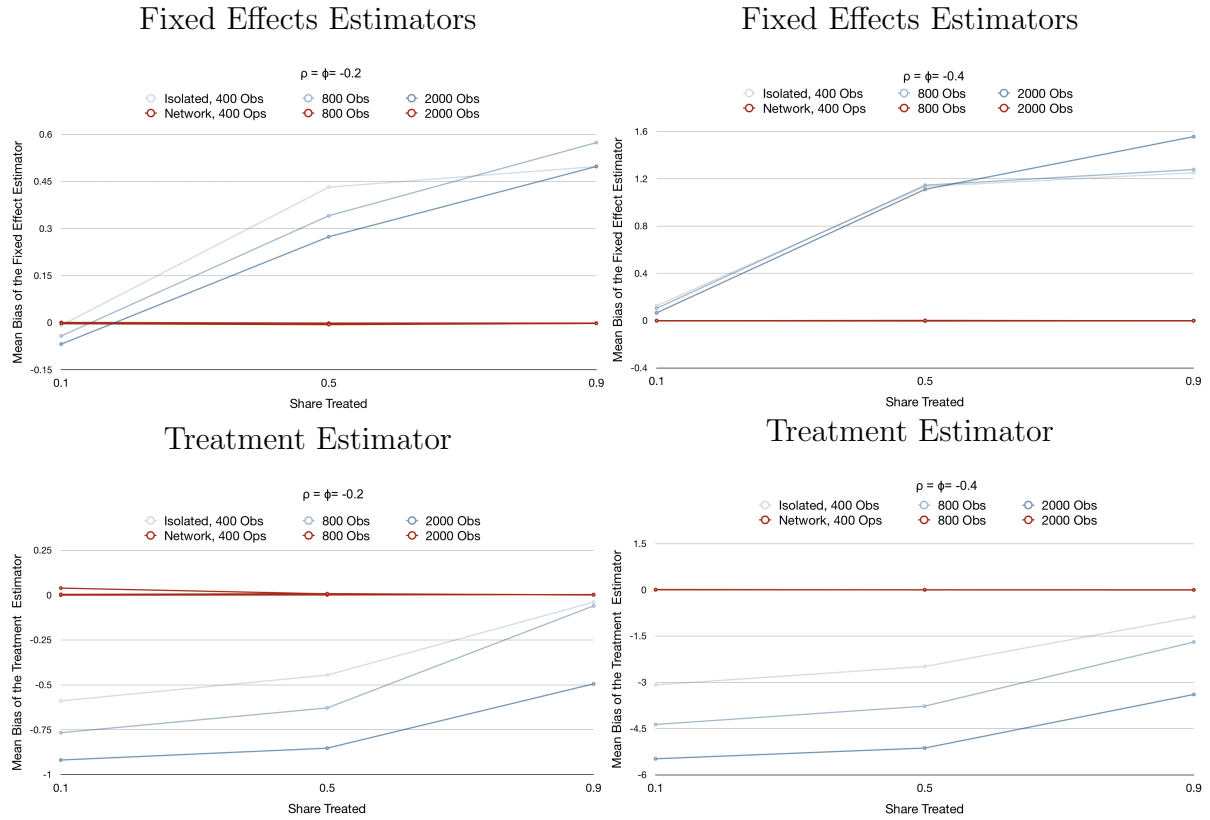
(b) Banks and firms

Notes: We derive the Figure's network sampling 2000 credit relationships observed in 2012. In panel (a), the color of each node is proportional to its degree, more violet nodes represent more connected banks and firms. In panel (b), we represent banks in red and firms in blue. We plot the estimated network with a force-directed layout with five iterations. A force-directed layout uses attractive forces between adjacent nodes and repulsive forces between distant nodes. To ease the visualization, the size of the nodes is equal to the (log) of their degree. See Fruchterman and Reingold (1991) for more details.

In panel (b) we color the nodes denoting firms in blue and banks in red. Looking at the Figure, we note a tight group of banks with a prominently central position in the network, surrounded by a cloud of more peripheral banks. We also highlight a large layer of firms connected to both central and peripheral banks, whose credit relationships connect indirectly many banks. This core-periphery structure resembles that observed in interbank networks (see Boss et al., 2004; Craig and Von Peter, 2014; Iori et al., 2008; Soramäki et al., 2007, among others), only, in this case, connections are not directly

through banks, but indirectly through firms.²⁷

Figure 12: ICM Bias and CNM Performance in a Real Network



Notes: We present results from sampled networks, chosen as random subsets of 400, 800, 2000 relationships from the 2012 Italian Credit Register. In different shades of blue, we plot how the mean bias of the isolated-credit model’s estimates changes as the share of treated increase. In red, how the mean bias for the credit network model does so. Darker shades signal larger real samples. We report the underlying simulation results in Table A.7.

We report mean bias for idiosyncratic shocks (upper panels) and treatment effects (lower panels) in Figure 12, varying the intensity of spillovers, the share of treated units and the sample size.²⁸ First, we note that the mean bias of the CNM estimator is always around zero for both treatment effects and idiosyncratic shocks. On the contrary, the ICM estimator presents a relevant bias, which on average increases with spillovers’ magnitude. Second, we note that the ICM fixed effects estimates’ bias grows with the number of nodes we sample. This growth is also due to the higher density of bigger networks. Indeed, the bigger the network sampled, the less we censor links between sampled and left-out nodes. Accordingly, our 400 nodes sample has a density of 1.25, while the 800 and 2000 samples have respectively 1.75 and 2.4. Third, we highlight how the treatment effect’s

²⁷ As we are working with a sample, the position of each node is not necessarily correctly represented. Nevertheless, the sampling procedure guarantees that all the connections among sampled nodes are included. In the empirical analysis we consider the whole credit network. Unfortunately, its huge dimension does not allow to examine it visually with standard software.

²⁸ Clearly, as we are sampling from a real network, we cannot vary the density here.

bias increase with the share of treated units, exactly as discussed for Figure 8 in Section 4.2.1.

In conclusion, the results of our last Monte Carlo experiment confirm that our estimator provides consistent estimates for treatment effects and idiosyncratic shocks even under a more complex structure of credit relationships, while conventional estimators are still severely biased.

5 Empirical Application

In our empirical exercise, we focus on the dynamics of credit to firms at the intensive margin. Following the vast majority of recent and past works in empirical banking, we employ the yearly log change of credit granted on each relationship as our dependent variable of interest.

As we want to study the bias of estimated bank and firm fixed effects (idiosyncratic shocks) as well as of treatment effects, we do not focus on point-in-time treatments at the firm or bank-level (e.g., banks' interbank market exposure during a freeze, as in Bonaccorsi di Patti and Sette, 2016; Iyer et al., 2013). We instead focus on *relationship-level* exposure to recurring macroeconomic shocks. The relationship-level effect of macroeconomic shocks has been often explored in the literature by interacting relationships' characteristics with changes in macroeconomic variables, which allows joint estimation of banks and firms' fixed effects (see Jiménez et al., 2014, among others).

Our aim here is not validating a specific strategy used in the literature. Our focus is on the quantification of relationship-level endogenous spillovers, as well as their effects on treatment and fixed effects' estimates. The specific setting, or the endogeneity challenge each single study aims to tackle, is really not the point here, also because we have shown analytically and numerically (see Section 3.4.2 and 4.2.1, respectively) that we can precisely quantify spillovers and their importance for other parameters even if the treatment is endogenous.

For all these reasons, we do not replicate a specific analysis. We instead focus on relationship-level effects of changes in the policy rates. Monetary policy is a consistent source of changes in banks cost of funding, allows this paper not to be time or experiment-specific. Furthermore, it is especially interesting nowadays, when we want to better understand the foreseeable effects of rates increases on firms. In the following, we will use the interaction between the one-year lag of relationships' revolving intensity and percentage points changes in the Italian interbank overnight rate as our primary independent variable. Indeed, revolving credit lines have either variable or relatively easy-to-re-bargain rates; hence, revolving-intensive credit relationships will likely bear the most immediate effects of changes in banks' cost of funding.

5.1 Data Description

To perform such an exercise, we use 2012 to 2018 data from (i) the Italian credit register, which tracks all credit relationships between Italian firms and banks whose total exposure in terms of granted credit is greater than 30 thousand euros. We match this data with (ii) the Company Accounts Data System (CADS), balance sheet information for the universe of Italian non-financial corporations provided by the Cerved group, and (iii) the Italian Supervisory Reports, which contain Italian banks' balance sheets and group structure.

Following the literature, we focus on firms with multiple credit relationships so that we can estimate firm fixed effects. Thus, we drop observations belonging to firms with only one relationship each year. Furthermore, we drop all observations belonging to firms with troubled credit relationships (*deteriorati* and *sofferenze*), as well as relationships belonging to foreign banks or non-bank financial intermediaries. Finally, we drop observations with missing granted credit data. We obtain seven yearly samples as a result. Each covers between five and four hundred thousand observations, belonging to about 150 thousand unique firms and five-to-four hundred banks.

Table 3 documents the basic characteristics of each yearly sample. First, the dynamic of granted credit in log changes is mostly negative across the whole study period, with the worst performance recorded in 2012, immediately after the European sovereign debt crisis. Then, we report our relationship-level explanatory variables, the ratio between revolving and total credit (F. Revolving); the ratio between credit granted on the relationship and total credit granted to the firm (F. Granted); a dummy taking value one if the firm and bank's headquarters are located in the same province (Same Prov.). These variables are mostly stable across the years we consider.

As our independent variable of interest is the interaction between lag revolving ratios and overnight rate changes, tracking the heterogeneous impact of movements in banks' refinancing rates, we also document the end-of-year change (year-on-year) in the Italian banks overnight refinancing rate in percentage points.²⁹ Finally, due to our interest in fixed effects estimation, we report data on the number of relationships per firm and bank, which are stable over the years. The number of firms is stable over time, while the number of banks decreases gradually, mainly because of consolidation in the banking sector.

²⁹ 3-Month or 90-day Interbank Rates for Italy, retrieved from the ECB Statistical Data Warehouse.

Table 3: Descriptive Statistics

	2012			2013			2014		
	<i>Mean</i>	<i>Std Dev.</i>	<i>Count</i>	<i>Mean</i>	<i>Std Dev.</i>	<i>Count</i>	<i>Mean</i>	<i>Std Dev.</i>	<i>Count</i>
Δ Log Granted	-0.092	0.445	517,885	-0.070	0.418	484,953	-0.028	0.431	446,107
F. Revolving	0.289	0.336	517,885	0.298	0.342	484,953	0.295	0.342	446,107
F. Granted	0.084	0.068	517,885	0.084	0.067	484,953	0.085	0.066	446,107
Same Prov.	0.275	0.446	517,885	0.270	0.444	484,953	0.258	0.438	446,107
	<i>Value</i>			<i>Value</i>			<i>Value</i>		
Δ Overnight Rate %	-1.3			0.046			-0.16		
	<i>Mean</i>	<i>Median</i>	<i>Max</i>	<i>Mean</i>	<i>Median</i>	<i>Max</i>	<i>Mean</i>	<i>Median</i>	<i>Max</i>
N. Rel. Firm	3.922	2.239	27	3.907	2.238	25	3.898	2.258	23
N. Rel. Bank	30,417	20,668	70,708	27,964	19,480	66,748	25,437	18,037	60,554
	<i>Count</i>			<i>Count</i>			<i>Count</i>		
N. Firm	159,893			157,352			145,474		
N. Banks	546			542			525		
	2015			2016			2017		
	<i>Mean</i>	<i>Std Dev.</i>	<i>Count</i>	<i>Mean</i>	<i>Std Dev.</i>	<i>Count</i>	<i>Mean</i>	<i>Std Dev.</i>	<i>Count</i>
Δ Log Granted	-0.024	0.464	423,056	-0.021	0.452	423,829	-0.011	0.453	410,045
F. Revolving	0.281	0.337	423,056	0.267	0.332	423,829	0.255	0.327	410,045
F. Granted	0.086	0.069	423,056	0.086	0.071	423,829	0.086	0.071	410,045
Same Prov.	0.247	0.431	423,056	0.235	0.424	423,829	0.225	0.418	410,045
	<i>Value</i>			<i>Value</i>			<i>Value</i>		
Δ Overnight Rate %	-0.171			-0.223			-0.016		
	<i>Mean</i>	<i>Median</i>	<i>Max</i>	<i>Mean</i>	<i>Median</i>	<i>Max</i>	<i>Mean</i>	<i>Median</i>	<i>Max</i>
N. Rel. Firm	3.917	2.290	27	3.969	2.364	25	3.929	2.345	21
N. Rel. Bank	24,336	17,638	57,891	23,635	17,733	56,275	23,372	18,053	54,870
	<i>Count</i>			<i>Count</i>			<i>Count</i>		
N. Firm	141,297			137,321			128,953		
N. Banks	493			467			428		
	2018								
	<i>Mean</i>	<i>Std Dev.</i>	<i>Count</i>						
Δ Log Granted	-0.019	0.467	402,919						
F. Revolving	0.245	0.322	402,919						
F. Granted	0.085	0.070	402,919						
Same Prov.	0.206	0.404	402,919						
	<i>Value</i>								
Δ Overnight Rate %	0.013								
	<i>Mean</i>	<i>Median</i>	<i>Max</i>						
N. Rel. Firm	3.908	2.339	37						
N. Rel. Bank	25,034	27,786	59,289						
	<i>Count</i>								
N. Firm	138,328								
N. Banks	400								

Notes: This Table presents descriptives for the samples used in the estimation. Each panel covers one year. In each panel, the first four lines record descriptives for the dependent and main relationship-level independent variables. The fifth line reports the value of the end-of-year change in the bank overnight rate in percentage points. The sixth and seventh lines report descriptives concerning the number of relationships per firm and bank. Finally, the last two lines report the total firm and bank count.

5.2 Empirical Specification

We estimate model (10) using the estimators in (21)-(22), and data pooled from 2012 to 2018. As said, we want to measure the impact on relationships' granted credit growth of the interaction between the one-year lag of revolving intensity and percentage point changes in overnight interbank rate. Our main concern with such a specification is that revolving credit intensity can correlate with relationship lending, as revolving lines embody the actual bank-firm relationship (Berger and Udell, 1995), while other forms of credit most often come from non-recurring needs for funding.

First, we notice that this potential bias is acceptable in principle, because it implies underestimating rate changes' effect. There is indeed ample evidence (e.g. Berlin and Mester, 1998; Sette and Gobbi, 2015) that relationship lenders at least partially shelter their long-term customers from the impact of shocks. Moreover, we use our rich dataset to mitigate this concern further. As relationship lending can be correlated with unobservable firm and bank characteristics, we fully control for these with bank-time and firm-time fixed effects. Then, we add two different proxies of relationship lending, i.e., the lag of the ratio between credit granted on the relationship and credit granted to the firm, measuring the relationship's importance to the firm, and a dummy taking the value of one if the firm and bank's headquarters are located in the same province.³⁰

The resulting Isolated Credit Model (ICM) is as follows:

$$\begin{aligned} \Delta \text{Log Granted}_{ibt} = & \delta_{it} + \gamma_{bt} + \beta \Delta \text{Overnight Rate}_{t-1} * \text{F. Revolving}_{ibt-1} + \dots \\ & \mu \text{Controls}_{ibt-1} + \varepsilon_{ibt} \end{aligned} \quad (29)$$

where δ_{it} is the firm-time fixed effect; γ_{bt} the bank-time fixed effect; $\Delta \text{Overnight Rate}_t * \text{F. Revolving}_{ibt}$ is our main variable of interest, which we will refer to as Treat_{ibt} when displaying results in Table 4; Controls_{ibt} is a matrix containing the lag in the relationship revolving ratio not interacted with changes in rates, the ratio between credit granted on the relationship and total granted, and the headquarter location dummy.

The ICM's β tracks differences in the growth of credit granted on bank-firm relationships that are more revolving credit intensive, after a rate change, within the *same* firm, absorbing all time-varying bank unobservables, and controlling for relationship characteristics. However, the ICM does not control for spillover effects among relationships. For example, following the arguments in Section 2.1, after the rates drop, a firm could reallocate its credit demand towards relationships for which the discount's pass-through is greater, from relationships for which it is smaller. Comparing credit growth on two such relationships will bias the estimated β 's magnitude upward, if we do not account for

³⁰ On the relationship between distance and lending, see Agarwal and Hauswald (2010); Degryse and Ongena (2005).

endogenous reallocation first.

To account for such bias, as well as from higher order indirect effects (derived and presented in Section 3.4), we estimate the following Credit Network Model (CNM):

$$\begin{aligned} \Delta \text{Log Granted}_{ibt} = & \delta_{it} + \gamma_{bt} + \beta \Delta \text{Overnight Rate}_t * \text{F. Revolving}_{ibt-1} + \dots \\ & \phi N_B \Delta \text{Log Granted}_{ibt} + \rho N_F \Delta \text{Log Granted}_{ibt} + \dots \quad (30) \\ & \mu \text{Controls}_{ibt-1} + \varepsilon_{ibt} \end{aligned}$$

here, we introduce as further controls the bank- and firm-network lags of the dependent variable, which keep track of the connections among relationships through banks and firms. We formalize this addition with the N_B and N_F operators, such that $N_B x_{ibt} = \sum_{j \in \mathbb{F} \setminus i} a_{ib,jb} x_{jbt}$ is the bank-network lag of x_{ibt} and $N_F x_{ibt} = \sum_{k \in \mathbb{B} \setminus b} a_{ib,ik} x_{ikt}$ the firm-network lag of x_{ibt} .

5.3 Main Results

Estimated Spillovers and Treatment Effects. Table 4 reports the results, with the 2SLS' second stage estimates of the ICM and the CNM on the left panel and the CNM's first stages for the two endogenous variables on the right panel.

First, we note that the bank and firm's spillover coefficients are highly statistically significant, with ratios between coefficients and standard errors well above three. Concerning magnitudes, the estimate of ρ (FDA) is about -0.6, while the estimate of ϕ (BSA) is much smaller. We expect this size disparity, as banks have many more relationships than firms. Thus, we re-scale the ϕ estimate by 10,000, or slightly less than half the number of corporate credit relationships of the average bank (see Table 3), obtaining the final ϕ^* estimate of -0.07.

We can perform the following two back-of-the-envelope calculations to understand the above estimates' magnitudes and their economic significance. For firms, we can think of one with three relationships (the average number we report in Table 3), seeing granted credit shrinking by 20 percent from two of its three banks. Such a firm will likely ask for an extension of the remaining line, and our ρ estimate suggests that this last will grow by about 24 percent ($0.24 = -0.6 * 2 * -0.20$). For banks, we can think of one granting 20 percent more credit to 10,000 borrowers. Our ϕ^* estimate suggests that such action would imply a 1.5 percent crowd-out on all bank's other credit lines ($-0.0146 = 0.2 * -0.073$).

Then, we report the estimates of treatment and controls' coefficients. First, we see that credit relationships belonging to the same firm with greater revolving ratios expand

more after overnight rate decreases.³¹ Then, overall, credit relationships that are more revolving intensive (F. Revolving_{ibt-1}), as well as relationships that are more important for the total credit access of the firm (F. Granted_{ibt-1}), grow less, while relationships for which firm and bank’s headquarters are in the same province grow slightly more. These conclusions hold in both the estimated CNM and ICM.

Focusing on the main variable of interest in the CNM (Table 4, Column (1), line 4) and re-scaling the treatment effect by the standard deviation in revolving fractions (about 0.3, see Table 3), we observe that credit relationships one standard deviation more revolving-intensive grow by 6 percent more after a one percent decrease in the overnight rate. We would have overestimated the magnitude of this coefficient by a factor of 2.5 had we not accounted for the network nature of credit relationships, as we can appreciate from the ICM panel (Table 4, Column (3), line 4). We expect such a large bias, as, looking at within-firm changes, the ICM estimator is likely to sum the endogenous demand reallocation induced by supply shifts, magnifying the estimates.

Regarding the first stage, we derive the dependent variables computing, for each relationship, the total growth in credit on other relationships by the same bank (Table 4, Columns (5-6)) or firm (Table 4, Columns (7-8)).³² Then, we regress these quantities on all the bank and firm-network lags of controls we include in the second stage. As we employ both network lags jointly in each first stage, the coefficients we display are net of the bank-lags and firm-lags effects and are difficult to interpret meaningfully. Nonetheless, the critical insight from the first stage panel in Table 4 is that the values of F_{SW} , the F-test for weak instruments in linear IV models with multiple endogenous variables proposed by Sanderson and Windmeijer (2016), are large and do not lend support to weak IV concerns (see Table 4, Columns (5) through (8), third line from the bottom).³³ These large F_{SW} s indicate that the instruments are relevant and both endogenous variables are sharply identified.

³¹ We comment on negative variations of the overnight rate as most large changes in our sample, documented in Table 3, are negative. Hence, the most empirically relevant variation is the effect of a one percent drop, as the one registered in 2012.

³² Coefficients and errors for variables’ bank-network lag in the first stage of $N_F \Delta \text{Log Credit}_{ibt}$ are multiplied by 10,000 (Table’s Columns (7-8), first four lines). Moreover, in the $N_B \Delta \text{Log Credit}_{ibt}$ first stage, the coefficients and errors for the variables’ firm-network lags are divided by 10,000 (Table’s Columns (5-6), last four lines).

³³ In greater detail, both figures are largely above the relevant minimum eigenvalue’s threshold for the Cragg-Donald statistic, tabulated by Stock and Yogo (2002).

Table 4: Spillovers Estimates, First and Second Stage

	(1)	(2)	(3)	(4)		(5)	(6)	(7)	(8)
Dependent Variable:	Second Stage				First Stage				
	$\Delta\text{Log Credit}_{ibt}$				$N_B\Delta\text{Log Credit}_{ibt}$ $N_F\Delta\text{Log Credit}_{ibt}$				
	CNM		ICM		IV	$N_B\Delta\text{Log Credit}_{ibt}$		$N_F\Delta\text{Log Credit}_{ibt}$	
	coeff.	std. err.	coeff.	std. err.		coeff.	std. err.	coeff.	std. err.
Firm Spillover: ρ	-0.6204	0.0032			$N_B\text{Treat}_{ibt-1}$	0.1050	0.0002	-0.0161	0.0011
Bank Spillover: ϕ^*	-0.0703	0.0001			$N_B\text{F. Revolving}_{ibt-1}$	-0.2388	0.0004	0.0252	0.0029
					$N_B\text{Same Prov.}_{ibt-1}$	-0.1419	0.0003	0.0353	0.0019
Treat_{ibt-1} : β	-0.1887	0.0016	-0.4661	0.0020	$N_B\text{F. Granted}_{ibt}$	0.5796	0.0015	-0.2028	0.0103
F. Revolving $_{ibt-1}$: μ_1	-0.1616	0.0014	-0.4148	0.0011	$N_F\text{Treat}_{ibt-1}$	-0.0201	0.0004	-0.1760	0.0029
Same Prov. $_{ibt}$: μ_2	0.0112	0.0003	0.0258	0.0006	$N_F\text{F. Revolving}_{ibt-1}$	-0.0087	0.0002	-0.1847	0.0016
F. Granted $_{ibt-1}$: μ_3	-0.5302	0.0049	-1.4116	0.0042	$N_F\text{Same Prov.}_{ibt-1}$	-0.0023	0.0001	-0.0115	0.0007
					$N_F\text{F. Granted}_{ibt}$	-0.3282	0.0048	-0.5356	0.0322
N	3,108,758		3,108,758			3,108,758		3,108,758	
F_{SW}						154,602		2,144	
Bank-FE	Yes		Yes			Yes		Yes	
Firm-FE	Yes		Yes			Yes		Yes	

Notes: Estimated coefficients and standard errors for the second and first stages of model (11) employing the 2SLS estimator in Equation (22), in the particular case of Equation (30). The second stage dependent variable $\Delta\text{Log Credit}_{ibt}$ is the yearly log growth rate of the credit relationship. N_B is the bank-network lag operator, and it equals to $\sum_{j \in \mathbb{F} \setminus i} a_{ib,jb} x_{jbt}$ for every x covariate; N_F is the firm-network lag operator, and it equals to $\sum_{k \in \mathbb{B} \setminus b} a_{ib,ik} x_{ikt}$ for every x covariate. Treat_{ibt-1} is the revolving ratio multiplied by the change in the overnight interest rate. Bank-related coefficients are re-scaled to account for the disparity in the number of relationships between banks and firms. In particular, ϕ^* and its errors are multiplied by 10,000, as well as coefficients and errors for variables' bank-network lag in the first stage of $N_B\Delta\text{Log Credit}_{ibt}$ (Table's columns 7 and 8, first four lines). Moreover, in the $N_B\Delta\text{Log Credit}_{ibt}$ first stage, the coefficients and errors for the variables' firm-network lags are divided by 10,000 (Table's columns 5 and 6, last four lines). F_{SW} statistics are reported for the first stages. The F_{SW} is the F-test for weak instruments in linear IV models with multiple endogenous variables proposed by Sanderson and Windmeijer (2016).

Idiosyncratic Shocks' Bias. We now analyze the estimated banks' and firms' fixed effects. Following our simulation study in Section 4, we focus on the distortion that occurs when spillovers are not accounted for. Given that we found significant spillover effects, we use the estimates from the CNM as unbiased measures of γ_{bt} and δ_{it} and compare them with ICM estimates, i.e., the ones not including the endogenous terms capturing the BSA and the FDA.

In Table 5, we report the empirical Mean and Median Bias, as well as the Mean and Median Absolute Bias, separately for firms and banks' fixed effects (FEs henceforth) estimates.

Table 5: Fixed Effects, Empirical Bias Measures

	Bank	Firm
Mean Bias	1.329	-0.750
Median Bias	0.735	-0.444
Mean Absolute Bias	1.414	1.297
Median Absolute Bias	0.742	0.564

Notes: The Table reports the empirical mean bias, mean absolute bias, median bias, and median absolute bias for firm and bank estimated fixed effects. We consider the fixed effects estimated correcting for network structure as the true parameters. We compute the bias measures with the following formulas, here reported only for the bank fixed effect case: $MB_{\text{Bank}} = 1/B \left\{ \sum_{k=1}^B (\hat{\gamma}_{ICM}^k - \hat{\gamma}_{CNM}^k) / |\hat{\gamma}_{CNM}^k| \right\}$, $MAB_{\text{Bank}} = 1/B \left\{ \sum_{k=1}^B |\hat{\gamma}_{ICM}^k - \hat{\gamma}_{CNM}^k| / |\hat{\gamma}_{CNM}^k| \right\}$, $MedB_{\text{Bank}} = Med \{ (\hat{\gamma}_{ICM} - \hat{\gamma}_{CNM}) \text{Diag} [|\hat{\gamma}_{CNM}|^{-1}] \}$, $MedAB_{\text{Bank}} = Med \{ (|\hat{\gamma}_{ICM} - \hat{\gamma}_{CNM}|) \text{Diag} [|\hat{\gamma}_{CNM}|^{-1}] \}$. Where B is the number of banks in the market, Med is the median operator, MB stands for Mean Bias, MedB for Median Bias, MAB for Mean Absolute Bias, and MedAB for Median Absolute Bias.

In Table 5, each indicator aggregates the difference between the FE estimated by the ICM and the FE estimated by the CNM. We focus on differences divided by the absolute value of the CNM's FEs so that magnitudes are in percentage of the unbiased estimate and easier to compare. First, looking at the mean and median bias, bank FEs are overestimated on average by the ICM. On the contrary, firm FEs are underestimated. The positive bias for banks ranges between 73 (median) to 133 (mean) percent of the true parameter, while the negative bias for firms ranges between 44 (median) to 75 (mean) percent. Second, we notice that the bias's average and median absolute values are prominent for bank and firm fixed effects. The median absolute bias stands around 74 percent for banks and 56 percent for firms. The values for the mean absolute bias are even higher, pointing to the presence of highly biased FE for some banks and firms.

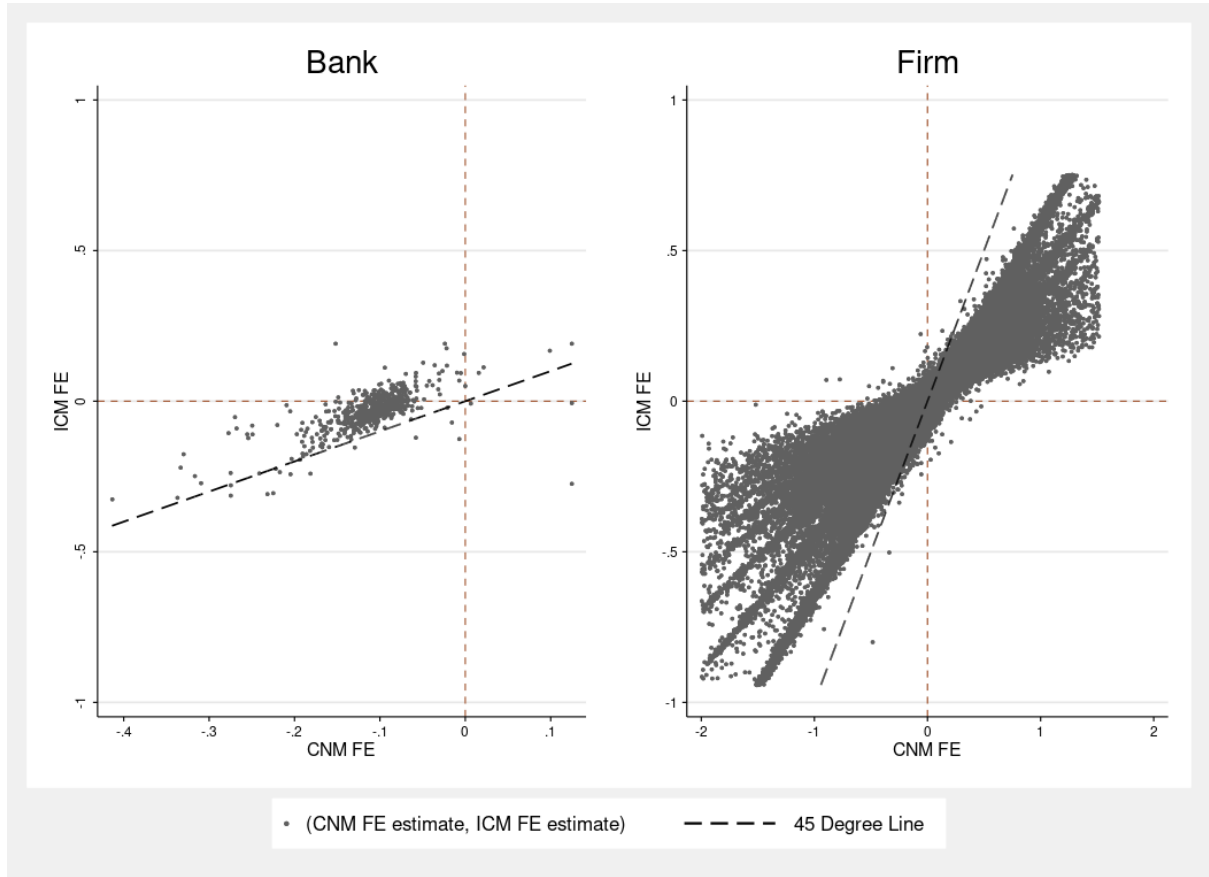
The large magnitude of the bias is easily explained by looking at the economic importance of spillover effects in Table 4's estimates and thinking of links' density over banks. At the firm level, first-order spillovers are substantial, while many relationships are connected through banks. When not accounting for the network nature of credit relationships, first-order spillovers through firm links lead to overestimating treatment effects. This overestimation happens at the expense of the idiosyncratic firm component if the latter is correlated with the treatment status. Then, as banks connect many nodes in the network, the network-wide transmission of spillovers through bank links leads us to overestimate the banks' idiosyncratic effects. When accounting for the network structure, we see that this last bias was just the result of the high centrality of banks in the credit relationships' network, as suggested before by simulation results (see Figure 10).

We can deepen the nature of the ICM-estimated fixed effects bias by plotting the

ICM estimates against the CNM unbiased estimates for each firm and bank. We do so in Figure 13, which we read as follows. First, on the left panel, we see that almost all ICM-estimated bank fixed effects are above the 45-degree. Hence, the ICM model is systematically overestimating idiosyncratic credit supply changes. In particular, it *underestimates the magnitude* of bank-level credit *contractions*. Furthermore, one-quarter of the scatter points lie in the upper-left quadrant; which is, we have numerous cases in which the CNM unbiased estimate is negative while the ICM biased estimate is positive. This sign discrepancy means that the bias in the ICM model would point to nonexistent idiosyncratic supply expansions for one in four banks.

Then, in the right panel, we can see a graphical depiction of how the ICM systematically underestimates the magnitude of idiosyncratic demand changes. Moreover, even if less often than with banks, the ICM switches the sign of firms' idiosyncratic demand changes for seven percent of firms. In conclusion, the evidence we gather points to the fact that estimating bank and firm fixed effects in an ICM may lead to a large bias in a real empirical setting. The bias's size and magnitude are node-specific, often switching the sign of the estimated idiosyncratic shock. Thus, using such estimates may be highly misleading.

Figure 13: The Empirical Distribution of Fixed Effects Estimates' Bias



Notes: The Figure displays, on the left, the scatterplot of bank fixed effects' estimates in the CNM (x -axis) and ICM (y -axis); on the right, the scatterplot of firm fixed effects' estimates in the CNM (x -axis) and ICM (y -axis). The short-dashed line is the 45-degree line.

6 Extensions and Discussions

In this section, we present some extensions of the CNM, to allow for the endogeneity of credit relationships and treatments, and for heterogeneous ϕ and ρ , which may be of practical relevance in empirical studies. We then compare our model with other approaches used in the literature.

6.1 Endogenous Credit Relationships Network

A common concern in the econometric analysis of spillovers over networks is the possible endogeneity of the network itself. Issues can arise if, for example, some unobserved factors drive the formation of the links (here the credit relationships) and the outcomes (credit quantity). For the link formation, the following type of dyadic network formation model is a standard choice

$$g_{ib} = I(d(h_i, h_b) \geq u_{ib}), \quad (31)$$

where $g_{ib} = 1$ if there exist a credit relationship between firm i and bank b , h_i and h_b are unobserved individual specific characteristics, u_{ib} is a link-specific random component, and $d(.,.)$ is some matching function. The unobserved node-specific characteristic h_i can be interpreted as a factor that increases the likelihood of forming a link. Network endogeneity may arise if firm (bank) individual unobserved characteristic h_i (h_b), affecting link formation, is correlated with firm's (bank's) unobserved characteristic determining the outcome c_{ib} . It is indeed easy to imagine that lower firm risk or higher profitability of projects, as well as advantages in monitoring or screening for banks, affect match-likelihood and credit granted both.

Compared to models in which the outcome is at the node level (see Arduini, Patacchini, and Rainone, 2015; Auerbach, 2022; Goldsmith-Pinkham and Imbens, 2013; Hsieh and Lee, 2016; Johnsson and Moon, 2021; Patacchini and Rainone, 2017; Qu and Lee, 2015, among the others), a key difference in our context is that we model outcomes at the link level. This feature allows us to include node fixed effects (δ_i and γ_b), which alleviates this type of concerns. Nevertheless, one can also assume that there are link level correlated unobservables. For example, think of

$$g_{ib} = I(d(h_i, h_b, h_{ib}) \geq u_{ib}), \quad (32)$$

with h_{ib} correlated with ϵ_{ib} . In this case, a connection between two credit relationships is observed if

$$a_{ib,jb} = g_{ib}g_{jb} = I(d(h_i, h_b, h_{ib}) \geq u_{ib})I(d(h_j, h_b, h_{jb}) \geq u_{jb}). \quad (33)$$

It follows that even if $E[h'_{ib}\epsilon_{ib}] \neq 0$ it does not imply that $E[a'_{ib,jb}\epsilon_{ib}] \neq 0$ because h_{ib} enters $a_{ib,jb}$ in a highly non linear form and it is multiplied by terms not necessarily correlated with ϵ_{ib} . In addition, all the terms in (33) only share b -indexed variables, but node b specific factors are absorbed by γ_b .

It is, however, possible to apply a control function approach at the link level similar to those developed by Arduini, Patacchini, and Rainone (2015); Johnsson and Moon (2021),³⁴ or a Bayesian approach (in the spirit of Goldsmith-Pinkham and Imbens, 2013; Hsieh and Lee, 2016; Patacchini and Rainone, 2017). Therefore, controlling for h_{ib} , the network A and ϵ_{ib} become mean independent, that is,

$$E(\epsilon_{ib}|A, h_{ib}) = E(\epsilon_{ib}|h_{ib}) =: k(h_{ib}). \quad (34)$$

³⁴See Rainone (2020b) for an application of interbank markets. Even if his model does not include node fixed effects, the control function approach does not change the estimated parameters significantly.

We can then consider the outcome equation that controls for \hat{h}_{ib} nonparametrically,

$$c_{ib} = \alpha + \phi \sum_{j \in \mathbb{F} \setminus i} a_{ib,jb} c_{jb} + \rho \sum_{k \in \mathbb{B} \setminus b} a_{ib,ik} c_{ik} + \delta_i + \gamma_b + x_{ib} \beta + k(\hat{h}_{ib}) + u_{ib}, \quad (35)$$

where $u_{ib} := \epsilon_{ib} - k(\hat{h}_{ib})$. Once we control the endogeneity of the network, the regressor of the spillover effect becomes exogenous, and we can estimate coefficient ϕ using the conventional partially linear regression estimation method (Robinson, 1988).

It is worth noting that in our model the endogeneity of credit relationship may not only affect the estimate of β or other coefficients per se, it can also affect the estimates of ϕ and ρ through the potential endogeneity of A_B and A_F , as discussed above. To the best of our knowledge, the only paper that accounts for endogeneity of links in the credit market is Jiménez et al. (2014). They exploit unique data on loan applications to restrict the dimensionality problem, estimate a selection equation that involves the granting of loans in the first stage and credit outcome equations for the applications granted in the second stage. Differently from our study, they focus on sample selection issues, and do not model interdependence among credit relationships.³⁵ Unfortunately, loan application information is not always available in credit register data. Further, we note that staple models of network formation do not allow for network effects in link formation, as a link between i and b depends only on the characteristics of i and b . Indeed, especially in this context an interesting extension could incorporate a transferable matching step in the spirit of Fox (2009). As this is a non-trivial extension, we leave this work for future research. For all these reasons, we focus on conditionally exogenous networks for the time being.

6.2 Endogenous Treatments

In the previous analysis, we assumed that the main regressor is exogenous. Our method can accommodate an instrumental variable strategy. For example, let us now allow X to be an endogenous regressor and assume that a valid instrument W is available, $X = W\kappa + \omega$, and $E[\epsilon'\omega] \neq 0$, then we can include the instrument and its network-lags in the first step in a quite straightforward way (see Anselin and Lozano-Gracia, 2008; Dall’Erba and Le Gallo, 2008, for applications of this procedure).³⁶ In practice, the empirical IV in

³⁵In particular, Jiménez et al. (2014) use the method proposed by Kyriazidou (1997) (see also Arellano and Honoré, 2001; Honore, Kyriazidou, and Powell, 2000), which does not require distributional assumptions (like normality of the errors in the selection equation) and differences out both the sample selection effect and the unobservable individual effect from the equation of interest, under the conditional exchangeability assumption. This assumption could be stronger for our model, given the presence of interdependence.

³⁶ See Fingleton and Le Gallo (2008) for the finite sample properties of this type of estimators.

Equations (16) and (17) can be augmented in the following way

$$EIV_F^{1W} = [A_F X, A_F W, W], \quad (36)$$

$$EIV_B^{1W} = [A_B X, A_B W, W]. \quad (37)$$

6.3 Heterogeneous BSA and FDA

The CNM can be augmented to have bank- and firm-type specific spillovers, for example if one is interested in studying the substitution of credit from a type of banks to another type of banks by firms, or the reallocation of credit from a certain type of firms to another one by banks, and eventually combinations of the two depending on the specific empirical questions.

A typical example for the FDA is the substitution of credit from low technology banks to high technology ones by firms demanding new and more modern financial services.³⁷ Another one, focusing this time on the BSA, is the reallocation of credit by banks from sectors hit by specific shocks to unaffected ones.³⁸ Let us focus on the first case and suppose that there are H (high tech) and L (low tech) banks.

If we have these two types, we will have four types of ρ s. ρ^H captures the spillovers among relationships involving type H banks; ρ^L captures the spillovers among relationships involving type L banks; ρ^{LH} captures the spillovers from relationships with type H bank to those with type L ; ρ^{HL} captures the spillovers from relationships with type L bank to those with type H . If, for example, one expects high substitution of credit from low tech banks to high tech banks, ρ^{HL} should be higher than the others.

Model (11) will then become:

$$C = (\rho^H A_F^H + \rho^L A_F^L + \rho^{HL} A_F^{HL} + \rho^{LH} A_F^{LH})C + \phi A_B C + Z\mu + \epsilon. \quad (38)$$

where the matrix A_F^H keeps track of connections among relationships involving the same firm and banks of type H ; the matrix A_F^L keeps track of connections among relationships involving the same firm and banks of type L ; A_F^{HL} and A_F^{LH} are symmetrically equal matrices that keep track of connections among relationships involving the same firm and banks of both types.

In such extension of the baseline model, the instrumental variables change accordingly to the different specification. First order approximations of the best IV for the five

³⁷ See Fuster et al. (2019), Fuster et al. (2018), Branzoli, Rainone, and Supino (2023), Core and De Marco (2021) and Kwan et al. (2021), among the others, for studies on the effects of bank technological adoption and credit.

³⁸ See Paravisini, Rappoport, and Schnabl (2023), Federico, Marinelli, and Palazzo (2023) and Federico, Hassan, and Rappoport (2023), among the others, for studies on lending behavior by banks more specialized or with loan portfolios concentrated in sectors more exposed to shocks.

endogenous variable are respectively:

$$\begin{aligned}
EIV_{FH}^1 &= A_F^H X, \\
EIV_{FL}^1 &= A_F^L X, \\
EIV_{FLH}^1 &= A_F^{LH} X, \\
EIV_{FHL}^1 &= A_F^{HL} X, \\
EIV_B^1 &= A_B X.
\end{aligned}$$

Second order approximations of the best IV for the five endogenous variable are respectively:

$$\begin{aligned}
EIV_{FH}^2 &= A_F^H [I, A_B] X, \\
EIV_{FL}^2 &= A_F^L [I, A_B] X, \\
EIV_{FLH}^2 &= A_F^{LH} [I, A_B] X, \\
EIV_{FHL}^2 &= A_F^{HL} [I, A_B] X, \\
EIV_B^2 &= A_B [I, A_F^H, A_F^L, A_F^{LH}, A_F^{HL}] X.
\end{aligned}$$

Similar derivations can be computed for higher order approximations.

If we are interested in heterogeneous reallocation policies by the banks, suppose within and between two sectors (say S and T), we will have again four types of ϕ s. ϕ^S captures the spillovers among relationships involving firms in sector S ; ϕ^T captures the spillovers among relationships involving firms in sector T ; ϕ^{TS} captures the spillovers from relationships with firms in sector S to those in sector T ; ϕ^{ST} captures the spillovers from relationships with firms in sector T to those in sector S . If, for example, one expects high substitution of credit from low tech banks to high tech banks, ρ^{LH} should be higher than the others.

Model (11) will then become:

$$C = (\phi^S A_B^S + \phi^T A_B^T + \phi^{ST} A_B^{ST} + \phi^{TS} A_B^{TS})C + \rho A_F C + Z\mu + \epsilon. \quad (39)$$

where the matrix A_B^S keeps track of connections among relationships involving the same bank and firms of sector S ; the matrix A_B^T keeps track of connections among relationships involving the same bank and firms of sector T ; A_B^{ST} and A_B^{TS} are symmetrically equal matrices that keep track of connections among relationships involving the same bank and firms of both sectors.

In such extension of the baseline model, the instrumental variables change accordingly to the different specification. First order approximations of the best IV for the five

endogenous variable are respectively:

$$\begin{aligned}
EIV_{FH}^1 &= A_B^S X, \\
EIV_{FL}^1 &= A_B^T X, \\
EIV_{FLH}^1 &= A_B^{TS} X, \\
EIV_{FHL}^1 &= A_B^{ST} X, \\
EIV_B^1 &= A_F X.
\end{aligned}$$

Second order approximations of the best IV for the five endogenous variable are respectively:

$$\begin{aligned}
EIV_{FH}^2 &= A_B^S [I, A_F] X, \\
EIV_{FL}^2 &= A_B^T [I, A_F] X, \\
EIV_{FLH}^2 &= A_B^{TS} [I, A_F] X, \\
EIV_{FHL}^2 &= A_B^{ST} [I, A_F] X, \\
EIV_B^2 &= A_F [I, A_B^S, A_B^T, A_B^{TS}, A_B^{ST}] X.
\end{aligned}$$

Similar derivations can be computed for higher order approximations.

Overall, when we are interested in heterogeneous FDA and BSA, instrumental variables depend on the final specification and thus on the specific research question. We thus do not provide sufficient identification conditions as those in Proposition 5, for any possible configuration. Nevertheless, a more general condition for identification is that the matrix including the expected value of the endogenous variables and the other covariates in the model has full rank. In this example, we need the matrix

$$[E(A_B^S C), E(A_B^T C), E(A_B^{ST} C), E(A_B^{TS} C), E(A_F C), Z]$$

to have full rank. The intuition is that as long as there are intransitive quadriads for any combination of heterogeneous BSA and FDA resulting from the selected choices, this condition is always respected, the parameters are identified and the IVs can be constructed as linear combinations of the vectors appearing in the expected value of each endogenous terms.

6.4 Comparison between OPIVs, SSIVs, and GIVs

In general terms, OPIVs are related to other approaches decomposing market's aggregate outcomes to derive instrumental variables, such as shift share instrumental variables (Bartik, 1991; Blanchard et al., 1992; Borusyak, Hull, and Jaravel, 2022; Goldsmith-Pinkham, Sorkin, and Swift, 2020, SSIVs, see), and granular instrumental variables (GIVs, Gabaix

and Kojien, 2020). However, our approach differs from the GIVs and SSIVs approaches substantially. OPIVs and both these approaches are actually complements, because they can be used in different types of markets.

The GIVs and SSIVs are procedures designed to estimate price elasticities in centralized markets, where there is only one price, while we design OPIVs to estimate objects more similar to elasticities of substitution in decentralized markets, where the price varies at the pair level and the identity of counterparties matters. For example, GIVs exploit the fact that in centralized markets single agents demand (or supply) depends on the aggregate price, but the aggregate price does not depend on the demand of the single agent, rather the aggregate demand. The instrumental variable is obtained when few large actors account for a substantial fraction of aggregate demand (supply) and idiosyncratic shocks are volatile relative to the volatility of aggregate shocks. Although being deeply different from each other, both GIVs and SSIVs derive instruments by decomposing aggregate quantities (like the aggregate demand, for example) and using exogenous components, under different assumptions.

The OPIV does not decompose any aggregate quantity, it instead derives instrumental variables for endogenous disaggregated outcomes, exploiting intransitivity in decentralized markets. The OPIVs exploits the fact that in decentralized markets agents demand (or supply) in a single contract depends on what happens in other contracts involving the same parties. We obtain OPIVs under the condition that agents have not fully overlapping portfolios of counterparties. Whereas GIVs use players' size disparities to derive exclusion restrictions, OPIVs use intransitivity. GIVs need that idiosyncratic shocks to large players can be separated from systemic ones (granularity). What OPIVs need is that not all banks lend to all firms.

6.5 Comparison with Spillover Papers in Corporate Finance

In this section we discuss the main differences between our model and those of the family which Huber (2022), Berg, Reisinger, and Streitz (2021) belong to. The first difference is about the main outcome variables and the source of spillovers considered. We study the formation of outcomes in the credit market, and specifically we focus on quantities of loans. They study the effects of shocks to banks to firms' outcomes, such as employment. They assume that spillovers come from firms operating in the same region or sector. Our spillovers come from relationships that share the very same counterparty. In other words, they provide tools to account for spillovers among firms when the effects of financial shocks on real outcomes are analyzed. We provide a tool to model spillovers among credit relationships and account for them when the effects of financial shocks on credit outcomes are analyzed.

That we focus on credit outcomes implies another difference: we study outcomes

at the bilateral level, because we look at credit market outcomes that are bilateral by construction, the other papers look at individual outcomes. This feature allows us to exploit the micro structure of the market for identification and infer on counterparties specific behaviors. It comes with non trivial differences in the complexity of the model and in the interpretation of the results. Indeed, we can provide a structural model that not only controls for spillovers, but also allows to recover parameters that have a direct behavioral interpretation in terms of credit substitution and reallocation, as discussed above.

In summary, we focus on endogenous effects and they focus on exogenous effects. The endogenous effects capture agents' choices/outcomes that depend on that of others, and can be interpreted through assumed response functions of agents, while exogenous (or contextual) effects control for others' exogenous characteristics or treatment status. From this perspective, the two approaches can be useful complements to study the effects of financial shocks.

7 Conclusion

In the last three decades, empirical research has consistently demonstrated that shocks to banks significantly impact firms' credit availability (e.g., Behn, Haselmann, and Wachtel, 2016; Jiménez et al., 2017; Khwaja and Mian, 2008; Paravisini, 2008; Paravisini et al., 2014; Peek and Rosengren, 1995, 2000), subsequently influencing employment and investment decisions (e.g., Amiti and Weinstein, 2018; Chodorow-Reich, 2014; Chodorow-Reich and Falato, 2022; Cingano, Manaresi, and Sette, 2016; Jiménez et al., 2020).

A pivotal tool in achieving these findings has been the within-firm (and often within-bank) estimation strategy, popularized initially by Khwaja and Mian (2008). This tool is crucial due to non-random matching between banks and firms (see, e.g., Schwert, 2018), necessitating the separation of shocks to banks from the unobservable idiosyncratic shocks affecting their partner firms. The ability to disentangle these two types of shocks is a key initial step in assessing the impact of credit supply contractions and expansions on the real economy.

However, fixed effects have inherent limitations. First, they assume a uniform distribution of idiosyncratic shocks across different credit relationships within the same firm or bank. If the data violates this assumption, residual bias may persist, affecting estimates (as noted in Paravisini, Rappoport, and Schnabl, 2023). Second, fixed effects cannot account for contextual spillovers, and they isolate sets of observations where these are more likely (as noted in, e.g. Berg and Streitz, 2019; Huber, 2018). Last, fixed effects cannot address the endogenous reallocation of demand and supply of credit across relationships after a shock.

In this paper, we introduced a methodology for measuring exactly such reallocation

without relying on additional information such as price data. To do so, we adapted standard results in network theory and estimation (Ballester, Calvó-Armengol, and Zenou, 2006; Bramoullé, Djebbari, and Fortin, 2009) to the credit market. Using overlapping portfolio instrumental variables, we demonstrated how to estimate the extent to which firms and banks reallocate credit across their relationships' portfolios following a financial shock. Furthermore, by controlling for reallocation, we provided a possible avenue to tackle the reflection bias in treatment and fixed effects estimates. Our approach strengthens the standard methodology and mitigates the concern that reallocation contaminates fixed effects estimates, when retrieved to proxy for demand and supply idiosyncratic shifts.

The ability to directly quantify reallocation across relationship portfolios is essential to comprehend *how* credit supply movements translate into real effects. For instance, how much does credit market concentration matter for the shocks' pass-through? Do specialized lenders play an oversized role, and how? Our methodology should aid researchers in addressing similar questions, even in the absence of extremely rich datasets or quasi-experimental settings to account for reallocation.³⁹ Moreover, the tool we propose could be especially useful for studying the effects of regulation changes and monetary policy, often enacted when reallocation is possible and easy. Indeed, as a practical application, we used the Italian Credit Register data to demonstrate how failing to account for endogenous reallocation of credit demand and supply would lead to overestimating the pass-through of monetary policy from 2012 to 2018.

The primary contribution of this paper is providing an innovative tool to estimate and control for credit reallocation effects. Our ongoing work will expand this framework to account for relationship heterogeneity, the creation margin of credit relationships, and facilitating accessibility by offering computational tools for estimation.

³⁹ In terms of setup, researchers could use a partial population design to account for reallocations and endogenous effects, as Angelucci and De Giorgi (2009), which uses village-level randomization of treatment to assess cash transfers' effects. Nonetheless, such a setup is extremely difficult to run into in reality.

Appendix

A.1 Proofs

A.1.1 Model

Full derivation of the toy model, proving Proposition 1.

Bank Problem:

$$\text{Bank } b: \max_{c_{ib}, c_{jb}} (r_{ib} - \omega(c_{ib} - \xi x_{ib} - \theta c_{jb} - \nu_{ib}))c_{ib} + (r_{jb} - \omega(c_{jb} - \xi x_{jb} - \theta c_{ib} - \nu_{jb}))c_{jb}$$

$$\text{Bank } a: \max_{c_{ia}} (r_{ia} - \omega(c_{ia} - \xi x_{ia} - \nu_{ia}))c_{ia}$$

FOC deliver:

$$\begin{aligned} r_{ib} &= \omega c_{ib} - \omega \underbrace{(\xi x_{ib} + \nu_{ib} - \theta c_{jb})}_{u_{ib}} \\ r_{jb} &= \omega c_{jb} - \omega \underbrace{(\xi x_{jb} + \nu_{jb} - \theta c_{ib})}_{u_{jb}} \\ r_{ia} &= \omega c_{ia} - \omega \underbrace{(\xi x_{ia} + \nu_{ia})}_{u_{ia}} \end{aligned}$$

(A.1)

Firm problem:

$$\text{Firm i: } \max_{c_{ia}, c_{ib}} (e_i - \alpha(c_{ia} + c_{ib}))(c_{ia} + c_{ib}) - \sum_{K=a,b} c_{iK} \omega (c_{iK} - u_{iK})$$

$$\text{Firm j: } \max_{c_{jb}} (e_j - \alpha c_{jb}) c_{jb} - c_{jb} \omega (c_{jb} - u_{jb})$$

FOC deliver:

$$e_i - 2\alpha c_{ia} - 2\alpha c_{ib} - 2\omega c_{ia} + \omega(\xi x_{ia} + \nu_{ia}) = 0$$

$$e_i - 2\alpha c_{ib} - 2\alpha c_{ia} - 2\omega c_{ib} + \omega(\xi x_{ib} + \nu_{ib} - \theta x_{jb}) = 0$$

$$e_j - 2\alpha c_{jb} - 2\omega c_{jb} + \omega(\xi x_{jb} + \nu_{jb} - \theta x_{ib}) = 0$$

Which simplifies to:

$$c_{ia} = -\frac{\alpha}{\alpha+\omega} c_{ib} + \frac{1}{2(\alpha+\omega)} e_i + \frac{\omega}{2(\alpha+\omega)} (\xi x_{ia} + \nu_{ia})$$

$$c_{ib} = -\frac{\alpha}{\alpha+\omega} c_{ia} + \frac{1}{2(\alpha+\omega)} e_i + \frac{\omega}{2(\alpha+\omega)} (\xi x_{ib} + \nu_{ib} - \theta c_{jb})$$

$$c_{jb} = \frac{1}{2(\alpha+\omega)} e_j + \frac{\omega}{2(\alpha+\omega)} (\xi x_{jb} + \nu_{jb} - \theta c_{ib})$$

And delivers the following structural demand system:

$$c_{ia} = \rho c_{ib} + \beta x_{ia} + \delta_i + \epsilon_{ia}$$

$$c_{ib} = \rho c_{ia} + \phi c_{jb} + \beta x_{ib} + \delta_i + \epsilon_{ib}$$

$$c_{jb} = \phi c_{ib} + \beta x_{jb} + \delta_j + \epsilon_{jb}$$

Calling:

$$\rho = -\frac{\alpha}{\alpha+\omega}$$

$$\phi = -\frac{\theta\omega}{2(\alpha+\omega)}$$

$$\beta = \frac{\xi\omega}{2(\alpha+\omega)}$$

$$\delta_{i,j} = \frac{1}{2(\alpha+\omega)} e_{i,j}$$

$$\epsilon_{ia,ib,jb} = \frac{\omega\nu_{ia,ib,jb}}{2(\alpha+\omega)}$$

From the above, we can derive the following reduced form system:

$$\begin{aligned}
c_{ia} &= \rho c_{ib} + \beta x_{ia} + \delta_i + \epsilon_{ia} \\
c_{ib} &= \rho(\rho c_{ib} + \beta x_{ia} + \delta_i + \epsilon_{ia}) \dots \\
&+ \phi(\phi c_{ib} + \beta x_{jb} + \delta_j + \epsilon_{jb}) + \beta x_{ib} + \delta_i + \epsilon_{ib} \\
c_{jb} &= \phi c_{ib} + \beta x_{jb} + \delta_j + \epsilon_{jb} \\
\\
c_{ia} &= \rho c_{ib} + \beta x_{ia} + \delta_i + \epsilon_{ia} \\
(1 - \rho^2 - \phi^2)c_{ib} &= \beta(\rho x_{ia} + x_{ib} + \phi x_{jb}) \dots \\
&+ (1 + \rho)\delta_i + \phi\delta_j + \rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib} \\
c_{jb} &= \phi c_{ib} + \beta x_{jb} + \delta_j + \epsilon_{jb} \\
\end{aligned} \tag{A.2}$$

$$\begin{aligned}
c_{ia} &= \rho \left(\frac{(1+\rho)\delta_i + \phi\delta_j}{1-\phi^2-\rho^2} + \beta \frac{\rho x_{ia} + \phi x_{jb} + x_{ib}}{1-\phi^2-\rho^2} + \frac{\rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2} \right) \dots \\
&+ \beta x_{ia} + \delta_i + \epsilon_{ia} \\
c_{ib} &= \frac{(1+\rho)\delta_i + \phi\delta_j}{1-\phi^2-\rho^2} + \beta \frac{\rho x_{ia} + \phi x_{jb} + x_{ib}}{1-\phi^2-\rho^2} + \frac{\rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2} \\
c_{jb} &= \phi \left(\frac{(1+\rho)\delta_i + \phi\delta_j}{1-\phi^2-\rho^2} + \beta \frac{\rho x_{ia} + \phi x_{jb} + x_{ib}}{1-\phi^2-\rho^2} + \frac{\rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2} \right) \dots \\
&+ \beta x_{jb} + \delta_j + \epsilon_{jb} \\
\\
c_{ia} &= \frac{\rho(1+\rho-\phi^2)\delta_i + \rho\phi\delta_j}{1-\phi^2-\rho^2} + \beta \frac{(1-\phi^2)x_{ia} + \rho\phi x_{jb} + \rho x_{ib}}{1-\phi^2-\rho^2} + \frac{(1-\phi^2)\epsilon_{ia} + \rho\phi\epsilon_{jb} + \rho\epsilon_{ib}}{1-\phi^2-\rho^2} \\
c_{ib} &= \frac{(1+\rho)\delta_i + \phi\delta_j}{1-\phi^2-\rho^2} + \beta \frac{\rho x_{ia} + \phi x_{jb} + x_{ib}}{1-\phi^2-\rho^2} + \frac{\rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2} \\
c_{jb} &= \frac{\phi(1+\rho)\delta_i + (1-\rho^2)\delta_j}{1-\phi^2-\rho^2} + \beta \frac{\rho\phi x_{ia} + (1-\rho^2)x_{jb} + \phi x_{ib}}{1-\phi^2-\rho^2} + \frac{\phi\rho\epsilon_{ia} + (1-\rho^2)\epsilon_{jb} + \phi\epsilon_{ib}}{1-\phi^2-\rho^2}
\end{aligned}$$

Adding Assumption 3 and considering the case in which the econometrician ignores spillovers and correlated demand shocks both, we obtain:

$$\begin{aligned}
c_{ia} &= \beta x_{ia} + \varepsilon_{ia} \\
\varepsilon_{ia} &= \delta_i + \rho c_{ib} + \epsilon_{ia} \\
c_{ib} &= \frac{(1+\rho)}{1-\phi^2-\rho^2} \delta_i + \beta \frac{\rho}{1-\phi^2-\rho^2} x_{ia} + \frac{\rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2}
\end{aligned} \tag{A.3}$$

which results in

$$\hat{\beta}_{\text{OLS}} = \frac{\text{cov}(c_{ia}, x_{ia})}{\text{var}(x_{ia})} = \beta + \underbrace{\rho \frac{\text{cov}(x_{ia}, c_{ib})}{\text{var}(x_{ia})}}_{\text{spillover bias}} + \underbrace{\frac{\text{cov}(x_{ia}, \delta_i)}{\text{var}(x_{ia})}}_{\text{demand bias}} \tag{A.4}$$

The fact that $\frac{\text{cov}(x_{ia}, c_{ib})}{\text{var}(x_{ia})} = \frac{1+\rho}{1-\rho^2-\phi^2} \frac{\text{cov}(x_{ia}, \delta_i)}{\text{var}(x_{ia})} + \beta \frac{\rho}{1-\rho^2-\phi^2} \neq 0$ concludes the proof.

■

Proof of Proposition 2.

Indicating averages with bars, so that, for example, $\bar{c}_i = \frac{c_{ia} + c_{ib}}{2}$, we have that:

$$c_{ia} = \beta x_{ia} + \delta_i + \varepsilon_{ia},$$

$$c_{ib} = \delta_i + \varepsilon_{ib}.$$

$$\Rightarrow \hat{\delta}_i = c_{ib}$$

$$\hat{\beta}_{FE} = \frac{\text{cov}(c_{ia} - \bar{c}_i, x_{ia} - \bar{x}_i)}{\text{var}(x_{ia} - \bar{x}_i)} = \frac{\text{cov}\left(\frac{c_{ia} - c_{ib}}{2}, \frac{x_{ia}}{2}\right)}{\text{var}\left(\frac{x_{ia}}{2}\right)} = \frac{\text{cov}(c_{ia} - c_{ib}, x_{ia})}{\text{var}(x_{ia})} = \dots$$

$$\dots \frac{\text{cov}(\beta x_{ia} + \varepsilon_{ia} - \varepsilon_{ib}, x_{ia})}{\text{var}(x_{ia})}$$

From the structural demand system:

$$\varepsilon_{ia} = \rho c_{ib} + \epsilon_{ia} \tag{A.5}$$

$$\varepsilon_{ib} = \rho c_{ia} + \phi c_{jb} + \epsilon_{ib}$$

$$\Rightarrow \frac{\text{cov}(\beta x_{ia} + \varepsilon_{ia} - \varepsilon_{ib}, x_{ia})}{\text{var}(x_{ia})} = \beta + \frac{\text{cov}(\rho(c_{ib} - c_{ia}) - \phi c_{jb}, x_{ia})}{\text{var}(x_{ia})} = \dots$$

$$\dots \beta + \frac{\text{cov}(\rho((1-\rho)c_{ib} - \beta x_{ia} - \delta_i) - \phi c_{jb}, x_{ia})}{\text{var}(x_{ia})} = \dots$$

$$\dots \beta(1-\rho) + \rho(1-\rho) \frac{\text{cov}(c_{ib}, x_{ia})}{\text{var}(x_{ia})} - \rho \frac{\text{cov}(\delta_i, x_{ia})}{\text{var}(x_{ia})} - \phi \frac{\text{cov}(c_{jb}, x_{ia})}{\text{var}(x_{ia})}$$

From the reduced form system, simplified thanks to Assumption 3:

$$c_{ib} = \frac{(1+\rho)\delta_i}{1-\phi^2-\rho^2} + \beta \frac{\rho x_{ia}}{1-\phi^2-\rho^2} + \frac{\rho \epsilon_{ia} + \phi \epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2}$$

$$c_{jb} = \frac{\phi(1+\rho)\delta_i}{1-\phi^2-\rho^2} + \beta \frac{\rho \phi x_{ia}}{1-\phi^2-\rho^2} + \frac{\phi \rho \epsilon_{ia} + (1-\rho^2)\epsilon_{jb} + \phi \epsilon_{ib}}{1-\phi^2-\rho^2}$$

From the above, and the reduced form of System 2 displayed in the last passage, it is evident that β_{FE} is biased, and that correlated demand shocks still play a role, as they are reflected back in the estimator through reallocation spillovers. β_{FE} is indeed a function of δ_i through the $-\rho \text{cov}(\delta_i, x_{ia})/\text{var}(x_{ia})$ element, from demand reallocation within the relationships of the same firm, and through the impact of δ_i on *all* other bias components.

■

Proof of Proposition 3. From the end of Proposition 1's Proof and the absence of demand bias it follows that:

$$\hat{\beta}_{OLS} = \beta \left(1 + \frac{\rho^2}{1-\phi^2-\rho^2}\right) = \beta \frac{1-\phi^2}{1-\phi^2-\rho^2}$$

From the reduced form demand system, Proposition 2's and the absence of demand bias:

$$\hat{\beta}_{FE} = \beta(1-\rho) + \rho(1-\rho) \frac{\text{cov}(c_{ib}, x_{ia})}{\text{var}(x_{ia})} - \phi \frac{\text{cov}(c_{jb}, x_{ia})}{\text{var}(x_{ia})} = \dots$$

$$\beta(1-\rho) + \beta(1-\rho) \frac{\rho^2}{1-\phi^2-\rho^2} - \beta \frac{\phi^2 \rho}{1-\phi^2-\rho^2} = \dots$$

$$\beta \frac{(1-\rho)(1-\phi^2-\rho^2) + \rho^2(1-\rho) - \phi^2 \rho}{1-\phi^2-\rho^2} = \beta \frac{1-\rho-\phi^2 + \rho\phi^2 - \rho^2 + \rho^3 + \rho^2 - \rho^3 - \phi^2 \rho}{1-\phi^2-\rho^2} = \dots$$

$$\beta \frac{1-\phi^2 + \rho}{1-\phi^2-\rho^2} \tag{A.6}$$

■

A.1.2 The econometric framework

Proof of Proposition 5. It follows from the proof of Proposition 1 in Arduini, Patacchini, and Rainone (2020) when G , the network among nodes, and its sub-matrices are replaced by A and its sub-matrices, the network among links. Moving from nodes to links implies that quadriads instead of triads intransitivity is needed. Quadriads intransitivity is implied by linear independence of I_F , $A_B A_F A_B$ and A_F and I_B , $A_F A_B A_F$ and A_B . See condition 1 of the proposition. Alternatively, also the proof of Proposition 1 in Rainone (2020a) brings to the same result if multiple endogenous terms are considered.

■

Details for the identification example in Section 3.3.

$$\begin{aligned}
 c_{ia} &= \rho c_{ib} + \beta x_{ia} + \epsilon_{ia} \\
 c_{ib} &= \rho c_{ia} + \phi c_{jb} + \epsilon_{ib} \\
 c_{jb} &= \phi c_{ib} + \beta x_{jb} + \epsilon_{jb}
 \end{aligned}
 \tag{A.7}$$

$$\begin{aligned}
c_{ia} &= \rho c_{ib} + \beta x_{ia} + \epsilon_{ia} \\
c_{ib} &= \rho(\rho c_{ib} + \beta x_{ia} + \epsilon_{ia}) + \phi(\phi c_{ib} + \beta x_{jb} + \epsilon_{jb}) + \epsilon_{ib} \\
c_{jb} &= \phi c_{ib} + \beta x_{jb} + \epsilon_{jb}
\end{aligned}$$

$$\begin{aligned}
c_{ia} &= \left(\frac{\rho^2\beta}{1-\phi^2-\rho^2} + \beta\right)x_{ia} + \rho\frac{\phi\beta}{1-\phi^2-\rho^2}x_{jb} + \frac{(1-\phi^2)\epsilon_{ia} + \rho\phi\epsilon_{jb} + \rho\epsilon_{ib}}{1-\phi^2-\rho^2} \\
c_{ib} &= \frac{\rho\beta}{1-\phi^2-\rho^2}x_{ia} + \frac{\phi\beta}{1-\phi^2-\rho^2}x_{jb} + \frac{\rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2} \\
c_{jb} &= \left(\frac{\phi^2\beta}{1-\phi^2-\rho^2} + \beta\right)x_{jb} + \frac{\phi\rho\beta}{1-\phi^2-\rho^2}x_{ia} + \frac{(1-\rho^2)\epsilon_{jb} + \phi\rho\epsilon_{ia} + \phi\epsilon_{ib}}{1-\phi^2-\rho^2}
\end{aligned}$$

$$\begin{aligned}
&\text{call } \mu = \frac{\rho\epsilon_{ia} + \phi\epsilon_{jb} + \epsilon_{ib}}{1-\phi^2-\rho^2} \\
&\text{and } \pi_\rho = \frac{\beta\phi}{1-\phi^2-\rho^2} \\
&\text{and } \pi_\phi = \frac{\beta\rho}{1-\phi^2-\rho^2} \\
\Rightarrow c_{ia} &= \rho\pi_\rho x_{jb} + \left(\beta + \frac{\beta\rho^2}{1-\phi^2-\rho^2}\right)x_{ia} + \rho\mu + \epsilon_{ia} \\
c_{ib} &= \pi_\rho x_{jb} + \pi_\phi x_{ia} + \mu \\
c_{jb} &= \phi\pi_\phi x_{ia} + \left(\frac{\phi^2\beta}{1-\phi^2-\rho^2} + \beta\right)x_{jb} + \phi\mu + \epsilon_{jb}
\end{aligned} \tag{A.8}$$

$$\begin{aligned}
\hat{\pi}_{\rho,OLS} &= \pi_\rho \\
\hat{\pi}_{\phi,OLS} &= \pi_\phi \\
\widehat{\pi_\rho\rho}_{OLS} &= \pi_\rho\rho \\
\widehat{\pi_\phi\phi}_{OLS} &= \pi_\phi\phi \\
\hat{\rho}_{IV} &= \frac{\widehat{\pi_\rho\rho}_{OLS}}{\hat{\pi}_{\rho,OLS}} = \rho \\
\hat{\phi}_{IV} &= \frac{\widehat{\pi_\phi\phi}_{OLS}}{\hat{\pi}_{\phi,OLS}} = \phi \\
\hat{\beta}_{OLS} &= \beta\left(1 + \frac{\rho^2}{1-\phi^2-\rho^2}\right) \\
\hat{\beta}_{IV} &= \hat{\beta}_{OLS} - \frac{\hat{\rho}_{IV}^2}{1-\hat{\phi}_{IV}^2-\hat{\rho}_{IV}^2}
\end{aligned}$$

■

A.1.3 Further results for the econometric framework

Econometric model with fixed effects. Let us now introduce also the presence of bank and firm fixed effects. Suppose now that the DGP is

$$\begin{aligned}
C &= \phi A_F C + \rho A_B C + X\beta + \Delta + \Gamma + \epsilon, \\
&= +\phi A_B C + \rho A_F C + Z\mu + \epsilon,
\end{aligned} \tag{A.9}$$

where $\Delta = D\delta$ is a matrix of firm FEs and $\Gamma = G\gamma$ is a matrix of bank FEs. For simplicity let again $\rho = \phi$. Suppose we estimate

$$C = X\beta + \Delta + \Gamma + U, \tag{A.10}$$

then the error term has the following form

$$\begin{aligned}
U &= \phi A_F C + \rho A_B C + \epsilon \\
&= (\phi A_B + \rho A_F)(I - \phi A_B - \rho A_F)^{-1}[X\beta + \Delta + \Gamma + \epsilon] + \epsilon \\
&= M[X\beta + \Delta + \Gamma] + (M + I)\epsilon
\end{aligned}$$

Suppose again that X is univariate and $X \perp \epsilon$, we have

$$\begin{aligned}
X'U &= X'M[X\beta + \Delta + \Gamma] + (M + I)\epsilon = X'\phi A \sum_{k=0}^{\infty} (\phi A)^k [X\beta + \Delta + \Gamma] \\
&= X' \sum_{k=1}^{\infty} (\phi A)^k [X\beta + \Delta + \Gamma] = S + \sum_{k=1}^{\infty} \phi^k X' A^k [\Delta + \Gamma]
\end{aligned}$$

Here, $A^k \Delta = D_k$ is a vector whose generic element contains the FEs of borrowers of credit relationships at distance k from the relative relationship. Symmetrically, $A^k \Gamma = G_k$ is a vector whose generic element contains the FEs of lenders of credit relationships at distance k from the relative relationship. If $X \perp \Delta, \Gamma$, then the bias obtained estimating (A.1.4) is equal to S , similarly to random effects models. If not, $\phi > 0$ and X is positively correlated with Δ and Γ , then $B = (X'X)^{-1}X'U > 0$. Otherwise, the sign of the bias is again ambiguous.

A.1.4 Bias of Idiosyncratic Firm and Bank Shocks

We can now derive the bias of firm and bank FEs defining the DGP as $C = X\beta + \Delta + \Gamma + U$. Let Δ_i (Γ_b) be the i th (b th) column of Δ (Γ) and Δ_{-i} (Γ_{-b}) the matrix containing all the columns of Δ (Γ) but the i th (b th). Following the same derivations as above, we have

$$\begin{aligned}
D'_i U &= D'_i \phi A M [X\beta + \Delta_i + \Delta_{-i} + \Gamma] + (M + I)\epsilon = \delta_i \sum_{k=1}^{\infty} \phi^k D'_i A^k D_i \\
&= \delta_i \sum_{k=1}^{\infty} \phi^k \mathbf{1} = \delta_i \left(\sum_{k \text{ even}} \phi^k l_k^i - \sum_{k \text{ odd}} \phi^k l_k^i \right) = \delta_i (E_i - O_i) \tag{A.11}
\end{aligned}$$

where l_k^i is the number of loops from (to) firm i that involve chains of length k in the credit network.

A similar derivation can be done for $G'_i U$:

$$\begin{aligned}
G'_i U &= D'_i \phi A M [X\beta + \Delta + \Gamma_b + \Gamma_{-b}] + (M + I)\epsilon \\
&= \gamma_b \left(\sum_{k \text{ even}} \phi^k l_k^b - \sum_{k \text{ odd}} \phi^k l_k^b \right) = \gamma_b (E_b - O_b) \tag{A.12}
\end{aligned}$$

where l_k^b is the number of loops from (to) bank b that involve chains of length k in the

credit network. Given that both O_b and E_b are positive, the sign of the bias depends on the credit network topology. In indirect networks, we may expect $E_b > O_b$, as loops generate paths of even length going back and forth at each step.

A.2 Monte Carlo Tables

Table A.1: Simulation study - ICM treatment effect bias under different share of treated, spillovers and density

		m=2		m=4		m=6		m=8		m=10	
$\phi = \rho$	% of treated	mean bias	mse	mean bias	mse	mean bias	mse	mean bias	mse	mean bias	mse
-0.2	0.10	-0.059	0.124	-0.198	0.110	-0.299	0.143	-0.390	0.199	-0.456	0.251
	0.25	-0.007	0.036	-0.037	0.026	-0.080	0.029	-0.117	0.031	-0.155	0.039
	0.50	0.142	0.035	0.238	0.067	0.295	0.095	0.314	0.106	0.328	0.114
	0.75	0.266	0.079	0.516	0.272	0.660	0.441	0.749	0.565	0.813	0.664
	0.90	0.339	0.123	0.685	0.473	0.879	0.774	1.006	1.014	1.095	1.201
-0.3	0.10	-0.253	0.213	-0.660	0.574	-0.958	1.016	-1.206	1.546	-1.352	1.909
	0.25	-0.105	0.063	-0.397	0.208	-0.588	0.388	-0.766	0.629	-0.901	0.848
	0.50	0.093	0.031	0.067	0.026	0.013	0.020	-0.064	0.022	-0.113	0.029
	0.75	0.313	0.110	0.517	0.277	0.601	0.369	0.642	0.421	0.661	0.444
	0.90	0.436	0.197	0.791	0.630	0.961	0.929	1.056	1.118	1.133	1.288
-0.4	0.10	-0.622	0.653	-2.029	4.629	-3.048	9.756	-3.766	14.635	-4.245	18.382
	0.25	-0.410	0.280	-1.512	2.479	-2.334	5.623	-2.890	8.530	-3.295	11.025
	0.50	-0.052	0.043	-0.638	0.491	-1.117	1.337	-1.457	2.199	-1.723	3.042
	0.75	0.273	0.095	0.201	0.073	0.081	0.042	-0.032	0.032	-0.099	0.040
	0.90	0.478	0.238	0.728	0.541	0.807	0.665	0.840	0.718	0.867	0.765

Notes. The mean bias and the MSE are computed across the 500 simulated samples. The number of nodes N in the network is 200, 100 firms and 100 banks, in each sample. The column spillovers intensity reports the value of $\phi = \rho$ in the simulations.

Table A.2: Simulation study - ICM treatment effect bias with FEs under different share of treated, spillovers and density

$\phi = \rho$	% of treated	m=2		m=4		m=6		m=8		m=10	
		mean bias	mse	mean bias	mse	mean bias	mse	mean bias	mse	mean bias	mse
-0.2	0.10	-0.085	0.121	-0.332	0.202	-0.855	0.872	-1.119	1.373	-1.221	1.574
	0.25	0.000	0.038	-0.211	0.083	-0.753	0.627	-1.073	1.196	-1.204	1.489
	0.50	0.136	0.037	0.028	0.019	-0.586	0.378	-0.990	1.020	-1.169	1.399
	0.75	0.263	0.078	0.312	0.108	-0.254	0.095	-0.755	0.615	-1.077	1.200
	0.90	0.340	0.123	0.512	0.269	0.125	0.042	-0.389	0.212	-0.790	0.684
-0.3	0.10	-0.284	0.217	-0.896	0.994	-1.936	4.023	-2.529	6.666	-2.785	7.962
	0.25	-0.116	0.064	-0.663	0.522	-1.776	3.282	-2.433	6.035	-2.773	7.776
	0.50	0.112	0.035	-0.271	0.114	-1.443	2.182	-2.255	5.170	-2.663	7.165
	0.75	0.306	0.104	0.235	0.075	-0.816	0.748	-1.833	3.480	-2.430	6.000
	0.90	0.439	0.201	0.564	0.328	-0.142	0.084	-1.099	1.347	-1.933	3.900
-0.4	0.10	-0.552	0.564	-2.471	6.830	-5.149	27.597	-6.650	45.138	-7.313	54.346
	0.25	-0.385	0.245	-2.082	4.646	-4.786	23.521	-6.530	43.221	-7.351	54.391
	0.50	-0.069	0.043	-1.222	1.657	-4.062	16.962	-6.022	36.735	-7.067	50.276
	0.75	0.277	0.094	-0.287	0.146	-2.669	7.507	-5.097	26.517	-6.503	42.740
	0.90	0.480	0.240	0.418	0.204	-1.162	1.610	-3.363	11.971	-5.331	29.257

Notes. The mean bias and the MSE are computed across the 500 simulated samples. The number of nodes N in the network is 200, 100 firms and 100 banks, in each sample. The column spillovers intensity reports the value of $\phi = \rho$ in the simulations.

Table A.3: Simulation study - CNM estimator performance under different share of treated, spillovers and density

$\phi = \rho$	% of treated	m=2		m=4		m=6		m=8		m=10	
		mean bias	mse	mean bias	mse	mean bias	mse	mean bias	mse	mean bias	mse
-0.2	0.10	0.058	0.118	0.022	0.059	0.013	0.039	0.000	0.030	-0.006	0.024
	0.25	0.006	0.034	0.002	0.017	0.005	0.013	0.009	0.009	0.006	0.008
	0.50	0.020	0.017	0.002	0.009	0.005	0.007	0.003	0.005	0.002	0.004
	0.75	0.010	0.015	0.004	0.009	0.013	0.007	0.006	0.005	0.005	0.004
	0.90	0.007	0.021	0.018	0.015	0.013	0.011	0.007	0.008	0.007	0.007
-0.3	0.10	0.009	0.098	0.016	0.058	0.000	0.045	0.000	0.032	0.011	0.033
	0.25	0.012	0.035	0.005	0.020	0.009	0.013	0.007	0.011	0.004	0.009
	0.50	0.010	0.016	-0.004	0.008	-0.001	0.006	-0.002	0.005	0.002	0.004
	0.75	0.009	0.015	0.003	0.007	0.004	0.005	0.008	0.003	-0.002	0.003
	0.90	0.012	0.014	0.011	0.010	0.012	0.007	0.000	0.005	0.008	0.005
-0.4	0.10	0.019	0.096	0.014	0.070	-0.001	0.058	0.019	0.045	0.007	0.040
	0.25	-0.009	0.034	0.001	0.022	-0.006	0.017	-0.001	0.011	0.009	0.012
	0.50	0.013	0.017	0.004	0.009	0.004	0.006	-0.001	0.005	0.000	0.004
	0.75	0.008	0.011	0.004	0.005	0.001	0.004	-0.001	0.000	0.000	0.002
	0.90	0.003	0.011	0.005	0.006	0.004	0.004	0.003	0.003	0.003	0.002

Notes. The mean bias and the MSE are computed across the 500 simulated samples. The number of nodes n in the network is 200, 100 firms and 100 banks, in each sample. The first column reports the spillovers intensity, i.e. the value of $\phi = \rho$ in the simulations.

Table A.4: Simulation study - CNM estimator performance with FEs under different share of treated, spillovers and density

$\phi = \rho$	% of treated	m=2		m=4		m=6		m=8		m=10	
		mean bias	mse	mean bias	mse	mean bias	mse	mean bias	mse	mean bias	mse
-0.2	0.10	0.025	0.120	0.037	0.069	0.008	0.139	-0.010	0.309	0.018	0.547
	0.25	0.014	0.034	0.005	0.023	0.027	0.045	0.018	0.091	0.012	0.209
	0.50	0.015	0.020	0.005	0.011	0.016	0.021	-0.001	0.040	0.009	0.081
	0.75	0.005	0.017	0.003	0.010	0.005	0.013	-0.010	0.026	-0.016	0.056
	0.90	0.003	0.022	0.007	0.016	0.007	0.014	-0.013	0.021	0.015	0.039
-0.3	0.10	0.002	0.098	0.011	0.072	0.015	0.146	0.035	0.309	0.033	0.835
	0.25	0.011	0.031	0.016	0.025	0.007	0.057	-0.003	0.116	-0.031	0.234
	0.50	0.016	0.018	-0.002	0.011	0.004	0.024	-0.009	0.052	-0.012	0.107
	0.75	0.001	0.012	0.014	0.008	0.006	0.015	0.002	0.032	-0.019	0.058
	0.90	0.016	0.018	0.005	0.009	0.003	0.013	-0.008	0.026	-0.018	0.049
-0.4	0.10	0.048	0.130	0.007	0.081	0.008	0.164	0.004	0.333	0.007	0.675
	0.25	0.017	0.034	-0.006	0.029	0.006	0.053	0.001	0.126	0.030	0.220
	0.50	0.004	0.016	0.011	0.012	-0.006	0.029	0.016	0.049	-0.029	0.102
	0.75	0.008	0.011	0.004	0.008	-0.001	0.018	-0.004	0.035	0.006	0.087
	0.90	0.015	0.014	0.003	0.003	0.012	0.015	-0.005	0.034	-0.005	0.065

Notes. The mean bias and the MSE are computed across the 500 simulated samples. The number of nodes n in the network is 200, 100 firms and 100 banks, in each sample. The first column reports the spillovers intensity, i.e. the value of $\phi = \rho$ in the simulations.

Table A.5: Simulation study - estimator performance under different spillovers, size and density

n	m	true		true		true		true		true		
		ϕ	ρ	ϕ	ρ	ϕ	ρ	ϕ	ρ	ϕ	ρ	
200		-0.1	-0.1	-0.1	-0.2	-0.1	-0.3	-0.1	-0.4	-0.4	-0.4	
	2	mean	-0.097	-0.100	-0.100	-0.209	-0.101	-0.306	-0.093	-0.414	-0.406	-0.406
		std	0.084	0.087	0.090	0.089	0.082	0.081	0.082	0.076	0.066	0.067
	4	mean	-0.099	-0.096	-0.095	-0.194	-0.094	-0.299	-0.096	-0.395	-0.402	-0.395
		std	0.040	0.041	0.040	0.040	0.040	0.044	0.040	0.042	0.039	0.040
	6	mean	-0.098	-0.098	-0.097	-0.197	-0.097	-0.295	-0.096	-0.398	-0.402	-0.395
		std	0.029	0.030	0.029	0.030	0.029	0.033	0.028	0.032	0.039	0.040
	8	mean	-0.101	-0.097	-0.101	-0.197	-0.100	-0.296	-0.098	-0.398	-0.403	-0.396
		std	0.024	0.022	0.025	0.026	0.024	0.026	0.024	0.028	0.033	0.032
	10	mean	-0.102	-0.096	-0.099	-0.198	0.000	-0.297	-0.098	-0.398	-0.402	-0.397
		std	0.021	0.020	0.023	0.023	0.022	0.024	0.021	0.024	0.031	0.030
800												
	2	mean	-0.102	-0.098	-0.100	-0.201	-0.097	-0.301	-0.097	-0.401	-0.398	-0.401
		std	0.041	0.043	0.044	0.042	0.042	0.042	0.040	0.037	0.034	0.033
	4	mean	-0.098	-0.098	-0.099	-0.198	-0.098	-0.300	-0.097	-0.400	-0.398	-0.401
		std	0.022	0.021	0.021	0.022	0.021	0.022	0.018	0.020	0.021	0.020
	6	mean	-0.099	-0.100	-0.099	-0.200	-0.097	-0.300	-0.099	-0.400	-0.398	-0.401
		std	0.015	0.014	0.015	0.016	0.014	0.016	0.014	0.016	0.021	0.020
	8	mean	-0.100	-0.099	-0.100	-0.200	-0.099	-0.300	-0.099	-0.399	-0.401	-0.399
		std	0.012	0.012	0.012	0.014	0.013	0.014	0.012	0.014	0.015	0.016
	10	mean	-0.100	-0.099	-0.100	-0.200	-0.099	-0.300	-0.099	-0.400	-0.399	-0.400
		std	0.010	0.011	0.010	0.011	0.011	0.012	0.010	0.012	0.015	0.015

Notes. The mean and the std are computed across the 500 simulated samples. n is the number of nodes in the network, m regulates the network density as described in Section 4.1.

Idiosyncratic Firm and Bank Shocks, Detailed Exposition. Here we deepen the properties firm and bank FEs' estimates, which account for idiosyncratic shocks, when using the isolated credit model (ICM) and the credit network model (CNM). To better assess the magnitude of the bias in finite samples, we let all FEs be positive by adding the minimum draw for each replication. To aggregate all the firms' and banks' specific parameters, we use the mean bias, the mean absolute bias and the root mean squared bias: $MB = 1/R \sum_{r=1}^R \{1/N [\sum_{j=1}^F (\hat{\delta}_j^r - \delta) + \sum_{k=1}^B (\hat{\gamma}_k^r - \gamma)]\}$, $MAB = 1/R \sum_{r=1}^R [1/N (\sum_{j=1}^F |\hat{\delta}_j^r - \delta| + \sum_{k=1}^B |\hat{\gamma}_k^r - \gamma|)]$, $RMSE = 1/R \sum_{r=1}^R [1/N \sqrt{\sum_{j=1}^F (\hat{\delta}_j^r - \delta)^2 + \sum_{k=1}^B (\hat{\gamma}_k^r - \gamma)^2}]$.

Table A.6 reports these indicators for different network sizes ($n = 200, 800$ and 2000), network densities ($m = 4, 6, 8$ and 10) and magnitude of spillovers (ϕ and ρ). The bias for the ICM is increasing in both the density of the network and the magnitude of spillovers, while the bias of the CNM is always close to zero and converges to it as n tends to infinity.

Given that $\theta = 0.1$ in our pivotal setting and we constrained the FEs to be positive, the average FE is greater than 0.3 with a probability lower than 0.001. The bias of the ICM ranges from about 0.3 (when $m = 4$ and $\phi = \rho = -0.2$) and 3 (when $m = 10$ and $\phi = \rho = -0.4$), which means that with low (high) density and small (large) spillovers the ICM estimate is on average about the double (ten times) the real idiosyncratic shock with very high probability. The intuition behind this result is that competitive interactions let positive idiosyncratic shocks diffuse through the credit network and amplify them, overestimating them, if such feedback loops are not accounted for.

Table A.6: Simulation study - ICM and CNM estimators performance under different spillovers, size and density

n	m	$\phi = \rho = -0.2$			$\phi = \rho = -0.3$			$\phi = \rho = -0.4$				
		MB	MAB	RMSE	MB	MAB	RMSE	MB	MAB	RMSE		
200	4	ICM	0.303	0.394	61	0.412	0.460	73	0.628	0.648	98	
		CNM	0.002	0.418	67	0.002	0.424	71	-0.005	0.420	69	
	6	ICM	0.435	0.532	627	0.676	0.739	843	1.283	1.330	1,446	
		CNM	-0.005	0.447	556	-0.009	0.448	556	-0.009	0.447	547	
	8	ICM	0.650	0.732	1,472	1.066	1.125	2,163	2.230	2.281	4,316	
		CNM	-0.013	0.472	1,025	0.001	0.466	1,002	-0.014	0.462	995	
	10	ICM	0.820	0.891	2,234	1.414	1.462	3,599	3.074	3.138	7,628	
		CNM	-0.027	0.498	1,376	-0.021	0.494	1,376	-0.018	0.474	1,312	
	800	4	ICM	0.276	0.377	484	0.376	0.429	527	0.594	0.620	729
			CNM	0.000	0.414	543	-0.002	0.415	535	0.001	0.417	539
		6	ICM	0.412	0.516	4,851	0.642	0.710	6,427	1.253	1.305	11,317
			CNM	-0.002	0.437	4,299	-0.002	0.436	4,284	-0.002	0.435	4,257
8		ICM	0.615	0.701	11,235	1.039	1.100	16,944	2.195	2.256	34,081	
		CNM	-0.004	0.448	7,707	0.000	0.448	7,664	-0.002	0.449	7,710	
10		ICM	0.789	0.862	17,467	1.374	1.430	28,187	3.015	3.081	59,562	
		CNM	-0.007	0.458	10,089	-0.002	0.454	10,007	-0.009	0.457	10,041	
2000		4	ICM	0.264	0.371	1,882	0.360	0.419	2,052	0.577	0.604	2,821
			CNM	0.000	0.415	2,144	-0.002	0.413	2,118	0.002	0.414	2,127
		6	ICM	0.399	0.507	18,845	0.632	0.703	25,099	1.238	1.295	44,536
			CNM	0.001	0.434	16,831	0.000	0.434	16,802	-0.001	0.435	16,858
	8	ICM	0.602	0.691	43,662	1.023	1.086	66,298	2.171	2.231	133,100	
		CNM	-0.001	0.444	30,017	0.000	0.443	30,005	0.000	0.444	30,024	
	10	ICM	0.775	0.851	68,219	1.354	1.412	109,951	3.001	3.069	234,902	
		CNM	0.001	0.450	39,115	-0.005	0.450	39,012	-0.004	0.448	38,915	

Notes. ICM and CNM stand respectively for isolated credit model and credit network model. n is the number of nodes in the network, m regulates the network density as described in Section 4.1. MB, MAB and RMSE stand respectively for mean bias, mean absolute bias and root mean square error. All are averaged across 500 replications and computed with the following formulas: $MB = 1/R \sum_{r=1}^R \{1/N[\sum_{j=1}^F (\hat{\delta}_j^r - \delta) + \sum_{k=1}^B (\hat{\gamma}_k^r - \gamma)]\}$, $MAB = 1/R \sum_{r=1}^R [1/N(\sum_{j=1}^F |\hat{\delta}_j^r - \delta| + \sum_{k=1}^B |\hat{\gamma}_k^r - \gamma|)]$, $RMSE = 1/R \sum_{r=1}^R [1/N \sqrt{\sum_{j=1}^F (\hat{\delta}_j^r - \delta)^2 + \sum_{k=1}^B (\hat{\gamma}_k^r - \gamma)^2}]$. $RMSE$ is reported without decimals.

Table A.7: Simulation study - ICM and CNM estimators performance with real credit networks

$\phi = \rho$		n =	mean bias						
			400		800		2000		
			treatment effect	idiosyncratic shocks	treatment effect	idiosyncratic shocks	treatment effect	idiosyncratic shocks	
-0.2	0.10	ICN	-0.589	-0.008	-0.766	-0.042	-0.918	-0.068	
		CNM	0.039	-0.003	0.004	-0.001	0.000	0.001	
	0.50	ICN	-0.445	0.432	-0.628	0.341	-0.852	0.274	
		CNM	0.007	-0.006	0.008	-0.005	0.001	-0.001	
	0.90	ICN	-0.039	0.498	-0.060	0.574	-0.494	0.498	
		CNM	0.001	-0.002	0.001	-0.002	0.003	-0.001	
-0.3	0.10	ICN	-1.244	0.025	-1.695	-0.012	-2.076	-0.045	
		CNM	0.024	-0.002	0.017	0.000	0.011	-0.002	
	0.50	ICN	-1.002	0.642	-1.438	0.556	-1.930	0.491	
		CNM	0.004	-0.002	0.005	-0.002	0.006	-0.002	
	0.90	ICN	-0.242	0.836	-0.266	0.764	-1.219	0.780	
		CNM	0.006	-0.005	0.005	-0.004	0.002	0.001	
	-0.4	0.10	ICN	-3.080	0.126	-4.364	0.104	-5.479	0.066
			CNM	0.012	-0.001	0.004	0.000	0.006	-0.001
		0.50	ICN	-2.484	1.135	-3.772	1.146	-5.130	1.111
			CNM	0.001	0.002	0.004	-0.001	0.005	-0.003
		0.90	ICN	-0.880	1.252	-1.691	1.279	-3.393	1.557
			CNM	0.004	-0.001	0.002	0.000	0.000	-0.001

Notes. ICM and CNM stand respectively for isolated credit model and credit network model. n is the number of nodes in the network. The links are extracted from realized credit relationships between a random sample of firms and banks from all credit relationships observed in 2016. The bias of the treatment effect is computed as in Table A.1, A.2, A.4 and 1. The bias of idiosyncratic shocks is computed as in Table A.6.

References

- Acharya, Viral V and Sascha Steffen. 2020. “The risk of being a fallen angel and the corporate dash for cash in the midst of COVID.” *The Review of Corporate Finance Studies* 9 (3):430–471.
- Agarwal, Sumit and Robert Hauswald. 2010. “Distance and private information in lending.” *The Review of Financial Studies* 23 (7):2757–2788.
- Alfaro, Laura, Manuel García-Santana, and Enrique Moral-Benito. 2021. “On the direct and indirect real effects of credit supply shocks.” *Journal of Financial Economics* 139 (3):895–921.
- Altavilla, Carlo, Miguel Boucinha, and Paul Bouscasse. 2022. “Supply or Demand: What Drives Fluctuations in the Bank Loan Market?” .
- Amiti, Mary and David E. Weinstein. 2018. “How much do idiosyncratic bank shocks affect investment? Evidence from matched bank-firm loan data.” *Journal of Political Economy* 126 (2):525–587.
- Andreeva, Desislava C and Miguel García-Posada. 2021. “The impact of the ECB’s targeted long-term refinancing operations on banks’ lending policies: The role of competition.” *Journal of Banking & Finance* 122:105992.
- Angelucci, Manuela and Giacomo De Giorgi. 2009. “Indirect effects of an aid program: How do cash transfers affect ineligibles’ consumption?” *American Economic Review* 99 (1):486–508.
- Anselin, Luc and Nancy Lozano-Gracia. 2008. “Errors in variables and spatial effects in hedonic house price models of ambient air quality.” *Empirical economics* 34:5–34.
- Arduini, Tiziano, Eleonora Patacchini, and Edoardo Rainone. 2015. “Parametric and semiparametric IV estimation of network models with selectivity.” *Einaudi Institute for Economics and Finance (EIEF) Working paper* 15/9.
- . 2020. “Identification and estimation of network models with heterogeneous interactions.” In *The Econometrics of Networks*. Emerald Publishing Limited.
- Arellano, Manuel and Bo Honoré. 2001. “Panel data models: some recent developments.” In *Handbook of econometrics*, vol. 5. Elsevier, 3229–3296.
- Auerbach, Eric. 2022. “Identification and estimation of a partially linear regression model using network data.” *Econometrica* 90 (1):347–365.

- Ballester, Coralio, Antoni Calvó-Armengol, and Yves Zenou. 2006. “Who’s who in networks. Wanted: The key player.” *Econometrica* 74 (5):1403–1417.
- Bartik, Timothy J. 1991. “Who benefits from state and local economic development policies?” .
- Bartoli, Francesca, Giovanni Ferri, Pierluigi Murro, and Zeno Rotondi. 2013. “SME financing and the choice of lending technology in Italy: Complementarity or substitutability?” *Journal of Banking & Finance* 37 (12):5476–5485.
- Battaglini, Marco, Forrest W Crawford, Eleonora Patacchini, and Sida Peng. 2020. “A Graphical Lasso Approach to Estimating Network Connections: The Case of US Lawmakers.” Tech. rep., National Bureau of Economic Research.
- Battaglini, Marco, Eleonora Patacchini, and Edoardo Rainone. 2022. “Endogenous social interactions with unobserved networks.” *The Review of Economic Studies* 89 (4):1694–1747.
- Behn, Markus, Rainer Haselmann, and Paul Wachtel. 2016. “Procyclical capital regulation and lending.” *The Journal of Finance* 71 (2):919–956.
- Benetton, Matteo. 2021. “Leverage regulation and market structure: A structural model of the uk mortgage market.” *The Journal of Finance* 76 (6):2997–3053.
- Benetton, Matteo and Davide Fantino. 2021. “Targeted monetary policy and bank lending behavior.” *Journal of Financial Economics* 142 (1):404–429.
- Berg, Tobias, Markus Reisinger, and Daniel Streitz. 2021. “Spillover effects in empirical corporate finance.” *Journal of Financial Economics* 142 (3):1109–1127.
- Berg, Tobias and Daniel Streitz. 2019. “Real effects of search frictions in consumer credit markets.” Available at SSRN 3377457. URL https://papers.ssrn.com/sol3/papers.cfm?abstract_id=3377457.
- Berger, Allen N and Gregory F Udell. 1995. “Relationship lending and lines of credit in small firm finance.” *Journal of business* :351–381.
- Berlin, Mitchell and Loretta J Mester. 1998. “On the profitability and cost of relationship lending.” *Journal of Banking & Finance* 22 (6-8):873–897.
- Bharath, Sreedhar, Sandeep Dahiya, Anthony Saunders, and Anand Srinivasan. 2007. “So what do I get? The bank’s view of lending relationships.” *Journal of Financial Economics* 85 (2):368–419.

- Blanchard, Olivier Jean, Lawrence F Katz, Robert E Hall, and Barry Eichengreen. 1992. “Regional evolutions.” *Brookings papers on economic activity* 1992 (1):1–75.
- Bolton, Patrick, Xavier Freixas, Leonardo Gambacorta, and Paolo Emilio Mistrulli. 2016. “Relationship and transaction lending in a crisis.” *The Review of Financial Studies* 29 (10):2643–2676.
- Bonaccorsi di Patti, Emilia and Enrico Sette. 2016. “Did the securitization market freeze affect bank lending during the financial crisis? Evidence from a credit register.” *Journal of Financial Intermediation* 25:54–76.
- Borusyak, Kirill, Peter Hull, and Xavier Jaravel. 2022. “Quasi-experimental shift-share research designs.” *The Review of Economic Studies* 89 (1):181–213.
- Boss, Michael, Helmut Elsinger, Martin Summer, and Stefan Thurner. 2004. “Network topology of the interbank market.” *Quantitative Finance* 4 (6):677–684.
- Bramoullé, Y., H. Djebbari, and B. Fortin. 2009. “Identification of peer effects through social networks.” *Journal of Econometrics* 150:41–55.
- Bramoullé, Yann, Habiba Djebbari, and Bernard Fortin. 2020. “Peer effects in networks: A survey.” *Annual Review of Economics* 12:603–629.
- Branzoli, Nicola, Edoardo Rainone, and Ilaria Supino. 2023. “The role of banks’ technology adoption in credit markets during the pandemic.” *Journal of Financial Stability, forthcoming* .
- Bripi, Francesco. 2021. “Substituting banks: Estimating credit demand in Italy.” *Unpublished Manuscript* .
- Calvó-Armengol, Antoni, Eleonora Patacchini, and Yves Zenou. 2009. “Peer effects and social networks in education.” *The Review of Economic Studies* 76 (4):1239–1267.
- Chodorow-Reich, Gabriel. 2014. “The employment effects of credit market disruptions: Firm-level evidence from the 2008–9 financial crisis.” *The Quarterly Journal of Economics* 129 (1):1–59.
- Chodorow-Reich, Gabriel and Antonio Falato. 2022. “The loan covenant channel: How bank health transmits to the real economy.” Tech. Rep. 1.
- Cingano, Federico, Francesco Manaresi, and Enrico Sette. 2016. “Does credit crunch investment down? New evidence on the real effects of the bank-lending channel.” *The Review of Financial Studies* 29 (10):2737–2773.

- Corbae, Dean and Pablo D’Erasmus. 2021. “Capital buffers in a quantitative model of banking industry dynamics.” *Econometrica* 89 (6):2975–3023.
- Core, Fabrizio and Filippo De Marco. 2021. “Public Guarantees for Small Businesses in Italy during COVID-19.” *C.E.P.R. Discussion Papers* (15799).
- Craig, Ben and Goetz Von Peter. 2014. “Interbank tiering and money center banks.” *Journal of Financial Intermediation* 23 (3):322–347.
- Dall’Erba, Sandy and Julie Le Gallo. 2008. “Regional convergence and the impact of European structural funds over 1989–1999: A spatial econometric analysis.” *Papers in Regional Science* 87 (2):219–244.
- Darmouni, Olivier and Andrew Sutherland. 2021. “Learning about competitors: Evidence from SME lending.” *The Review of Financial Studies* 34 (5):2275–2317.
- De Jonghe, Olivier, Hans Dewachter, Klaas Mulier, Steven Ongena, and Glenn Schepens. 2020. “Some borrowers are more equal than others: Bank funding shocks and credit reallocation.” *Review of Finance* 24 (1):1–43.
- De Paula, Áureo. 2020. “Econometric models of network formation.” *Annual Review of Economics* 12:775–799.
- De Paula, Áureo, Imran Rasul, and Pedro Souza. 2019. “Identifying network ties from panel data: Theory and an application to tax competition.” *arXiv preprint arXiv:1910.07452* .
- Degryse, Hans and Steven Ongena. 2005. “Distance, lending relationships, and competition.” *The Journal of Finance* 60 (1):231–266.
- Dewally, Michaël and Yingying Shao. 2014. “Liquidity crisis, relationship lending and corporate finance.” *Journal of Banking & Finance* 39:223–239.
- Federico, Stefano, Fadi Hassan, and Veronica Rappoport. 2023. “Trade shocks and credit reallocation.” Tech. rep., National Bureau of Economic Research.
- Federico, Stefano, Giuseppe Marinelli, and Francesco Palazzo. 2023. “The 2014 Russia shock and its effects on Italian firms and banks.” Tech. rep., National Bureau of Economic Research.
- Fingleton, Bernard and Julie Le Gallo. 2008. “Estimating spatial models with endogenous variables, a spatial lag and spatially dependent disturbances: finite sample properties.” *Papers in Regional Science* 87 (3):319–339.

- Foroughi, Pouyan, Alan J. Marcus, Vinh Nguyen, and Hassan Tehranian. 2022. “Peer effects in corporate governance practices: Evidence from universal demand laws.” *The Review of Financial Studies* 35 (1):132–167.
- Fox, Jeremy T. 2009. “Structural empirical work using matching models.” *New Palgrave Dictionary of Economics* .
- Fruchterman, Thomas MJ and Edward M Reingold. 1991. “Graph drawing by force-directed placement.” *Software: Practice and experience* 21 (11):1129–1164.
- Fuster, Andreas, Paul Goldsmith-Pinkham, Tarun Ramadorai, and Ansgar Walther. 2018. “Predictably Unequal? The Effects of Machine Learning on Credit Markets.” *SSRN Working Paper* (3072038).
- Fuster, Andreas, Matthew Plosser, Philipp Schnabl, and James Vickery. 2019. “The Role of Technology in Mortgage Lending.” *The Review of Financial Studies* 32 (5):1854–1899.
- Gabaix, Xavier and Ralph S. J. Koijen. 2020. “Granular instrumental variables.” Working Paper 28204, National Bureau of Economic Research. URL <https://www.nber.org/papers/w28204>.
- Galaasen, Sigurd, Rustam Jamilov, Ragnar Juelsrud, and H el ene Rey. 2020. “Granular credit risk.” Tech. rep., National Bureau of Economic Research.
- Giannetti, Mariassunta and Farzad Saidi. 2019. “Shock propagation and banking structure.” *The Review of Financial Studies* 32 (7):2499–2540.
- Giometti, Marco and Stefano Pietrosanti. 2022. “Bank specialization and the design of loan contracts.” Working Paper 14, Federal Deposit Insurance Corporation. <https://www.fdic.gov/analysis/cfr/working-papers/2022/cfr-wp2022-14.pdf>.
- Goldsmith-Pinkham, Paul and Guido W Imbens. 2013. “Social networks and the identification of peer effects.” *Journal of Business & Economic Statistics* 31 (3):253–264.
- Goldsmith-Pinkham, Paul, Isaac Sorkin, and Henry Swift. 2020. “Bartik instruments: What, when, why, and how.” *American Economic Review* 110 (8):2586–2624.
- Graham, Bryan and Aureo De Paula. 2020. *The Econometric Analysis of Network Data*. Academic Press.
- Grieser, William, Charles Hadlock, James LeSage, and Morad Zekhnini. 2022. “Network effects in corporate financial policies.” *Journal of Financial Economics* 144 (1):247–272.

- Gupta, Abhimanyu, Sotirios Kokas, Alexander Michaelides, and Raoul Minetti. 2023. “Networks and Information in Credit Markets.” Tech. rep.
- Honore, Bo E, Ekaterini Kyriazidou, and JL Powell. 2000. “Estimation of Tobit-type models with individual specific effects.” *Econometric reviews* 19 (3):341–366.
- Hsieh, Chih-Sheng and Lung Fei Lee. 2016. “A social interactions model with endogenous friendship formation and selectivity.” *Journal of Applied Econometrics* 31 (2):301–319.
- Huber, Kilian. 2018. “Disentangling the effects of a banking crisis: Evidence from German firms and counties.” *American Economic Review* 108 (3):868–98.
- . 2022. “Estimating General Equilibrium Spillovers of Large-Scale Shocks.” *The Review of Financial Studies* :hhac057.
- Huremovic, Kenan, Gabriel Jiménez, Enrique Moral-Benito, Fernando Vega-Redondo, and José-Luis Peydró. 2020. “Production and financial networks in interplay: Crisis evidence from supplier-customer and credit registers.” .
- Ioannidou, Vasso and Steven Ongena. 2010. “Time for a change: Loan conditions and bank behavior when firms switch banks.” *The Journal of Finance* 65 (5):1847–1877.
- Iori, Giulia, Giulia De Masi, Ovidiu Vasile Precup, Giampaolo Gabbi, and Guido Caldarelli. 2008. “A network analysis of the Italian overnight money market.” *Journal of Economic Dynamics and Control* 32 (1):259–278.
- Iyer, Rajkamal, José-Luis Peydró, Samuel da Rocha-Lopes, and Antoinette Schoar. 2013. “Interbank liquidity crunch and the firm credit crunch: Evidence from the 2007–2009 crisis.” *The Review of Financial Studies* 27 (1):347–372.
- Jackson, Matthew O. 2010. *Social and economic networks*. Princeton university press.
- Jackson, Matthew O, Brian W Rogers, and Yves Zenou. 2017. “The economic consequences of social-network structure.” *Journal of Economic Literature* 55 (1):49–95.
- Jackson, Matthew O and Yves Zenou. 2015. “Games on networks.” In *Handbook of game theory with economic applications*, vol. 4. Elsevier, 95–163.
- Jiménez, Gabriel, Atif Mian, José-Luis Peydró, and Jesús Saurina. 2020. “The real effects of the bank lending channel.” *Journal of Monetary Economics* 115:162–179.
- Jiménez, Gabriel, Steven Ongena, José-Luis Peydró, and Jesús Saurina. 2012. “Credit Supply and Monetary Policy: Identifying the Bank Balance-Sheet Channel with Loan Applications.” *The American Economic Review* 102 (5):2301–2326.

- Jiménez, Gabriel, Steven Ongena, José-Luis Peydró, and Jesús Saurina. 2014. “Hazardous Times for Monetary Policy: What Do Twenty-Three Million Bank Loans Say About the Effects of Monetary Policy on Credit Risk-Taking?” *Econometrica* 82 (2):463–505.
- Jiménez, Gabriel, Steven Ongena, José-Luis Peydró, and Jesus Saurina Salas. 2017. “Macroprudential policy, countercyclical bank capital buffers and credit supply: Evidence from the Spanish dynamic provisioning experiments.” *Journal of Political Economy* 125 (6):2126–2177.
- Johnsson, Ida and Hyungsik Roger Moon. 2021. “Estimation of peer effects in endogenous social networks: control function approach.” Tech. Rep. 2.
- Kashyap, Anil K and Jeremy C Stein. 2000. “What do a million observations on banks say about the transmission of monetary policy?” *American Economic Review* 90 (3):407–428.
- Kelejian, H. and I. R. Prucha. 1998. “A generalized spatial two-stage least squares procedure for estimating a spatial autoregressive model with autoregressive disturbances.” *The Journal of Real Estate Finance and Economics* 17 (1):99–121.
- . 1999. “A generalized moments estimator for the autoregressive parameter in a spatial model.” *International Economic Review* 40 (2):509–533.
- Khwaja, Asim Ijaz and Atif Mian. 2008. “Tracing the Impact of Bank Liquidity Shocks: Evidence from an Emerging Market.” *The American Economic Review* 98 (4):1413–1442.
- Kwan, Alan, Chen Lin, Vesa Pursianen, and Mingzhu Tai. 2021. “Stress Testing Banks’ Digital Capabilities: Evidence From the COVID-19 Pandemic.” *SSRN working papers* (3694288).
- Kyriazidou, Ekaterini. 1997. “Estimation of a panel data sample selection model.” *Econometrica: Journal of the Econometric Society* :1335–1364.
- Leary, Mark T. and Michael R. Roberts. 2014. “Do peer firms affect corporate financial policy?” *The Journal of Finance* 69 (1):139–178.
- Lee, L. F. 2004. “Asymptotic Distributions of Quasi-Maximum Likelihood Estimators for Spatial Autoregressive Models.” *Econometrica* 72 (6):1899–1925.
- . 2007. “Identification and estimation of econometric models with group interactions, contextual factors and fixed effects.” *Journal of Econometrics* 140:333–374.
- Lee, L. F., X. Liu, and X. Lin. 2010. “Specification and estimation of social interaction models with network structures.” *The Econometrics Journal* 13:145–176.

- Liu, X. and L. F. Lee. 2010. “GMM estimation of social interaction models with centrality.” *Journal of Econometrics* 159:99–115.
- Liu, Xiaodong, Eleonora Patacchini, and Yves Zenou. 2014. “Endogenous peer effects: local aggregate or local average?” *Journal of Economic Behavior & Organization* 103:39–59.
- Manski, Charles F. 1993. “Identification of endogenous social effects: The reflection problem.” *The Review of Economic Studies* 60 (3):531–542.
- Mian, Atif, Andrés Sarto, and Amir Sufi. 2022. “Estimating general equilibrium multipliers: With application to credit markets.” Tech. rep., Working Paper.
- Miraldo, Marisa, Carol Propper, and Christiern Rose. 2021. “Identification of Peer Effects using Panel Data.” *arXiv preprint arXiv:2108.11545* .
- Paravisini, Daniel. 2008. “Local bank financial constraints and firm access to external finance.” *The Journal of Finance* 63 (5):2161–2193.
- Paravisini, Daniel, Veronica Rappoport, and Philipp Schnabl. 2023. “Specialization in bank lending: Evidence from exporting firms.” *The Journal of Finance* 78 (4):2049–2085.
- Paravisini, Daniel, Veronica Rappoport, Philipp Schnabl, and Daniel Wolfenzon. 2014. “Dissecting the effect of credit supply on trade: Evidence from matched credit-export data.” *The Review of Economic Studies* 82 (1):333–359.
- Patacchini, Eleonora and Edoardo Rainone. 2017. “Social ties and the demand for financial services.” *Journal of Financial Services Research* 52 (1-2):35–88.
- Patacchini, Eleonora, Edoardo Rainone, and Yves Zenou. 2017. “Heterogeneous peer effects in education.” *Journal of Economic Behavior & Organization* 134:190–227.
- Peek, Joe and Eric Rosengren. 1995. “Bank regulation and the credit crunch.” *Journal of Banking & Finance* 19 (3):679–692.
- Peek, Joe and Eric S. Rosengren. 2000. “Collateral damage: Effects of the Japanese bank crisis on real activity in the United States.” *The American Economic Review* 90 (1):30–45.
- Petersen, Mitchell A and Raghuram G Rajan. 1994. “The benefits of lending relationships: Evidence from small business data.” *The journal of finance* 49 (1):3–37.
- Qu, Xi and Lung-fei Lee. 2015. “Estimating a spatial autoregressive model with an endogenous spatial weight matrix.” *Journal of Econometrics* 184 (2):209–232.

- Rainone, Edoardo. 2020a. “Estimating Spillover Effects with Bilateral Outcomes.” In *The Econometrics of Networks*. Emerald Publishing Limited.
- . 2020b. “The network nature of over-the-counter interest rates.” *Journal of Financial Markets* 47:100525.
- Robinson, Peter M. 1988. “Root-N-consistent semiparametric regression.” *Econometrica: Journal of the Econometric Society* :931–954.
- Sanderson, Eleanor and Frank Windmeijer. 2016. “A weak instrument F-test in linear IV models with multiple endogenous variables.” *Journal of econometrics* 190 (2):212–221.
- Schnabl, Philipp. 2012. “The international transmission of bank liquidity shocks: Evidence from an emerging market.” *The Journal of Finance* 67 (3):897–932.
- Schwert, Michael. 2018. “Bank capital and lending relationships.” *The Journal of Finance* 73 (2):787–830.
- Sette, Enrico and Giorgio Gobbi. 2015. “Relationship lending during a financial crisis.” *Journal of the European Economic Association* 13 (3):453–481.
- Silva, André F. 2019. “Strategic liquidity mismatch and financial sector stability.” *The Review of Financial Studies* 32 (12):4696–4733.
- Soramäki, Kimmo, Morten L Bech, Jeffrey Arnold, Robert J Glass, and Walter E Beyeler. 2007. “The topology of interbank payment flows.” *Physica A: Statistical Mechanics and its Applications* 379 (1):317–333.
- Stock, James H and Motohiro Yogo. 2002. “Testing for weak instruments in linear IV regression.”

RECENTLY PUBLISHED “TEMI” (*)

- N. 1409 – *EU structural funds and GDP per capita: spatial var evidence for the European regions*, by Sergio Destefanis and Valter Di Giacinto (April 2023).
- N. 1410 – *Should inequality factor into central banks’ decisions?*, by Niels-Jakob H. Hansen, Alessandro Lin and Rui C. Mano (April 2023).
- N. 1411 – *The anatomy of labor cost adjustment to demand shocks: Germany and Italy during the Great Recession*, by Francesco D’Amuri, Salvatore Lattanzio and Benjamin S. Smith (June 2023).
- N. 1412 – *Quantitative easing, accounting and prudential frameworks, and bank lending*, by Andrea Orame, Rodney Ramcharan and Roberto Robatto (June 2023).
- N. 1413 – *How costly are cartels?*, by Flavien Moreau and Ludovic Panon (June 2023).
- N. 1414 – *Trade in services related to intangibles and the profit shifting hypothesis*, by Nadia Accoto, Stefano Federico and Giacomo Oddo (June 2023).
- N. 1415 – *Currency risk premiums redux*, by Federico Calogero Nucera, Lucio Sarno and Gabriele Zinna (July 2023).
- N. 1416 – *The external financial spillovers of CBDCs*, by Alessandro Moro and Valerio Nispi Landi (July 2023).
- N. 1417 – *Parental retirement and fertility decisions across family policy regimes*, by Edoardo Frattola (July 2023).
- N. 1418 – *Entry, exit, and market structure in a changing climate*, by Michele Cascarano, Filippo Natoli and Andrea Petrella (July 2023).
- N. 1419 – *Temperatures and search: evidence from the housing market*, by Michele Cascarano and Filippo Natoli (July 2023).
- N. 1420 – *Flight to climatic safety: local natural disasters and global portfolio flows*, by Fabrizio Ferriani, Andrea Gazzani and Filippo Natoli (July 2023).
- N. 1421 – *The effects of the pandemic on households’ financial savings: a Bayesian structural VAR analysis*, by Luigi Infante, Francesca Lilla and Francesco Vercelli (October 2023).
- N. 1422 – *Decomposing the monetary policy multiplier*, by Piergiorgio Alessandri, Fabrizio Venditti and Oscar Jordà (October 2023).
- N. 1423 – *The short and medium term effects of full-day schooling on learning and maternal labor supply*, by Giulia Bovini, Nicolò Cattadori, Marta De Philippis and Paolo Sestito (October 2023).
- N. 1424 – *Subsidizing business entry in competitive credit markets*, by Vincenzo Cuciniello, Claudio Michelacci and Luigi Paciello (October 2023).
- N. 1425 – *Drivers of large recessions and monetary policy responses*, by Giovanni Melina and Stefania Villa (October 2023).
- N. 1426 – *The performance of household-held mutual funds: evidence from the euro area*, by Valerio Della Corte and Raffaele Santioni
- N. 1427 – *Trade in the time of COVID-19: an empirical analysis based on Italian data*, by Raffaele De Marchi and Alessandro Moro (March 2023).
- N. 1428 – *Natural gas and the macroeconomy: not all energy shocks are alike*, by Nicola Branzoli, Edoardo Rainone and Iliaria Supino (March 2023).
- N. 1429 – *Inflation is not equal for all: the heterogenous effects of energy shocks*, by Filippo Natoli (March 2023).
- N. 1430 – *Labor market dynamics and geographical reallocations*, by Gaetano Basso, Salvatore Lo Bello and Francesca Subioli

(*) Requests for copies should be sent to:

Banca d’Italia – Servizio Studi di struttura economica e finanziaria – Divisione Biblioteca e Archivio storico – Via Nazionale, 91 – 00184 Rome – (fax 0039 06 47922059). They are available on the Internet www.bancaditalia.it.

2021

- ACCETTURO A., A. LAMORGESE, S. MOCETTI and D. PELLEGRINO, *Housing Price elasticity and growth: evidence from Italian cities*, Journal of Economic Geography, v. 21, 3, pp. 367-396, **WP 1267 (March 2020)**.
- AFFINITO M. and M. PIAZZA, *Always look on the bright side? Central counterparties and interbank markets during the financial crisis*, International Journal of Central Banking, v. 17, 1, pp. 231-283, **WP 1181 (July 2018)**.
- ALBANESE G., E. CIANI and G. DE BLASIO, *Anything new in town? The local effects of urban regeneration policies in Italy*, Regional Science and Urban Economics, v. 86, **WP 1214 (April 2019)**.
- ALBANESE G., G. DE BLASIO and A. LOCATELLI, *Does EU regional policy promote local TFP growth? Evidence from the Italian Mezzogiorno*, Papers in Regional Science, v. 100, 2, pp. 327-348, **WP 1253 (December 2019)**.
- ALBERTAZZI A., A. NOBILI and F. M. SIGNORETTI, *The bank lending channel of conventional and unconventional monetary policy*, Journal of Money, Credit and Banking, v. 53, 2-3, pp. 261-299, **WP 1094 (Jan 2017)**.
- ANZUINI A. and L. ROSSI, *Fiscal policy in the US: a new measure of uncertainty and its effects on the American economy*, Empirical Economics, v. 61, 6, pp. 2613-2634, **WP 1197 (November 2018)**.
- APRIGLIANO V. and D. LIBERATI, *Using credit variables to date business cycle and to estimate the probabilities of recession in real time*, The Manchester School, v. 89, 51, pp. 76-96, **WP 1229 (July 2019)**.
- AUER S., M. BERNARDINI and M. CECIONI, *Corporate leverage and monetary policy effectiveness in the euro area*, European Economic Review, v. 140, Article 103943, **WP 1258 (December 2019)**.
- BANERJEE R, L. GAMBACORTA and E. SETTE, *The real effects of relationship lending*, Journal of Financial Intermediation, v. 48, Article 100923, **WP 1133 (September 2017)**.
- BAMIEH O and A. CINTOLESI, *Intergenerational Transmission in regulated professions and the role of familism*, Journal of Economic Behavior & Organization, v. 192, pp. 857-879, **WP 1350 (October 2021)**.
- BARONE G., F. DAVID, G. DE BLASIO and S. MOCETTI, *How do house prices respond to mortgage supply?*, Journal of Economic Geography, v. 21, 1, pp.127-140, **WP 1282 (June 2020)**.
- BARONE G. and S. MOCETTI, *Intergenerational mobility in the very long run: Florence 1427-2011*, Review of Economic Studies, v. 88, 4, pp. 1863–1891, **WP 1060 (April 2016)**.
- BARTOCCI A., L. BURLON, A. NOTARPIETRO and M. PISANI, *Macroeconomic effects of non-standard monetary policy measures in the Euro Area: the role of corporate bond purchases*, The Manchester School, v. 89, S1, pp. 97-130, **WP 1241 (October 2019)**.
- BATINI N., A. CANTELMO, G. MELINA and S. VILLA, *How loose, how tight? A measure of monetary and fiscal stance for the euro area*, Oxford Economic Papers, v. 73, 4, pp. 1536-1556, **WP 1295 (September 2020)**.
- BENETTON M. and D. FANTINO, *Targeted monetary policy and bank lending behavior*, Journal of Financial Economics, v. 142, 1, pp. 404-429, **WP 1187 (September 2018)**.
- BUSETTI F., M. CAIVANO and D. DELLE MONACHE, *Domestic and global determinants of inflation: evidence from expectile regression*, Oxford Bulletin of Economics and Statistics, v. 83, 4, pp. 982-1001, **WP 1225 (June 2019)**.
- BUSETTI F., M. CAIVANO, D. DELLE MONACHE and C. PACELLA, *The time-varying risk of Italian GDP*, Economic Modelling, v. 101, Article 105522, **WP 1288 (July 2020)**.
- BUSETTI F., S. NERI, A. NOTARPIETRO and M. PISANI, *Monetary Policy strategies in the new normal: a model-based analysis for the Euro Area*, Journal of Macroeconomics, v. 70, Article 103366, **WP 1308 (December 2020)**.
- CAPOLONGO A. and C. PACELLA, *Forecasting inflation in the Euro Area: countries matter*, Empirical Economics, v. 61, 4, pp. 2477-2499, **WP 1224 (June 2019)**.
- CARMIGNANI A., G. DE BLASIO, C. DEMMA and A. D'IGNAZIO, *Urbanization and firm access to credit*, Journal of Regional Science, v. 61, 3, pp. 597-622, **WP 1222 (June 2019)**.
- CORNELI F., *Financial integration without financial development*, Atlantic Economic Journal, v. 49, 2, pp. 201-220, **WP 1120 (June 2017)**.
- COVA P., P. PAGANO, A. NOTARPIETRO and M. PISANI, *Secular stagnation, R&D, public investment and monetary policy: a global-model perspective*, Macroeconomic Dynamics, v. 25, 5, pp. 1267-1287, **WP 1156 (December 2017)**.
- DE PHILIPPIS M., *Multitask agents and incentives: the case of teaching and research for university professors*, Economic Journal, v. 131, 636, pp. 1643-1681, **WP 1042 (December 2015)**.
- DEL PRETE S. and M. L. STEFANI, *Women as "Gold Dust": gender diversity in top boards and the performance of Italian banks*, Economic Notes, Monte dei Paschi di Siena, v. 50, 2, e12183, **WP 1014 (June 2015)**.

- FERRERO G., M. LOBERTO and M. MICCOLI, *The assets' pledgeability channel of unconventional monetary policy*, Economic Inquiry, v. 59, 4, pp. 1547-1568, **WP 1119 (June 2017)**.
- FIDORA M., C. GIORDANO and M. SCHMITZ, *Real exchange rate misalignments in the Euro Area*, Open Economies Review, v. 32, 1, pp. 71-107, **WP 1162 (January 2018)**.
- GAMBACORTA L., G. RICOTTI, S. SUNDARESAN and Z. WANG, *Tax effects on bank liability structure*, European Economic Review, v. 138, Article 103820, **WP 1101 (February 2017)**.
- HERTWECK M., V. LEWIS and S. VILLA, *Going the extra mile: effort by workers and job-seekers*, Journal of Money, Credit and Banking, v. 54, 8, pp. 2099-2127, **WP 1277 (June 2020)**.
- LI F., A. MERCATANTI, T. MAKINEN and A. SILVESTRINI, *A regression discontinuity design for ordinal running variables: evaluating central bank purchases of corporate bonds*, The Annals of Applied Statistics, v. 15, 1, pp. 304-322, **WP 1213 (March 2019)**.
- LOSCHIAVO D., *Big-city life (dis)satisfaction? The effect of urban living on subjective well-being*, Journal of Economic Behavior & Organization, vol. 192, pp. 740-764, **WP 1221 (June 2019)**.
- LOSCHIAVO D., *Household debt and income inequality: evidence from Italian survey data*, Review of Income and Wealth, v. 67, 1, pp. 61-103, **WP 1095 (January 2017)**.
- METELLI L. and F. NATOLI, *The international transmission of US tax shocks: a proxy-SVAR approach*, IMF Economic Review, v. 69, 2, pp. 325-356, **WP 1223 (June 2019)**.
- NISPI LANDI V. and A. SCHIAVONE, *The effectiveness of capital controls*, Open Economies Review, v. 32, 1, pp. 183-211, **WP 1200 (November 2018)**.
- PAPETTI A., *Demographics and the natural real interest rate: historical and projected paths for the Euro Area*, Journal of Economic Dynamics and Control, v. 132, Article 04209, **WP 1306 (November 2020)**.
- PEREDA FERNANDEZ S., *Copula-based random effects models for clustered data*, Journal of Business & Economic Statistics, v. 39, 2, pp. 575-588, **WP 1092 (January 2017)**.

2022

- ANDINI M., M. BOLDRINI, E. CIANI, G. DE BLASIO, A. D'IGNAZIO and A. PALADINI, *Machine learning in the service of policy targeting: the case of public credit guarantees*, Journal of Economic Behavior & Organization, v. 198, pp. 434-475, **WP 1206 (February 2019)**.
- ANGELICO C., J. MARCUCCI, M. MICCOLI and F. QUARTA, *Can we measure inflation expectations using twitter?*, Journal of Econometrics, v. 228, 2, pp. 259-277, **WP 1318 (February 2021)**.
- BARTOCCI A., A. NOTARPIETRO and M. PISANI, *Covid-19 shock and fiscal-monetary policy mix in a monetary union*, Economic challenges for Europe after the pandemic, Springer Proceedings in Business and Economics, Berlin-Heidelberg, Springer, **WP 1313 (December 2020)**.
- BOTTERO M., C. MINOIU, J. PEYDRÒ, A. POLO, A. PRESBITERO and E. SETTE, *Expansionary yet different: credit supply and real effects of negative interest rate policy*, Journal of Financial Economics, v. 146, 2, pp. 754-778, **WP 1269 (March 2020)**.
- BRONZINI R., A. D'IGNAZIO and D. REVELLI, *Financial structure and bank relationships of Italian multinational firms*, Journal of Multinational Financial Management, v. 66, Article 100762, **WP 1326 (March 2021)**.
- CANTELMO A., *Rare disasters, the natural interest rate and monetary policy*, Oxford Bulletin of Economics and Statistics, v. 84, 3, pp. 473-496, **WP 1309 (December 2020)**.
- CARRIERO A., F. CORSELLO and M. MARCELLINO, *The global component of inflation volatility*, Journal of Applied Econometrics, v. 37, 4, pp. 700-721, **WP 1170 (May 2018)**.
- CIAPANNA E. and G. ROVIGATTI, *The grocery trolley race in times of Covid-19. Evidence from Italy*, Italian Economic Journal / Rivista italiana degli economisti, v. 8, 2, pp. 471-498, **WP 1341 (June 2021)**.
- CONTI A. M., A. NOBILI and F. M. SIGNORETTI, *Bank capital requirement shocks: a narrative perspective*, European Economic Review, v.151, Article 104254, **WP 1199 (November 2018)**.
- FAIELLA I. and A. MISTRETTA, *The net zero challenge for firms' competitiveness*, Environmental and Resource Economics, v. 83, pp. 85-113, **WP 1259 (February 2020)**.
- FERRIANI F. and G. VERONESE, *Hedging and investment trade-offs in the U.S. oil industry*, Energy Economics, v. 106, Article 105736, **WP 1211 (March 2019)**.
- GUISSO L., A. POZZI, A. TSOY, L. GAMBACORTA and P. E. MISTRULLI, *The cost of steering in financial markets: evidence from the mortgage market*, Journal of Financial Economics, v.143, 3, pp. 1209-1226, **WP 1252 (December 2019)**.
- LAMORGESE A. and D. PELLEGRINO, *Loss aversion in housing appraisal: evidence from Italian homeowners*, Journal of Housing Economics, v. 56, Article 101826, **WP 1248 (November 2019)**.

- LI F., T. MÄKINEN, A. MERCATANTI and A. SILVESTRINI, *Causal analysis of central bank holdings of corporate bonds under interference*, *Economic Modelling*, v.113, Article 105873, **WP 1300 (November 2020)**.
- LOBERTO M., A. LUCIANI and M. PANGALLO, *What do online listings tell us about the housing market?*, *International Journal of Central Banking*, v. 18, 4, pp. 325-377, **WP 1171 (April 2018)**.
- MIRENDA L., M. SAURO and L. RIZZICA, *The economic effects of mafia: firm level evidence*, *American Economic Review*, vol. 112, 8, pp. 2748-2773, **WP 1235 (October 2019)**.
- MOCETTI S., G. ROMA and E. RUBOLINO, *Knocking on parents' doors: regulation and intergenerational mobility*, *Journal of Human Resources*, v. 57, 2, pp. 525-554, **WP 1182 (July 2018)**.
- PERICOLI M. and M. TABOGA, *Nearly exact Bayesian estimation of non-linear no-arbitrage term-structure models*, *Journal of Financial Econometrics*, v. 20, 5, pp. 807-838, **WP 1189 (September 2018)**.
- ROSSI P. and D. SCALISE, *Financial development and growth in European regions*, *Journal of Regional Science*, v. 62, 2, pp. 389-411, **WP 1246 (November 2019)**.
- SCHIVARDI F., E. SETTE and G. TABELLINI, *Credit misallocation during the European financial crisis*, *Economic Journal*, v. 132, 641, pp. 391-423, **WP 1139 (September 2017)**.
- TABOGA M., *Cross-country differences in the size of venture capital financing rounds: a machine learning approach*, *Empirical Economics*, v. 62, 3, pp. 991-1012, **WP 1243 (November 2019)**.

2023

- APRIGLIANO V., S. EMILIOZZI, G. GUAITOLI, A. LUCIANI, J. MARCUCCI and L. MONTEFORTE, *The power of text-based indicators in forecasting Italian economic activity*, *International Journal of Forecasting*, v. 39, 2, pp. 791-808, **WP 1321 (March 2021)**.
- BARTOCCI A., A. NOTARPIETRO and M. PISANI, *Non-standard monetary policy measures in non-normal times*, *International Finance*, v. 26, 1, pp. 19-35, **WP 1251 (November 2019)**.
- DI ADDARIO S., P. KLINE, R. SAGGIO and M. SØLVSTEN, *It ain't where you're from, it's where you're at: hiring origins, firm heterogeneity, and Wages*, *Journal of Econometrics*, v. 233, 2, pp. 340-374, **WP 1374 (June 2022)**.
- FERRIANI F., *Issuing bonds during the Covid-19 pandemic: was there an ESG premium?*, *International Review of Financial Analysis*, v. 88, Article 102653, **WP 1392 (November 2022)**.
- GIORDANO C., *Revisiting the real exchange rate misalignment-economic growth nexus via the across-sector misallocation channel*, *Review of International Economics*, v. 31, 4, pp. 1329-1384, **WP 1385 (October 2022)**.
- GUGLIELMINETTI E., M. LOBERTO and A. MISTRETTA, *The impact of COVID-19 on the European short-term rental market*, *Empirica*, v. 50, 3, pp. 585-623, **WP 1379 (July 2022)**.
- LILLA F., *Volatility bursts: a discrete-time option model with multiple volatility components*, *Journal of Financial Econometrics*, v. 21, 3, pp. 678-713, **WP 1336 (June 2021)**.
- LOBERTO M., *Foreclosures and house prices*, *Italian Economic Journal / Rivista italiana degli economisti*, v. 9, 1, pp. 397-424, **WP 1325 (March 2021)**.
- LOMBARDI M. J., M. RIGGI and E. VIVIANO, *Worker's bargaining power and the Phillips curve: a micro-macro analysis, and wages*, *Journal of the European Economic Association*, v. 21, 5, pp. 1905-1943, **WP 1302 (November 2020)**.
- NERI S., *Long-term inflation expectations and monetary policy in the Euro Area before the pandemic*, *European Economic Review*, v. 154, Article 104426, **WP 1357 (December 2021)**.
- ORAME A., *Bank lending and the European debt crisis: evidence from a new survey*, *International Journal of Central Banking*, v. 19, 1, pp. 243-300, **WP 1279 (June 2020)**.
- RIZZICA L., G. ROMA and G. ROVIGATTI, *The effects of shop opening hours deregulation: evidence from Italy*, *The Journal of Law and Economics*, v. 66, 1, pp. 21-52, **WP 1281 (June 2020)**.
- TANZI G. M., *Scars of youth non-employment and labour market conditions*, *Italian Economic Journal / Rivista italiana degli economisti*, v. 9, 2, pp. 475-499, **WP 1312 (December 2020)**.

FORTHCOMING

- BALTRUNAITE A., M. CANNELLA, S. MOCETTI and G. ROMA, *Board composition and performance of state-owned enterprises: quasi experimental evidence*, *The Journal of Law, Economics, and Organization*, **WP 1328 (April 2021)**.
- CIAPANNA E., S. MOCETTI and A. NOTARPIETRO, *The macroeconomic effects of structural reforms: an empirical and model-based approach*, *Economic Policy*, **WP 1303 (November 2022)**.

"TEMI" LATER PUBLISHED ELSEWHERE

- FERRARI A. and V. NISPI LANDI, *Whatever it takes to save the planet? Central banks and unconventional green policy*, *Macroeconomic Dynamics*, **WP 1320 (February 2021)**.
- FERRARI A. and V. NISPI LANDI, *Toward a green economy: the role of central bank's asset purchases*, *International Journal of Central Banking*, **WP 1358 (February 2022)**.
- MICHELANGELI V. and E. VIVIANO, *Can internet banking affect households' participation in financial markets and financial awareness?*, *Journal of Money, Credit and Banking*, **WP 1329 (April 2021)**.
- MISTRETTA A., *Synchronization vs transmission: the effect of the German slowdown on the Italian business cycle*, *International Journal of Central Banking*, **WP 1346 (October 2021)**.
- RAINONE E., *Real-time identification and high frequency analysis of deposits outflows*, *Journal of Financial Econometrics*, **WP 1319 (February 2021)**.